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FLUCTUATION IN THE ZERO  
SET OF THE PARABOLIC  
GAUSSIAN ANALYTIC  
FUNCTION

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# Abstract

In this project we study the fluctuation of the zero set process of the *parabolic Gaussian analytic function*, denoted  $\mathbb{S}^2$ -GAF and where  $\mathbb{S}^2$  is the Riemann sphere. There exist several ways to measure such fluctuations. One of them is to compute the *variance* of certain variables counting the number of points of the process inside a given region. Some asymptotics of such variables will lead us to conclude that the  $\mathbb{S}^2$ -GAF process is more rigid than the Poisson process on  $\mathbb{S}^2$  having, in mean, the same number of points as the  $\mathbb{S}^2$ -GAF process. Also, we will see that the  $\mathbb{S}^2$ -GAF process tends, as the intensity goes to infinity, to the planar GAF. Another point of view to study the fluctuations of the  $\mathbb{S}^2$ -GAF is the so-called *large deviations*, i.e., to measure how certain linear statistics deviate from its average by a fraction of its same average. The latter allows us to estimate the *hole probability*, i.e., the probability that the point process does not meet a given disk.

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# Chapter 1

## Introduction

The main object of study in this project is a specific type of *point process* in the Riemann sphere  $\mathbb{S}^2 := \mathbb{C} \cup \{\infty\}$ , that is, a random sequence of points in  $\mathbb{S}^2$ . Probably the best known point process is the *Poisson process*, denoted  $\mathcal{X}$ , which is characterized by two properties:

- For all  $A \subseteq \mathbb{S}^2$ , the random variable  $n_A := \#(A \cap \mathcal{X})$  follows a Poisson distribution of parameter  $\lambda(A)$ , which is the area of  $A$  in  $\mathbb{S}^2$ .
- If for all  $A, B \subset \mathbb{S}^2$  we have  $A \cap B = \emptyset$ , then  $n_A$  and  $n_B$  are independent random variables.

Independence is natural sometimes, but sometimes is not. For example, some physical phenomena in quantum mechanics cannot be explained by using such a process, due to the clumping points. In the 90's some physicists realized that a good model for point processes with local repulsion are the zero sets of some *random analytic functions*.

### 1.1 Gaussian analytic functions

Several random analytic functions can be considered, but we are going to focus on *Gaussian analytic functions*, GAFs for short, and see some properties of their zero sets.

We say that the random variable  $Z$  follows a *standard complex Gaussian* distribution, denoted  $Z \sim N_{\mathbb{C}}(0, 1)$ , if its density function, with respect to the Lebesgue measure, is

$$f_Z(z) = \frac{1}{\pi} e^{-|z|^2}, \quad z \in \mathbb{C}.$$

Now, assume that  $(e_n)_{n=0}^{+\infty}$  is a sequence of analytic functions in a region  $\Omega \subseteq \mathbb{C}$  and that  $(\xi_n)_{n=0}^{+\infty}$  is a sequence of i.i.d.  $N_{\mathbb{C}}(0, 1)$  random variables. Under some assumptions about convergence, we say that  $f$  is a GAF in  $\Omega$  if:

$$f(z) = \sum_{n=0}^{+\infty} \xi_n e_n(z), \quad z \in \Omega.$$

From this definition and fixing  $z \in \Omega$ ,  $f$  follows a complex Gaussian distribution with null mean due to the linear combination of  $(e_n)_{n=0}^{+\infty}$  and  $(\xi_n)_{n=0}^{+\infty} \sim N_{\mathbb{C}}(0, 1)$  and since the linear combination of Gaussian random variables is a Gaussian random variable. So, if we denote the *normalized GAF* as

$$\hat{f}(z) = \frac{f(z)}{\sqrt{\text{Var}[f(z)]}}, \quad z \in \Omega,$$

it follows a  $N_{\mathbb{C}}(0, 1)$  distribution. We have that the *covariance kernel* of a GAF  $f$  is

$$\mathcal{K}_f(z, w) := \text{Cov} [f(z), \overline{f(w)}] = \mathbb{E} [f(z)\overline{f(w)}] = \sum_{n=0}^{+\infty} e_n(z)\overline{e_n(w)}, \quad z, w \in \Omega.$$

In particular we have:

$$\text{Var}[f(z)] = \sqrt{\mathcal{K}_f(z, z)}.$$

Therefore, all the probabilistic properties of a GAF are encoded in its covariance kernel.

Some remarkable properties about GAFs are:

- A GAF is an analytic function in  $\Omega$  a.s. Fixed  $z \in \Omega$ ,  $f(z)$  converges up to a possible set, that depends on  $z$ , with probability zero. The problem is that the uncountable union of these sets can be the whole space or a significant part of it. This can be solved by using a version of Kolmogorov's inequality for Hilbert spaces.
- A standard way to construct a GAF on a given space  $\Omega$  as in the definition is to consider the orthonormal basis  $(e_n)_{n=0}^{+\infty}$  of a Hilbert space  $\mathcal{H}$  of analytic functions in  $\Omega$ . In that case the covariance kernel coincides with the *Bergman kernel* of  $\mathcal{H}$ , defined as

$$B(z, w) = \sum_{n=0}^{+\infty} e_n(z)\overline{e_n(w)}, \quad z, w \in \Omega.$$

Notice that  $B_w(z) = B(z, w)$  is the *reproducing kernel* of  $\mathcal{H}$  at  $w \in \Omega$ . Such kernel is independent of the choice of the orthonormal basis.

- Let  $(e_n)_{n=0}^{+\infty}$  be an orthonormal basis in  $\mathcal{H}$ . Then, the GAF does not belong to  $\mathcal{H}$  a.s., because on the contrary we would have that  $\sum_{n=0}^{+\infty} |\xi_n|^2$  converges, something that happens with probability zero for  $(\xi_n)_{n=0}^{+\infty}$  i.i.d. with  $N_{\mathbb{C}}(0, 1)$  distribution.

As canonical examples of this construction, we introduce three of the most studied families of Hilbert spaces of analytic functions.

- The *planar space* or the *Bargmann-Fock space* with real parameter  $L > 0$  in  $\mathbb{C}$  is defined as

$$\mathcal{F}_L := \left\{ f \in \mathcal{A}(\mathbb{C}) : \|f\|_{\mathcal{F}_L}^2 = \frac{L}{\pi} \int_{\mathbb{C}} |f(z)|^2 e^{-L|z|^2} dm(z) < +\infty \right\},$$

where  $z \in \mathbb{C}$  and  $\mathcal{A}(\mathbb{C})$  is the space of analytic functions in  $\mathbb{C}$ . The orthonormal basis obtained by normalizing the monomials  $z^n$ ,  $n \in \mathbb{N}$ , is:

$$e_n(z) = \sqrt{\frac{L^n}{n!}} z^n.$$

The GAF, also called  $\mathbb{C}$ -GAF or *planar GAF*, is therefore

$$f_L(z) = \sum_{n=0}^{+\infty} \xi_n \sqrt{\frac{L^n}{n!}} z^n,$$

where  $(\xi_n)_{n=0}^{+\infty}$  is a sequence of i.i.d.  $N_{\mathbb{C}}(0, 1)$  random variables. The covariance kernel is

$$\mathcal{K}_{f_L}(z, w) = \sum_{n=0}^{+\infty} \frac{L^n}{n!} (z\bar{w})^n = e^{Lz\bar{w}}.$$

- The *hyperbolic space* or the *weighted Bergman space* with real parameter  $L > 1$  in  $\mathbb{D}$  is defined as

$$\mathcal{B}_L := \left\{ f \in \mathcal{A}(\mathbb{D}) : \|f\|_{\mathcal{B}_L}^2 = \frac{L}{\pi} \int_{\mathbb{D}} |f(z)|^2 (1 - |z|^2)^{L-2} dm(z) < +\infty \right\},$$

where  $z \in \mathbb{D}$  and  $\mathcal{A}(\mathbb{D})$  is the space of analytic functions in  $\mathbb{D}$ . An orthonormal basis is:

$$e_n(z) = \binom{L+n-1}{n}^{1/2} z^n, \quad n \geq 0.$$

The GAF, also called  $\mathbb{D}$ -GAF or *hyperbolic GAF*, is then

$$f_L(z) = \sum_{n=0}^{+\infty} \xi_n \binom{L+n-1}{n}^{1/2} z^n,$$

which has sense for  $L > 0$ . The covariance kernel is

$$\mathcal{K}_{f_L}(z, w) = \sum_{n=0}^{+\infty} \binom{L+n-1}{n} (z\bar{w})^n = (1 - z\bar{w})^{-L}.$$

- The *parabolic space* or the *space of polynomials of degree at most  $L \in \mathbb{N}$*  in  $\mathbb{C}$  is described as:

$$\mathcal{P}_L := \left\{ f \in P_L[\mathbb{C}] : \|f\|_{\mathcal{P}_L}^2 = \frac{L+1}{\pi} \int_{\mathbb{C}} \frac{|f(z)|^2}{(1 + |z|^2)^{L+2}} dm(z) < +\infty \right\},$$

where  $z \in \mathbb{C}$  and  $P_L[\mathbb{C}]$  is the vector space of polynomials of degree at most  $L$  with complex coefficients. The space  $P_L[\mathbb{C}]$  can be seen as the projection to  $\mathbb{C}$  of the space of sections of the  $L$ -th power of the canonical bundle of  $\mathbb{S}^2$ . The norm defined in  $\mathcal{P}_L$  makes sense. Indeed, the term

$$\frac{|f(z)|^2}{(1 + |z|^2)^L}$$

is the normalization by the degree of  $f$ , and

$$\frac{dm(z)}{\pi(1 + |z|^2)^2}$$

is the area measure of  $\mathbb{S}^2$  projected to  $\mathbb{C}$ .

An orthonormal basis is:

$$e_n(z) = \binom{L}{n}^{1/2} z^n, \quad 0 \leq n \leq L.$$

The GAF, also called  $\mathbb{S}^2$ -GAF or *parabolic GAF*, is then

$$f_L(z) = \sum_{n=0}^L \xi_n \binom{L}{n}^{1/2} z^n,$$

and the covariance kernel is

$$\mathcal{K}_{f_L}(z, w) = \sum_{n=0}^L \binom{L}{n} (z\bar{w})^n = (1 + z\bar{w})^L.$$

## 1.2 First intensity and the Edelman-Kostlan formula

We will focus our attention on this last space and on the zero set of an  $\mathbb{S}^2$ -GAF  $f_L$  of parameter  $L \in \mathbb{N}$  in a region  $\Omega \subseteq \mathbb{C}$ . The zero set of  $f_L$ , denoted  $\mathcal{Z}_{f_L}$ , will be studied through its *empirical measure*

$$\nu_{f_L} = \sum_{a \in \mathcal{Z}_{f_L}} \delta_a = \frac{1}{2\pi} \Delta \log |f_L|,$$

where  $\delta_a$  is the Dirac delta measure at  $a$ . Notice that  $\nu_{f_L}$  is a measure supported precisely on the zeros of  $f_L$ . The *first intensity* of the GAF  $f_L$  is the measure  $\mathbb{E}[\nu_{f_L}]$  defined by the action

$$\int_{\Omega} \varphi d\mathbb{E}[\nu_{f_L}] = \mathbb{E} \left[ \int_{\Omega} \varphi d\nu_{f_L} \right], \quad \varphi \in C_c^\infty(\Omega).$$

The first intensity measures the average number of points of the point process. According to the well-known *Edelman-Kostlan formula*:

$$\mathbb{E}[\nu_{f_L}] = \frac{1}{2\pi} \Delta \log \mathbb{E}[|f_L|] = \frac{1}{2\pi} \Delta \log \sqrt{\mathcal{K}_{f_L}(z, z)}, \quad z \in \Omega.$$

For the spaces before introduced, we have that:

- For a  $\mathbb{C}$ -GAF  $f_L$  of real parameter  $L > 0$ , the first intensity is

$$\mathbb{E}[\nu_{f_L}] = \frac{1}{2\pi} \Delta \log e^{L|z|^2/2} = \frac{L}{\pi} dm(z),$$

where  $dm$  stands for the Lebesgue measure on the plane.

- For a  $\mathbb{D}$ -GAF  $f_L$  of real parameter  $L > 1$ , the first intensity is

$$\mathbb{E}[\nu_{f_L}] = \frac{1}{4\pi} \Delta \log(1 - zw)^{-L} = \frac{L}{\pi} \frac{dm(z)}{(1 - |z|^2)^2}.$$

- For an  $\mathbb{S}^2$ -GAF  $f_L$  of parameter  $L \in \mathbb{N}$ , the first intensity is

$$\mathbb{E}[\nu_{f_L}] = \frac{1}{4\pi} \Delta \log [(1 + |z|^2)^L] = \frac{L}{\pi} \frac{dm(z)}{(1 + |z|^2)^2}.$$

A remarkable feature of these processes is the invariance by the natural translations on each space.

- For a  $\mathbb{C}$ -GAF the zero point process is invariant by translations

$$\phi_a(z) = z - a, \quad z, a \in \mathbb{C}.$$

- For a  $\mathbb{D}$ -GAF the zero point process is invariant by automorphisms in  $\mathbb{D}$

$$\phi_a(z) = \frac{z - a}{1 - \bar{a}z} e^{i\theta}, \quad z, a \in \mathbb{D}, \quad \theta \in [0, 2\pi).$$

- For an  $\mathbb{S}^2$ -GAF the zero point process is invariant by rotations in  $\mathbb{S}^2$ , which in the  $\mathbb{C}$ -chart are seen as the Möbius transformations

$$\phi_a(z) = \frac{z - a}{1 + \bar{a}z}, \quad z, a \in \mathbb{C}.$$

Since the first intensity determines the distribution in mean of  $\mathcal{Z}_{f_L}$ , it is also invariant by the suitable transformations just introduced.

## 1.3 Fluctuations of the parabolic GAF

Having this basic background in GAF theory, we can face the main problem of this project. We know how the zeros of a GAF are distributed in average according to the Edelman-Kostlan formula, but how do they interact? Or what is equivalent, how do they *fluctuate*?

We quantify this fact from different points of view:

- The variance of the random variable  $\nu_{f_L}(U) = \#(\mathcal{Z}_{f_L} \cap U)$ , where  $U \subseteq \mathbb{C}$ . We are going to make an exhaustive study of the variance, denoted  $\mathbb{V}$ , in a disk of radius  $2\rho$ ,  $\rho > 0$ , endowed with the *chordal metric* (see (2.4.4)). Denoting the chordal disk as  $D_{ch} := D_{ch}(z_0, 2\rho)$ , for  $z_0 \in \mathbb{C}$ , using the definition of the variance and the properties of the GAF we prove that

$$\mathbb{V}[\nu_{f_L}(D_{ch})] = \frac{L^2}{2\pi} \rho \sqrt{1 - \rho^2} \int_0^{4\rho^2(1-\rho^2)} \frac{(1-x)^L}{1 - (1-x)^L} \frac{\sqrt{x}}{1-x} \frac{1}{\sqrt{1 - \frac{x}{4\rho^2(1-\rho^2)}}} dx.$$

For the proof see Theorem 3.2.2.

This is the integral of a positive function in a bounded interval, from which we can extract some information. For example:

1. *Asymptotics as  $L \rightarrow +\infty$*  (see Subsection 3.2.3). We will show that:

**Proposition 1.3.1.** *Let  $f_L$  be an  $\mathbb{S}^2$ -GAF of intensity  $L \in \mathbb{N}$ . Consider a chordal disk  $D_{ch} := D_{ch}(z_0, 2\rho)$ , for  $z_0 \in \mathbb{C}$ . Then*

$$\mathbb{V}[\nu_{f_L}(D_{ch})] = \left( \frac{\sqrt{L}}{4\sqrt{\pi}} \zeta(3/2) \rho \sqrt{1 - \rho^2} \right) (1 + o(1)), \quad \text{as } L \rightarrow +\infty.$$

Here  $\zeta$  stands for the Riemann's zeta function and  $o(1)$  is a term tending to 0 as  $L \rightarrow +\infty$ .

For the Poisson process  $\mathcal{X}$  with underlying measure

$$\frac{L}{\pi} dm,$$

the random variable

$$n_L(D(0, r)) := \#(\mathcal{X} \cap D(0, r)) \tag{1.3.1}$$

has the same average number of points as our GAF:

$$\mathbb{E}[n_L(D(0, r))] = Lr^2. \tag{1.3.2}$$

But the variance is much larger:

$$\mathbb{V}[n_L(D(0, r))] = Lr^2. \tag{1.3.3}$$

It is in this sense that the GAF process is more *rigid*.

2. *Asymptotics as  $\rho \rightarrow 0$*  (see Subsection 3.2.4). It is intuitive that the variance will tend to zero as  $\rho \rightarrow 0$ . We will quantify the speed of convergence. More precisely:

**Proposition 1.3.2.** *Let  $f_L$  be an  $\mathbb{S}^2$ -GAF of intensity  $L \in \mathbb{N}$ . Consider a chordal disk  $D_{ch} := D_{ch}(z_0, 2\rho)$ , for  $z_0 \in \mathbb{C}$ . Then*

$$\mathbb{V}[\nu_{f_L}(D_{ch})] = L\rho^2(1 + o(1)) \quad \text{as } \rho \rightarrow 0.$$

In here, we see that the speed of convergence is equivalent to the Poisson process as the radius tends to zero (see 1.3.3).

3.  $\mathcal{F}_1$  as the limit of  $\mathcal{P}_L$  as  $L \rightarrow +\infty$  (see Subsection 3.2.5). There is a result stating that the functions in  $\mathcal{F}_1$  can be seen as limits of rescaled polynomials of  $\mathcal{P}_L$  as  $L \rightarrow +\infty$ . More precisely:

**Lemma 1.3.3.** *Given a GAF  $f_1^{\mathbb{C}} \in \mathcal{F}_1$  and a constant  $M > 0$ , there is  $L_0 \in \mathbb{N}$  such that for all  $L \geq L_0$ , there exist GAFs  $f_L^{\mathbb{S}^2} \in \mathcal{P}_L$  such that*

$$\int_{\{|z| \leq M/\sqrt{L}\}} \left| f_1^{\mathbb{C}}(\sqrt{L}z) - f_L^{\mathbb{S}^2}(z) \right|^2 e^{-L|z|^2} dm(z) \lesssim \frac{1}{L} \|f_1^{\mathbb{C}}\|_{\mathcal{F}_1}^2$$

and

$$\int_{\{|z| > M/\sqrt{L}\}} \frac{|f_L^{\mathbb{S}^2}(z)|^2}{\pi(1 + |z|^2)^{L+2}} dm(z) \lesssim \frac{1}{L} \|f_1^{\mathbb{C}}\|_{\mathcal{F}_1}^2.$$

In accordance with this lemma, we prove that:

**Proposition 1.3.4.** *The limit of the variance of an  $\mathbb{S}^2$ -GAF as  $L \rightarrow +\infty$  coincides with the variance of a  $\mathbb{C}$ -GAF of parameter  $L = 1$ , that is,*

$$\lim_{L \rightarrow +\infty} \mathbb{V} \left[ \nu_{f_L^{\mathbb{S}^2}} \left( D_{ch}(z_0, 2r/\sqrt{L}) \right) \right] = \mathbb{V} \left[ \nu_{f_1^{\mathbb{C}}} (D(z_0, r)) \right], \quad z_0 \in \mathbb{C}.$$

- *Large deviations* (see Section 4.1). Consider test-functions  $\varphi \in \mathcal{C}_c^2(\mathbb{C})$ . We define the *linear statistic* associated to  $\varphi$  as

$$I_L(\varphi) := \int_{\mathbb{C}} \varphi d\nu_{f_L}.$$

We have, by the Edelman-Kostlan formula,

$$\mathbb{E} [I_L(\varphi)] = \mathbb{E} \left[ \int_{\mathbb{C}} \varphi d\nu_{f_L} \right] = \int_{\mathbb{C}} \varphi d\mathbb{E}[\nu_{f_L}].$$

Here we will study how much  $I_L(\varphi)$  deviates from its mean  $\mathbb{E} [I_L(\varphi)]$  by a fraction of the same mean. More concretely:

**Theorem 1.3.5.** *For all  $\varphi \in \mathcal{C}_c^2(\mathbb{C})$  and for all  $\delta > 0$ , there exist constants  $c = c(\varphi, \delta)$  and  $L_0 = L_0(\varphi, \delta) \in \mathbb{N}$  such that, for all  $L \geq L_0$ ,*

$$\mathbb{P} \left[ \left| \frac{I_L(\varphi)}{\mathbb{E} [I_L(\varphi)]} - 1 \right| > \delta \right] \leq e^{-cL^2}.$$

- *Hole probability* (see Section 4.2). As a consequence of the large deviations we will estimate the probability that there is a hole in the zero point process, that is, a disk without zeros of  $f_L$ . More precisely:

**Theorem 1.3.6.** *For a given  $\rho > 0$ , there exist  $C_1 = C_1(\rho) > 0$ ,  $C_2 = C_2(\rho) > 0$  and  $L_0 \in \mathbb{N}$  such that, for all  $L \geq L_0$  and a disk  $D_{ch} := D_{ch}(z_0, \rho) \subset \mathbb{C}$  with  $z_0 \in \mathbb{C}$ ,*

$$e^{-C_1 L^2} \leq \mathbb{P}[\mathcal{Z}_{f_L} \cap D_{ch} = \emptyset] \leq e^{-C_2 L^2}.$$

For the Poisson process in the planar case with intensity  $L > 0$ , we have that

$$\mathbb{P}[n_{D(z_0, r)} = 0] = e^{-Lr^2}, \quad z_0 \in \mathbb{C}, \quad r > 0.$$

But for the  $\mathbb{S}^2$ -GAF we have lower and upper bounds such that the power of the exponential depends on  $L^2$ . Then, it is more unlikely to have a hole in the zero point process of an  $\mathbb{S}^2$ -GAF than in the Poisson process.

This manuscript is divided in three parts. Chapter 2 is devoted to introduce the basis of GAF theory. I followed [1], [2] and [9].

Chapter 3 is devoted to compute explicitly the variance of the random variables of the zero point process of a  $\mathbb{C}$ -GAF and an  $\mathbb{S}^2$ -GAF in suitable disks. Then we give the asymptotic results described above for an  $\mathbb{S}^2$ -GAF. Here I used [2], [5], [6] and [9].

In Chapter 4 we give a full description of the large deviations and the hole probability of an  $\mathbb{S}^2$ -GAF. The statements, results and proofs are analogous to [6]. Also [4] was helpful to write these pages.

# Chapter 2

## Preliminaries

In this chapter we will introduce the basic elements and results about the theory of *Gaussian analytic functions*. Since the project is focused on the zero set of Gaussian analytic functions on the Riemann sphere  $\mathbb{S}^2 := \mathbb{C} \cup \{\infty\}$ , we are going to explicit these elements on this space and the suitable Hilbert space we are going to use. The following pages are strongly based on [1] and [2], so all the statements and proofs come from those references. The source [9] was also helpful at some points.

### 2.1 Complex Gaussian distribution

Recall that a real-valued random variable  $X$  follows a *real Gaussian distribution*, denoted  $X \sim N_{\mathbb{R}}(\mu, \sigma^2)$ , if its density function is

$$f_X(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}, \quad x \in \mathbb{R},$$

where  $\mu \in \mathbb{R}$  is the *mean* and  $\sigma^2 \in (0, +\infty)$  is the *variance*. We can define also the complex version of the standard Gaussian distribution.

**Definition 2.1.1.** A complex-valued random variable  $Z$  follows a *standard complex Gaussian distribution*, denoted  $Z \sim N_{\mathbb{C}}(0, 1)$ , if its density function, with respect to the Lebesgue measure, is

$$f_Z(z) = \frac{1}{\pi} e^{-|z|^2}, \quad z \in \mathbb{C}.$$

We outline the following result:

**Proposition 2.1.2.** *i) If  $Z \sim N_{\mathbb{C}}(0, 1)$ , then  $|Z|^2$  is an exponential random variable of parameter 1, i.e.,  $\mathbb{P}[|Z|^2 > t] = e^{-t}$  for all  $t > 0$ .*

*ii) If  $(\xi_n)_{n=0}^{+\infty}$  is a sequence of i.i.d.  $N_{\mathbb{C}}(0, 1)$  random variables, then*

$$\limsup_{n \rightarrow +\infty} |\xi_n|^{1/n} = 1, \quad \text{a.s.}$$

*Proof.* i) By applying polar coordinates and the change of variable  $s = r^2$ , we have for all  $t > 0$ ,

$$\begin{aligned} \mathbb{P}[|Z|^2 > t] &= 1 - \mathbb{P}[|Z|^2 \leq t] = 1 - \int_{\{0 \leq |z|^2 \leq t\}} \frac{1}{\pi} e^{-|z|^2} dm(z) = 1 - \int_0^{\sqrt{t}} 2re^{-r^2} dr \\ &= 1 - \int_0^t e^{-s} ds = e^{-t}. \end{aligned}$$

ii) It is a consequence of the Borel-Cantelli lemma. See for instance [2], p. 15.  $\square$

## 2.2 Gaussian analytic functions

Here we introduce the main object of study of the project, the so-called *Gaussian analytic function*. For this we endow the space of analytic functions over a region  $\Omega$  with the topology of uniform convergence on compact sets of  $\Omega$ . Denote by  $\mathcal{A}(\Omega)$  the space just described. A very natural question is: how can we generate GAFs? The next definition addresses this.

**Definition 2.2.1.** Let  $(e_n)_{n=0}^{+\infty}$  be a sequence in  $\mathcal{A}(\Omega)$  and let  $(\xi_n)_{n=0}^{+\infty}$  be a sequence of i.i.d. random variables with  $N_{\mathbb{C}}(0, 1)$  distribution. Assume that  $\sum_{n=0}^{\infty} |e_n(z)|^2$  converges locally uniformly on  $\Omega$ . A *Gaussian analytic function* (GAF from now on) is the linear combination

$$f(z) = \sum_{n=0}^{+\infty} \xi_n e_n(z), \quad z \in \Omega. \quad (2.2.1)$$

Fixing  $z \in \Omega$ , the random variable  $f$  described in Definition 2.2.1 follows a complex Gaussian distribution with null mean due to the linear combination of  $(e_n)_{n=0}^{+\infty}$  and  $(\xi_n)_{n=0}^{+\infty} \sim N_{\mathbb{C}}(0, 1)$  and since the linear combination of Gaussian random variables is a Gaussian random variable. So, if we denote the *normalized GAF* as

$$\hat{f}(z) = \frac{f(z)}{\sqrt{\text{Var}[f(z)]}}, \quad z \in \Omega,$$

it follows a  $N_{\mathbb{C}}(0, 1)$  distribution. We have that the *covariance kernel* of a GAF  $f$  is

$$\mathcal{K}_f(z, w) := \text{Cov} \left[ f(z), \overline{f(w)} \right] = \mathbb{E} \left[ f(z) \overline{f(w)} \right] = \sum_{n=0}^{+\infty} e_n(z) \overline{e_n(w)}, \quad z, w \in \Omega. \quad (2.2.2)$$

In particular, we have

$$\text{Var}[f(z)] = \sqrt{\mathcal{K}_f(z, z)}.$$

Therefore, all the probabilistic properties of a GAF are encoded in its covariance kernel.

**Remark 2.2.2.** To justify the last equality of (2.2.2) notice that

$$\mathcal{K}_f(z, w) = \mathbb{E} \left[ f(z) \overline{f(w)} \right] = \sum_{n, m=0}^{+\infty} e_n(z) \overline{e_n(w)} \mathbb{E} \left[ \xi_n \overline{\xi_m} \right].$$

Since  $(\xi_n)_{n=0}^{+\infty}$  is a sequence of i.i.d. random variables with  $N_{\mathbb{C}}(0, 1)$  distribution, we have that

$$\mathbb{E} \left[ \xi_n \overline{\xi_m} \right] = \delta_{n, m},$$

where  $\delta_{n, m}$  denotes the Kronecker delta function. Thus the equality follows.

Let us state a few remarks and properties about GAFs:

- A GAF is an analytic function in  $\Omega$  a.s. Fixed  $z \in \Omega$ ,  $f(z)$  converges up to a possible set, that depends on  $z$ , with probability zero. The problem is that the uncountable union of these sets can be the whole space or a significant part of it. This can be solved by using a version of Kolmogorov's inequality for Hilbert spaces (see [2], Lemma 2.2.3).
- The radius of convergence of a GAF is computed with the conditions of growth of the sequence  $(e_n)_{n=0}^{+\infty}$  and with ii) of Proposition 2.1.2.
- A standard way to construct a GAF on a given space  $\Omega$  is to consider the orthonormal basis  $(e_n)_{n=0}^{+\infty}$  of a Hilbert space  $\mathcal{H}$  of analytic functions in  $\Omega$  and consider (2.2.1). In that case the covariance kernel (2.2.2) coincides with the *Bergman kernel* of  $\mathcal{H}$ , defined as

$$B(z, w) = \sum_{n=0}^{+\infty} e_n(z) \overline{e_n(w)}, \quad z, w \in \Omega.$$

Notice that  $B_w(z) = B(z, w)$  is the *reproducing kernel* of  $\mathcal{H}$  at  $w \in \Omega$ . Such kernel is independent of the choice of the orthonormal basis.

- Let  $(e_n)_{n=0}^{+\infty}$  be an orthonormal basis in  $\mathcal{H}$ . Then, (2.2.1) does not belong to  $\mathcal{H}$  a.s., because on the contrary we would have that  $\sum_{n=0}^{+\infty} |\xi_n|^2$  converges, something that happens with probability zero for  $(\xi_n)_{n=0}^{+\infty}$  i.i.d. with  $N_{\mathbb{C}}(0, 1)$  distribution.

**Lemma 2.2.3.** *The normalized kernel*

$$K(z, w) = \frac{|\mathcal{K}_f(z, w)|^2}{\mathcal{K}_f(z, z) \mathcal{K}_f(w, w)}$$

satisfies  $|K(z, w)| \leq 1$ .

*Proof.* To check this inequality just notice that, since  $B(z, w) = \mathcal{K}_f(z, w)$ ,

$$|\mathcal{K}_f(z, w)| = |(B_z, B_w)_{\mathcal{H}}| \leq \|B_z\|_{\mathcal{H}} \|B_w\|_{\mathcal{H}},$$

where  $\mathcal{H}$  is the Hilbert space we are working in and  $B_z$  (respectively  $B_w$ ) is the reproducing kernel of  $\mathcal{H}$  at the point  $z \in \Omega$  (respectively at  $w \in \Omega$ ). By squaring at both sides and passing the RHS to the left, we see that  $|K(z, w)| \leq 1$ .  $\square$

## 2.3 The Gamma function

The *Gamma function*, denoted  $\Gamma$ , is defined for all  $y > 0$  as

$$\Gamma(y) = \int_0^{+\infty} x^{y-1} e^{-x} dx.$$

We state a few well-known properties:

**Lemma 2.3.1.**    *i)*  $\Gamma(1) = 1$ .

*ii)* For all  $y > 0$ ,  $\Gamma(y + 1) = y\Gamma(y)$ .

*iii)*  $\Gamma(1/2) = \sqrt{\pi}$ .

*iv)* For all  $n \in \mathbb{N}$ ,  $\Gamma(n + 1) = n!$

*v)* For  $L \in \mathbb{N}$  and  $n \leq L$ :

$$\binom{L}{n} = \frac{\Gamma(L + 1)}{\Gamma(n + 1)\Gamma(L - n + 1)} = \frac{L!}{n!(L - n)!}. \quad (2.3.1)$$

*vi)* For all  $a \in \mathbb{R}$ ,

$$\lim_{\ell \rightarrow +\infty} \frac{\Gamma(\ell)}{\Gamma(\ell + a)} = \ell^{-a}. \quad (2.3.2)$$

We shall also use the *Beta function*, denoted  $B$ , which is defined, for all  $x, y > 0$ , as

$$B(x, y) = \int_0^1 s^{x-1}(1-s)^{y-1} ds = \int_0^{+\infty} \frac{s^{x-1}}{(1+s)^{x+y}} ds = \frac{\Gamma(x)\Gamma(y)}{\Gamma(x+y)}. \quad (2.3.3)$$

## 2.4 The $\mathbb{C}$ -GAF and the $\mathbb{S}^2$ -GAF

We are going to introduce two families of Hilbert spaces of analytic functions that depend on a positive parameter  $L$ . These will be used to generate GAFs and to study the properties of their zero sets.

### 2.4.1 The spaces $\mathcal{F}_L$

Given a real value  $L > 0$ , we define the *planar space*, also known as the *Bargmann-Fock space*, as

$$\mathcal{F}_L := \left\{ f \in \mathcal{A}(\mathbb{C}) : \|f\|_{\mathcal{F}_L}^2 = \frac{L}{\pi} \int_{\mathbb{C}} |f(z)|^2 e^{-L|z|^2} dm(z) < +\infty \right\},$$

where  $dm$  stands for the Lebesgue measure. The factor  $L/\pi$  is chosen so that

$$\frac{L}{\pi} e^{-L|z|^2} dm(z), \quad z \in \mathbb{C},$$

is a probability measure, i.e.,  $\|1\|_{\mathcal{F}_L} = 1$ .

In this space,

$$e_n(z) = \sqrt{\frac{L^n}{n!}} z^n, \quad n \geq 0,$$

is an orthonormal basis. Let us show this property.

**Lemma 2.4.1.**  $(e_n)_{n=0}^{+\infty}$  is an orthonormal basis in  $\mathcal{F}_L$ .

*Proof.* We have to show that  $(e_n, e_m)_{\mathcal{F}_L} = \delta_{n,m}$  and that  $(e_n)_{n=0}^{+\infty}$  is complete. Let us focus on the first part. For  $n, m \in \mathbb{N}$  we have, using polar coordinates:

$$\begin{aligned} (e_n, e_m)_{\mathcal{F}_L} &= \frac{L}{\pi} \int_{\mathbb{C}} \sqrt{\frac{L^{n+m}}{n!m!}} e^{-L|z|^2} dm(z) \\ &= \frac{L}{\pi} \sqrt{\frac{L^{n+m}}{n!m!}} \int_0^{2\pi} \int_0^{+\infty} r^{n+m+1} e^{i\theta(n-m)} e^{-Lr^2} dr d\theta. \end{aligned}$$

However, if  $n \neq m$ ,

$$\int_0^{2\pi} e^{i\theta(n-m)} d\theta = 0. \quad (2.4.1)$$

Hence, if  $n \neq m$ , we conclude that  $(e_n, e_m)_{\mathcal{F}_L} = 0$ . Otherwise, if  $n = m$ , we have, by using the definition and properties of the Gamma function and the change of variable  $t = Lr^2$ , that

$$(e_n, e_n)_{\mathcal{F}_L} = \frac{L^{n+1}}{n!} \int_0^{+\infty} 2r^{2n+1} e^{-Lr^2} dr = \frac{1}{n!} \int_0^{+\infty} t^n e^{-t} dt = \frac{\Gamma(n+1)}{n!} = 1.$$

Thus  $(e_n)_{n=0}^{+\infty}$  is an orthonormal system in  $\mathcal{F}_L$ .

For completeness just notice that  $(e_n)_{n=0}^{+\infty}$  is a normalization of the monomial basis.  $\square$

Using Definition 2.2.1, for every  $L > 0$  we can generate the GAF

$$f_L(z) = \sum_{n=0}^{+\infty} \xi_n \sqrt{\frac{L^n}{n!}} z^n, \quad (2.4.2)$$

where  $(\xi_n)_{n=0}^{+\infty}$  is a sequence of i.i.d. random variables with  $N_{\mathbb{C}}(0, 1)$  distribution. This is known as the *planar Gaussian analytic function* ( $\mathbb{C}$ -GAF, for short). The covariance kernel of this GAF is

$$\mathcal{K}_{f_L}(z, w) = \sum_{n=0}^{+\infty} \frac{L^n}{n!} z^n \bar{w}^n = e^{Lz\bar{w}}. \quad (2.4.3)$$

### 2.4.2 The Riemann sphere $\mathbb{S}^2$

We deal mainly with GAFs on the Riemann sphere. Let us reveal some properties of it. The topology is defined in the following way:

- If  $w \in \mathbb{C}$  the neighbourhood system is generated by the family of disks  $\{D(w, r)\}_{r>0}$ .
- Otherwise, if  $w = \infty$  the neighbourhood system is generated by the family  $\mathbb{S}^2 \setminus \{\overline{D(0, r)}\}_{r>0}$ .

Let us set  $S^2 := \{(x_1, x_2, x_3) \in \mathbb{R}^3 : x_1^2 + x_2^2 + x_3^2 = 1\}$ . Denote  $N := (0, 0, 1)$ .

**Lemma 2.4.2.** *The stereographic projection*

$$p : S^2 \setminus \{N\} \longrightarrow \mathbb{C}$$

$$x = (x_1, x_2, x_3) \longmapsto \frac{x_1}{1 - x_3} + i \frac{x_2}{1 - x_3}$$

*establishes a homeomorphism between  $S^2 \setminus \{N\}$  and  $\mathbb{C}$ .*

*Proof.* The mapping  $p$  is well-defined, continuous and bijective. For the inverse mapping, take  $p(x) = a + ib$ , for all  $a, b \in \mathbb{R}$ . So, imposing  $a = x_1/(1 - x_3)$  and  $b = x_2/(1 - x_3)$ , we get

$$|z|^2 = a^2 + b^2 = \frac{x_1^2 + x_2^2}{(1 - x_3)^2} = \frac{1 - x_3^2}{(1 - x_3)^2} = \frac{1 + x_3}{1 - x_3}.$$

Thus, since  $a = (z + \bar{z})/2$  and  $b = (z - \bar{z})/2i$ , from

$$(1 - x_3)|z|^2 = 1 + x_3$$

we deduce that

$$x_3 = \frac{|z|^2 - 1}{1 + |z|^2}.$$

Also it follows that

$$x_1 = a(1 - x_3) = a \left( 1 - \frac{|z|^2 - 1}{1 + |z|^2} \right) = \frac{z + \bar{z}}{1 + |z|^2},$$

$$x_2 = b(1 - x_3) = b \left( 1 - \frac{|z|^2 - 1}{1 + |z|^2} \right) = \frac{z - \bar{z}}{i(1 + |z|^2)}.$$

Hence the inverse mapping is defined as

$$p^{-1} : \mathbb{C} \longrightarrow S^2 \setminus \{N\}$$

$$z \longmapsto \left( \frac{z + \bar{z}}{1 + |z|^2}, \frac{z - \bar{z}}{i(1 + |z|^2)}, \frac{|z|^2 - 1}{1 + |z|^2} \right),$$

and it is continuous. Therefore  $p$  is a homeomorphism between  $S^2 \setminus \{N\}$  and  $\mathbb{C}$ . We extend  $p$  to the mapping

$$p_\infty : S^2 \longrightarrow \mathbb{S}^2 \\ x \longmapsto \begin{cases} p(x), & \text{if } x \neq N, \\ \infty, & \text{if } x = N. \end{cases}$$

Then we have that  $p_\infty$  is bijective and that  $\lim_{x \rightarrow N} |p_\infty(x)| = \infty$ . This implies that  $p_\infty$  is a homeomorphism between  $S^2$  and  $\mathbb{S}^2$  and, since  $S^2$  is compact,  $\mathbb{S}^2$  is also compact. Therefore  $\mathbb{S}^2$  can be understood as the Alexandroff compactification of  $\mathbb{C}$  with the point  $\{\infty\}$ .  $\square$

### The chordal distance

The metric we are going to consider is the *chordal distance*, which is the Euclidian distance in  $\mathbb{R}^3$  projected to  $\mathbb{C}$  by the stereographic projection. It has the expression:

$$d_{ch}(z, w) := \frac{2|z - w|}{(1 + |z|^2)^{1/2}(1 + |w|^2)^{1/2}}, \quad z, w \in \mathbb{C}. \quad (2.4.4)$$

**Remark 2.4.3.** Let us state a few properties about the chordal distance:

i) We define the *chordal disk* as

$$D_{ch}(z_0, \rho) := \{z \in \mathbb{C} : d_{ch}(z_0, z) < \rho\}, \quad z_0 \in \mathbb{C}, \rho > 0.$$

ii) The relation between the radii of  $D(0, r)$  and  $D_{ch}(0, \rho)$ , for  $r, \rho > 0$ , is:

$$D_{ch}(0, \rho) = D\left(0, \frac{\rho}{\sqrt{4 - \rho^2}}\right)$$

and

$$D(0, r) = D_{ch}\left(0, \frac{2r}{\sqrt{1 + r^2}}\right).$$

iii) The expression

$$dm^*(z) := \frac{dm(z)}{\pi(1 + |z|^2)^2}, \quad z \in \mathbb{C},$$

called the *parabolic measure*, is the push-forward in  $\mathbb{C}$  of the surface area measure in  $\mathbb{S}^2$ , by the stereographic projection. We will see in Corollary 2.4.5 that the parabolic measure is invariant by the rotations of  $\mathbb{S}^2$  projected to  $\mathbb{C}$ .

iv) Let us compute the parabolic measure of a chordal disk  $D_{ch} := D_{ch}(z_0, \rho)$  with  $z_0 \in \mathbb{C}$  and  $\rho > 0$ . Denote

$$|z| \leq \tilde{\rho} := \frac{\rho}{\sqrt{4 - \rho^2}}.$$

Using polar coordinates and the change of variable  $t = 1 + r^2$ :

$$\begin{aligned} m^*(D_{ch}) &= \int_{D_{ch}} dm^*(z) = \int_{D_{ch}} \frac{dm(z)}{\pi(1+|z|^2)^2} = \int_0^{\tilde{\rho}} \frac{2r}{(1+r^2)^2} dr \\ &= \int_1^{1+\tilde{\rho}^2} \frac{dt}{t^2} = \frac{\tilde{\rho}^2}{1+\tilde{\rho}^2} = \frac{\rho^2}{4}. \end{aligned}$$

### Transformations

The transformations we consider are the rotations of  $\mathbb{S}^2$  projected to  $\mathbb{C}$ , which are the Möbius transformations of the form

$$\phi_a^\theta(z) = \frac{z-a}{1+\bar{a}z} e^{i\theta}, \quad z, a \in \mathbb{C} \text{ and } \theta \in [0, 2\pi).$$

As notation,  $\phi_a \equiv \phi_a^0$ , for all  $a \in \mathbb{C}$ .

**Remark 2.4.4.** From the expression of  $\phi_a$  it follows

$$1 + |\phi_a(z)|^2 = \frac{(1+|a|^2)(1+|z|^2)}{|1+\bar{a}z|^2}, \quad a, z \in \mathbb{C}. \quad (2.4.5)$$

Due to (2.4.4) and (2.4.5), it can be easily checked that

$$(1 + |\phi_a(z)|^2)^{-1} = 1 - \left( \frac{d_{ch}(z, a)}{2} \right)^2, \quad a, z \in \mathbb{C}. \quad (2.4.6)$$

Also notice that

$$\phi_a'(z) = \frac{1+|a|^2}{(1+\bar{a}z)^2}. \quad (2.4.7)$$

**Corollary 2.4.5.** *The parabolic measure is invariant by  $\phi_a$ , for all  $a \in \mathbb{C}$ .*

*Proof.* Recalling that  $dm^*$  is the surface area measure of  $\mathbb{S}^2$  projected to  $\mathbb{C}$ , we have, by using (2.4.5) and (2.4.7), that for all  $z \in \mathbb{C}$ :

$$dm^*(\phi_a(z)) = \frac{|\phi_a'(z)|^2}{\pi(1+|\phi_a(z)|^2)^2} dm(z) = \frac{dm(z)}{\pi(1+|z|^2)^2} = dm^*(z).$$

□

### 2.4.3 The spaces $\mathcal{P}_L$

We represent the Hilbert spaces of holomorphic functions used in the definition of GAF in the  $\mathbb{C}$ -chart. Given  $L \in \mathbb{N}$  we define the *parabolic space* or the *space of polynomials of degree at most  $L$*  as:

$$\mathcal{P}_L := \left\{ f \in P_L[\mathbb{C}] : \|f\|_{\mathcal{P}_L}^2 = (L+1) \int_{\mathbb{C}} \frac{|f(z)|^2}{(1+|z|^2)^L} dm^*(z) < +\infty \right\},$$

where  $P_L[\mathbb{C}]$  is the vector space of polynomials of degree at most  $L$  with complex coefficients and  $dm$  stands for the Lebesgue measure on  $\mathbb{C}$ . The space  $P_L[\mathbb{C}]$  can be seen as the projection to  $\mathbb{C}$  of the space of sections of the  $L$ -th power of the canonical bundle of  $\mathbb{S}^2$ .

It is not strange to consider such a norm. We normalize  $|f(z)|^2$  by  $(1 + |z|^2)^L$ , which is rescaled by the degree of  $f$ , to avoid that the first terms tends to infinite. The factor  $L + 1$  is chosen so that

$$\frac{L + 1}{(1 + |z|^2)^L} dm^*(z), \quad z \in \mathbb{C},$$

is a probability measure, i.e.,  $\|1\|_{\mathcal{P}_L} = 1$ .

**Lemma 2.4.6.** *The family*

$$e_n(z) = \binom{L}{n}^{1/2} z^n, \quad n = 0, \dots, L,$$

*forms an orthonormal basis of  $\mathcal{P}_L$ .*

*Proof.* We have to show that  $(e_n, e_m)_{\mathcal{P}_L} = \delta_{n,m}$  and that  $(e_n)_{n=0}^{+\infty}$  is complete. For  $n, m \in \mathbb{N}$  we have, by applying a change of variables in polar coordinates:

$$\begin{aligned} (e_n, e_m)_{\mathcal{P}_L} &= \frac{L + 1}{\pi} \int_{\mathbb{C}} \binom{L}{n}^{1/2} \binom{L}{m}^{1/2} \frac{z^n \bar{z}^m}{(1 + |z|^2)^{L+2}} dm(z) \\ &= \frac{L + 1}{\pi} \binom{L}{n}^{1/2} \binom{L}{m}^{1/2} \int_0^{2\pi} \int_0^{+\infty} \frac{r^{n+m+1} e^{i\theta(n-m)}}{(1 + r^2)^{L+2}} dr d\theta. \end{aligned}$$

Thus, if  $n \neq m$ ,  $(e_n, e_m)_{\mathcal{P}_L} = 0$  by (2.4.1). Otherwise, if  $n = m$ , we get, by using the property (2.3.1), the definition of the Beta function (2.3.3) and the change of variable  $t = r^2$ , that

$$\begin{aligned} (e_n, e_n)_{\mathcal{P}_L} &= \frac{(L + 1)\Gamma(L + 1)}{\Gamma(n + 1)\Gamma(L - n + 1)} \int_0^{+\infty} \frac{2r^{2n+1}}{(1 + r^2)^{L+2}} dr \\ &= \frac{(L + 1)\Gamma(L + 1)}{\Gamma(n + 1)\Gamma(L - n + 1)} \int_0^{+\infty} \frac{t^n}{t^{L+2}} dt \\ &= \frac{(L + 1)\Gamma(L + 1)}{\Gamma(n + 1)\Gamma(L - n + 1)} \frac{\Gamma(n + 1)\Gamma(L - n + 1)}{\Gamma(L + 2)} \\ &= \frac{(L + 1)\Gamma(L + 1)}{\Gamma(L + 2)} = 1. \end{aligned}$$

Hence  $(e_n)_{n=0}^L$  is an orthonormal system in  $\mathcal{P}_L$ .

Completeness follows since  $(z^n)_{n=0}^L$  is an orthogonal basis and  $(e_n)_{n=0}^L$  is its orthonormalization.  $\square$

By using Definition 2.2.1, for all  $L \in \mathbb{N}$  we can define the GAF

$$f_L(z) = \sum_{n=0}^L \xi_n \binom{L}{n}^{1/2} z^n, \quad (2.4.8)$$

where  $(\xi_n)_{n=0}^L$  is a sequence of i.i.d. random variables  $N_{\mathbb{C}}(0, 1)$ . This is known as the *parabolic Gaussian analytic function* ( $\mathbb{S}^2$ -GAF, for short). The covariance kernel of this GAF is

$$\mathcal{K}_{f_L}(z, w) = \sum_{n=0}^L \binom{L}{n} z^n \bar{w}^n = (1 + z\bar{w})^L. \quad (2.4.9)$$

## 2.5 Distribution, intensity and invariance of the zero set of a GAF

In this section we would like to see how these zero point sets are distributed. Assume that  $f$  is a GAF in a region  $\Omega \subseteq \mathbb{C}$  and denote  $\mathcal{Z}_f$  its zero set.

**Definition 2.5.1.** The *empirical measure* of  $f$  is

$$\nu_f = \sum_{a \in \mathcal{Z}_f} \delta_a = \frac{1}{2\pi} \Delta \log |f|,$$

where  $\delta_a$  is the Dirac delta measure at  $a$  and the Laplacian  $\Delta$  must be understood in the distributional sense.

The empirical measure encodes all the information of  $\mathcal{Z}_f$ , and allows the use of the tools of the theory of distributions to extract information from it.

**Definition 2.5.2.** Let  $\mu$  be measure that is finite over compact sets of  $\Omega$ . We say that  $u \in L_{\text{loc}}^1(\Omega)$  is a *solution* of

$$\Delta u = \mu$$

on  $\Omega$  in the sense of distributions if for all  $\varphi \in C_c^2(\Omega)$  it is satisfied

$$\int_{\Omega} u(z) \Delta \varphi(z) dm(z) = \int_{\Omega} \varphi(z) d\mu(z).$$

Now we are ready to state a fundamental result describing the average distribution of  $\mathcal{Z}_f$ .

**Theorem 2.5.3. (The Edelman-Kostlan formula)** (see [2], p. 24, 25) *Assume that  $f$  is a GAF in  $\Omega$  with zero mean and covariance kernel  $\mathcal{K}_f(z, w)$ , for all  $z, w \in \Omega$ . Then*

$$\mathbb{E}[\nu_f] = \frac{1}{2\pi} \log \mathbb{E}[|f(z)|] = \frac{1}{4\pi} \Delta \log \mathcal{K}_f(z, z),$$

where  $\Delta$  must be understood in the sense of distributions and  $\mathbb{E}[\nu_f]$  is a deterministic measure called the first intensity.

*Proof.* Consider a function  $\varphi \in \mathcal{C}_c^2(\Omega)$ . By Definition 2.5.1 we get

$$\int_{\Omega} \varphi(z) d\nu_f(z) = \int_{\Omega} \frac{1}{2\pi} \log |f(z)| \Delta\varphi(z) dm(z),$$

which implies

$$\mathbb{E} \left[ \int_{\Omega} \varphi(z) d\nu_f(z) \right] = \mathbb{E} \left[ \int_{\Omega} \frac{1}{2\pi} \log |f(z)| \Delta\varphi(z) dm(z) \right].$$

In order to apply Fubini's theorem on the RHS of the equation we must verify that:

$$\mathbb{E} \left[ \int_{\Omega} \left| \frac{1}{2\pi} \log |f(z)| \Delta\varphi(z) dm(z) \right| \right] < \infty.$$

Using the linearity of the expectation and taking into account that  $\Delta\varphi$  is deterministic, we have

$$\mathbb{E} \left[ \int_{\Omega} \left| \frac{1}{2\pi} \log |f(z)| \Delta\varphi(z) dm(z) \right| \right] = \int_{\Omega} \frac{1}{2\pi} \mathbb{E} [|\log |f(z)||] |\Delta\varphi(z)| dm(z).$$

By denoting  $\hat{f}(z) = f(z)/\sqrt{\mathcal{K}_f(z, z)}$ , which is a  $N_{\mathbb{C}}(0, 1)$  random variable, we have:

$$\begin{aligned} \mathbb{E} [|\log |f(z)||] &= \mathbb{E} \left[ \left| \log \left| \frac{f(z)}{\sqrt{\mathcal{K}_f(z, z)}} \sqrt{\mathcal{K}_f(z, z)} \right| \right| \right] \\ &= \mathbb{E} [|\log |\hat{f}(z)||] + \log \left| \sqrt{\mathcal{K}_f(z, z)} \right| \\ &\stackrel{(*)}{=} \int_{\mathbb{C}} |\log |\xi|| \frac{e^{-|\xi|^2}}{\pi} dm(\xi) + \frac{1}{2} \log |\mathcal{K}_f(z, z)| \\ &\stackrel{(**)}{=} \int_0^{+\infty} 2r |\log(r)| e^{-r^2} dr + \frac{1}{2} \log |\mathcal{K}_f(z, z)| \\ &\stackrel{(***)}{=} \int_0^{+\infty} |\log(s)| e^{-s} ds + \frac{1}{2} \log |\mathcal{K}_f(z, z)| = K_1 + \frac{1}{2} \log |\mathcal{K}_f(z, z)|, \end{aligned}$$

where in (\*) we applied the definition of expectation, taking into account that the density function of a  $N_{\mathbb{C}}(0, 1)$  random variable is  $e^{-|z|^2}/\pi$ , for all  $z \in \mathbb{C}$ . The step (\*\*) follows by a change into polar coordinates and (\*\*\*) by the change of variable  $s = r^2$ . Recall that  $K_1$  is a constant value.

Notice that  $\log |\mathcal{K}_f(z, z)|$  is locally integrable for all  $z \in \Omega$ . However we must study the case when  $\mathcal{K}_f(a, a) = 0$ , for  $a \in \Omega$ . For this we will show that:

$$\mathcal{K}_f(z, z) = |z - a|^{2m} G(z, z),$$

where  $m$  is a natural number and  $G$  is a function such that  $G(a, a) \neq 0$ , for those values  $a \in \Omega$  such that  $\mathcal{K}_f(a, a) = 0$ .

Indeed, by Definition 2.2.1, we can write

$$\mathcal{K}_f(z, z) = \sum_{n=0}^{+\infty} e_n(z) \overline{e_n(z)}.$$

Each element of  $(e_n)_{n=0}^{+\infty}$  is of the form

$$e_n(z) = (z - a)^{m_n} g_n(z),$$

where  $m_n$  is the multiplicity of  $a$  and  $g_n$  is a function such that  $g_n(a) \neq 0$ . Thus, and by denoting  $m = \min_{n \in \mathbb{N}} m_n$ :

$$\begin{aligned} \mathcal{K}_f(z, z) &= \sum_{n=0}^{+\infty} |z - a|^{2m_n} |g_n(z)|^2 = |z - a|^{2m} \sum_{n=0}^{+\infty} |z - a|^{2m_n - 2m} |g_n(z)|^2 \\ &= |z - a|^{2m} G(z, z), \end{aligned}$$

where  $G(a, a) \neq 0$  since  $g_n(a) \neq 0$ . Hence  $\log |\mathcal{K}_f(z, z)|$  is locally integrable in a neighbourhood of  $a$ . Therefore

$$\mathbb{E} \left[ \int_{\Omega} \left| \frac{1}{2\pi} \log |f(z)| \Delta \varphi(z) dm(z) \right| \right] < \infty,$$

and we are under conditions of Fubini's theorem. We get:

$$\mathbb{E} \left[ \int_{\Omega} \frac{1}{2\pi} \log |f(z)| \Delta \varphi(z) dm(z) \right] = \int_{\Omega} \frac{1}{2\pi} \Delta \mathbb{E} [\log |f(z)|] \varphi(z) dm(z).$$

Proceeding in a similar way as before we have

$$\mathbb{E} [\log |f(z)|] = K_2 + \frac{1}{2} \log \mathcal{K}_f(z, z),$$

where  $K_2$  is a constant value. Finally, we reach the equality

$$\mathbb{E} \left[ \int_{\Omega} \varphi(z) d\nu_f(z) \right] = \int_{\Omega} \frac{1}{4\pi} \Delta \log \mathcal{K}_f(z, z) \varphi(z) dm(z).$$

By Fubini's theorem and recalling that  $\varphi$  is deterministic, we get that

$$\mathbb{E} \left[ \int_{\Omega} \varphi(z) d\nu_f(z) \right] = \int_{\Omega} \varphi(z) d\mathbb{E} [\nu_f].$$

By Definition 2.5.2 we conclude that the first intensity is, with respect to the Lebesgue measure:

$$\mathbb{E} [\nu_f] = \frac{1}{4\pi} \Delta \log \mathcal{K}_f(z, z).$$

□

### 2.5.1 First intensity of a $\mathbb{C}$ -GAF

We see here that the average number of points of the  $\mathbb{C}$ -GAF of intensity  $L > 0$  is  $L$  times the Lebesgue measure in  $\mathbb{C}$ .

Let  $f_L$  be a  $\mathbb{C}$ -GAF with real parameter  $L > 0$ . From its covariance kernel expression, (2.4.3), we have

$$\mathbb{E}[\nu_{f_L}] = \frac{1}{4\pi} \Delta \log \mathcal{K}_{f_L}(z, z) = \frac{1}{4\pi} \Delta \log (e^{Lz\bar{w}}) = \frac{L}{\pi} \frac{\partial}{\partial z} \frac{\partial}{\partial \bar{z}} (z\bar{z}) = \frac{L}{\pi} dm(z).$$

The zero set process of a  $\mathbb{C}$ -GAF has another property: it is invariant by translations in  $\mathbb{C}$ .

**Proposition 2.5.4.** *Let  $f_L$  be a  $\mathbb{C}$ -GAF of real parameter  $L > 0$ . Its zero point process is invariant under the transformations*

$$\phi_a(z) = z - a, \quad z, a \in \mathbb{C}.$$

*Proof.* Denote

$$f_a(z) = f_L(\phi_a(z)).$$

This function has covariance kernel

$$\mathcal{K}_{f_a}(z, w) = \mathcal{K}_{f_L}(\phi_a(z), \phi_a(w)) = e^{Lz\bar{w} - Lz\bar{a} - La\bar{w} + L|a|^2}. \quad (2.5.1)$$

Consider the function

$$T_a f_L(z) = f_a(z) e^{Lz\bar{a} - \frac{L}{2}|a|^2}.$$

We want to show that

$$f_L(z) = T_a f_L(z)$$

in distribution. To check this we must ensure that their covariance kernels are equal. Indeed:

$$\mathcal{K}_{T_a f_L}(z, w) = \mathcal{K}_{f_a}(z, w) e^{Lz\bar{a} - L|a|^2 + La\bar{w}} \stackrel{(*)}{=} e^{Lz\bar{w}} = \mathcal{K}_{f_L}(z, w),$$

where in (\*) we used (2.5.1). Since the covariance kernels coincide, the proof is finished.  $\square$

**Proposition 2.5.5.** *Let  $f_L$  be a  $\mathbb{C}$ -GAF of real parameter  $L > 0$ . Then  $f_L$  and  $T_a f_L$  are isometric.*

*Proof.* We must show that

$$\|f_L\|_{\mathcal{F}_L}^2 = \|T_a f_L\|_{\mathcal{F}_L}^2.$$

Noticing that

$$\|T_a f_L\|_{\mathcal{F}_L}^2 = \frac{L}{\pi} \int_{\mathbb{C}} |f_L(z - a)|^2 \left| e^{L\bar{a}z - \frac{L}{2}|a|^2} \right|^2 e^{-L|z|^2} dm(z),$$

doing the change of variable  $w = z - a$  and recalling that the Lebesgue measure is invariant under translations, the equality is simple to verify.  $\square$

Proposition 2.5.4 implies that the zero set of  $f_a$  is the same one in distribution than  $f_L$ . Indeed, since  $e^{Lz\bar{a} - \frac{L}{2}|a|^2}$  does not vanish anywhere, the property is trivially accomplished.

Since the first intensity of a  $\mathbb{C}$ -GAF determines in mean the distribution of its zero set process, it is also invariant by translations in  $\mathbb{C}$ .

### 2.5.2 First intensity of an $\mathbb{S}^2$ -GAF

We see here that the average number of points of the  $\mathbb{S}^2$ -GAF in a region is proportional to the area of the region at  $\mathbb{S}^2$ .

Let  $f_L$  be an  $\mathbb{S}^2$ -GAF of parameter  $L \in \mathbb{N}$ . From its covariance kernel, (2.4.9), we have

$$\mathbb{E}[\nu_{f_L}] = \frac{1}{4\pi} \Delta \log \mathcal{K}_{f_L}(z, z) = \frac{1}{4\pi} \Delta \log [(1 + |z|^2)^L] = \frac{L}{\pi} \frac{dm(z)}{(1 + |z|^2)^2} = L dm^*(z).$$

Notice that the Poisson process and the zero set process of an  $\mathbb{S}^2$ -GAF have the same average number of points (see (1.3.2)).

The zero set process of an  $\mathbb{S}^2$ -GAF is invariant by rotations in  $\mathbb{S}^2$ , which in the  $\mathbb{C}$ -chart are seen as certain Möbius transformations.

**Proposition 2.5.6.** *Let  $f_L$  be an  $\mathbb{S}^2$ -GAF. Its zero point process is invariant under the Möbius transformations*

$$\phi_a(z) = \frac{z - a}{1 + \bar{a}z}, \quad a, z \in \mathbb{C}. \quad (2.5.2)$$

*Proof.* As in the proof of Proposition 2.5.4, consider

$$f_a(z) = f_L(\phi_a(z)).$$

Its covariance kernel is

$$\mathcal{K}_{f_a}(z, w) = \mathcal{K}_{f_L}(\phi_a(z), \phi_a(w)) = \left( \frac{(1 + |a|^2)(1 + z\bar{w})}{(1 + \bar{a}z)(1 + a\bar{w})} \right)^L. \quad (2.5.3)$$

Consider

$$T_a f_L(z) = \left( \frac{1 + |a|^2}{(1 + \bar{a}z)^2} \right)^{-L/2} f_a(z).$$

What we want to prove is

$$f_L(z) = T_a f_L(z)$$

in distribution. To see this we must check that their covariance kernels coincide. We have:

$$\begin{aligned} \mathcal{K}_{T_a f_L}(z, w) &= \mathcal{K}_{f_a}(z, w) \left( \frac{1 + |a|^2}{(1 + \bar{a}z)^2} \right)^{-L/2} \overline{\left( \frac{1 + |a|^2}{(1 + \bar{a}w)^2} \right)^{-L/2}} \\ &\stackrel{(*)}{=} (1 + z\bar{w})^L = \mathcal{K}_{f_L}(z, w), \end{aligned}$$

where in (\*) we applied (2.5.3). Since the covariance kernels are equal, we finally have the result we desired.  $\square$

**Proposition 2.5.7.** *Let  $f_L$  be an  $\mathbb{S}^2$ -GAF of parameter  $L \in \mathbb{N}$ . We have that  $T_a$  is an isometry from  $\mathcal{P}_L$  to  $\mathcal{P}_L$ .*

*Proof.* The statement is equivalent to proving that

$$\|f_L\|_{\mathcal{P}_L}^2 = \|T_a f_L\|_{\mathcal{P}_L}^2.$$

Using (2.4.7) we have that

$$(1 + |z|^2) |\phi'_a(z)| = \frac{(1 + |a|^2)(1 + |z|^2)}{|1 + \bar{a}z|^2} = (1 + |\phi_a(z)|)^{-1}.$$

Hence, by doing the change of variable  $w = \phi_a(z)$  (see (2.5.2)), using (2.4.7) and applying Corollary 2.4.5, we get:

$$\begin{aligned} \|T_a f_L\|_{\mathcal{P}_L}^2 &= (L + 1) \int_{\mathbb{C}} \frac{|f_L(\phi_a(z))|^2}{(1 + |z|^2)^L} \frac{|1 + \bar{a}z|^{2L}}{(1 + |a|^2)^L} dm^*(z) \\ &= (L + 1) \int_{\mathbb{C}} \frac{|f_L(\phi_a(z))|^2}{(1 + |\phi_a(z)|^2)^L} dm^*(z) \\ &= (L + 1) \int_{\mathbb{C}} \frac{|f_L(w)|^2}{(1 + |w|^2)^L} dm^*(w) = \|f_L\|_{\mathcal{P}_L}^2. \end{aligned}$$

□

By Proposition 2.5.6 we get that the zero set point of an  $\mathbb{S}^2$ -GAF  $f_L$  is the same than  $T_a f_L$  in distribution. Furthermore, the zero set of  $f_L$  is equal to  $f_a$  in distribution. Indeed, if for all  $a, z \in \mathbb{C}$  we had

$$\left( \frac{1 + |a|^2}{(1 + \bar{a}z)^2} \right)^{-L/2} = 0,$$

we would conclude that  $|a|^2 = -1$ , which is a contradiction by the definition of modulus.

Since the first intensity of an  $\mathbb{S}^2$ -GAF determines in mean the distribution of its zero set process, it is also invariant by the Möbius rotational transformation. Indeed, by the change of variable  $w = \phi_a(z)$  (see (2.5.2)), using (2.4.5) and taking  $f_a$  as in the proof of Proposition 2.5.6, we get

$$\begin{aligned} \mathbb{E}[\nu_{f_L}] &= \frac{L}{\pi} \frac{dm(z)}{(1 + |\phi_a(z)|^2)^2} = \frac{L}{\pi} \frac{(1 - a\bar{w})^2(1 - \bar{a}w)^2(1 + |a|^2)^2}{(1 - a\bar{w})^2(1 - \bar{a}w)^2(1 + |a|^2)^2(1 + |w|^2)^2} dm(w) \\ &= \frac{L}{\pi} \frac{dm(w)}{(1 + |w|^2)^2} = \mathbb{E}[\nu_{f_a}]. \end{aligned}$$



# Chapter 3

## Fluctuation of the zero set of an $\mathbb{S}^2$ -GAF

Here we address the main goal of this project: the study of the fluctuation of an  $\mathbb{S}^2$ -GAF in  $\Omega$ . This can be measured in different ways. One possibility is to measure how the number of zeros in a given region  $D \subset\subset \Omega$ , with  $\partial D$  regular, deviates from its mean (given by the Edelman-Kostlan formula). This deviation can be quantified as the variance of the random variables counting the number of points on  $D$ , i.e.,  $\nu_f(D)$ , where  $f$  is a GAF in  $\Omega$ .

We will first give a general formula for the variance of  $\nu_f(D)$ . Later we shall use this formula to write the particular case of the  $\mathbb{S}^2$ -GAF  $f_L$  of parameter  $L \in \mathbb{N}$  and  $D$  a disk in  $\mathbb{C}$ . From this explicit formula we will deduce the asymptotic behaviour of the variance of  $\nu_{f_L}(D)$  in different ways.

To write this chapter I followed a similar scheme as in [5]. The source [6] was essential for this chapter. The references [2] and [9] were also helpful for the understanding of some proofs.

### 3.1 Variance of a GAF

We are going to compute the variance of a GAF  $f$  defined in  $\Omega$  in a region  $D \subset\subset \Omega$  with  $\partial D$  regular. For simplicity, we are going to denote

$$\nu_f(D) = I(\mathbb{1}_D) = \int_D d\nu_f = \#(\mathcal{Z}_f \cap D).$$

In addition, we will use  $\mathbb{V}$  for the variance. We state the following theorem.

**Theorem 3.1.1.** *Let  $f$  be a GAF in  $\Omega$  and let  $D \subset\subset \Omega$  with  $\partial D$  regular. Then*

$$\mathbb{V}[\nu_f(D)] = -\frac{1}{4\pi^2} \int_{\partial D} \int_{\partial D} \frac{1}{1 - K(z, w)} \frac{\partial}{\partial \bar{z}} \left( \frac{\mathcal{K}_f(w, z)}{\mathcal{K}_f(z, z)} \right) \frac{\partial}{\partial \bar{w}} \left( \frac{\mathcal{K}_f(z, w)}{\mathcal{K}_f(w, w)} \right) d\bar{z} d\bar{w}, \quad (3.1.1)$$

where

$$K(z, w) = \frac{|\mathcal{K}_f(z, w)|^2}{\mathcal{K}_f(z, z)\mathcal{K}_f(w, w)}. \quad (3.1.2)$$

*Proof.* By definition

$$\mathbb{V} [\nu_f(D)] = \mathbb{E} [(\nu_f(D) - \mathbb{E} [\nu_f(D)])^2].$$

By Definition 2.5.1 and Theorem 2.5.3 we can state:

$$\begin{aligned} \nu_f(D) - \mathbb{E} [\nu_f(D)] &= \int_D \left( \frac{1}{2\pi} \Delta \log |f(z)| - \frac{1}{2\pi} \Delta \sqrt{\mathcal{K}_f(z, z)} \right) \\ &= \int_D \frac{1}{2\pi} \Delta \log \left( \frac{|f(z)|}{\sqrt{\mathcal{K}_f(z, z)}} \right) \\ &\stackrel{(*)}{=} \int_D \frac{1}{2\pi} \Delta \log |\hat{f}(z)|, \end{aligned}$$

where in (\*) we denote  $\hat{f} = f/\sqrt{\mathcal{K}_f(z, z)} \sim N_{\mathbb{C}}(0, 1)$ . To go further in the calculations consider the 1-form

$$\omega = -\frac{i}{\pi} \bar{\partial}_z \log |\hat{f}(z)| d\bar{z}.$$

Applying the exterior derivative to  $\omega$ , remembering that  $d^2 = 0$  and recalling the fact that  $dz \wedge d\bar{z} = -2idx \wedge dy$ , we get:

$$\begin{aligned} d\omega &= -\frac{i}{\pi} \frac{\partial^2}{\partial z \partial \bar{z}} \log |\hat{f}(z)| d\bar{z} \wedge dz = \frac{i}{4\pi} \Delta \log |\hat{f}(z)| dz \wedge d\bar{z} = \frac{1}{2\pi} \Delta \log |\hat{f}(z)| dx \wedge dy \\ &= \frac{1}{2\pi} \Delta \log |\hat{f}(z)| dm(z). \end{aligned}$$

By Stokes' theorem we conclude that

$$\int_D \frac{1}{2\pi} \Delta \log |\hat{f}(z)| = - \int_{\partial D} \frac{i}{\pi} \bar{\partial}_z \log |\hat{f}(z)| d\bar{z}.$$

So we have

$$\begin{aligned} \mathbb{V} [\nu_f(D)] &= \mathbb{E} [(\nu_f(D) - \mathbb{E} [\nu_f(D)])^2] \\ &= \mathbb{E} \left[ \int_{\partial D} \frac{i}{\pi} \bar{\partial}_z \log |\hat{f}(z)| d\bar{z} \int_{\partial D} \frac{i}{\pi} \bar{\partial}_w \log |\hat{f}(w)| d\bar{w} \right] \\ &= -\frac{1}{\pi^2} \int_{\partial D} \int_{\partial D} \bar{\partial}_z \bar{\partial}_w \mathbb{E} [\log |\hat{f}(z)| \log |\hat{f}(w)|] d\bar{z} d\bar{w}, \end{aligned}$$

where in the last equality we used Fubini's theorem and the differentiation under the integral sign. Recall that, if  $X$  and  $Y$  are two random variables, their *covariance* is defined as

$$\text{Cov} [X, Y] = \mathbb{E} [XY] - \mathbb{E} [X] \mathbb{E} [Y].$$

Here

$$\text{Cov} [\log |\hat{f}(z)|, \log |\hat{f}(w)|] = \mathbb{E} [\log |\hat{f}(z)| \log |\hat{f}(w)|] - \mathbb{E} [\log |\hat{f}(z)|] \mathbb{E} [\log |\hat{f}(w)|].$$

Since  $\hat{f} \sim N_{\mathbb{C}}(0, 1)$ ,  $\mathbb{E} \left[ \log |\hat{f}(z)| \right]$  is constant (independent of  $z$ ). Hence we have

$$\bar{\partial}_z \bar{\partial}_w \text{Cov} \left[ \log |\hat{f}(z)|, \log |\hat{f}(w)| \right] = \bar{\partial}_z \bar{\partial}_w \mathbb{E} \left[ \log |\hat{f}(z)| \log |\hat{f}(w)| \right]$$

and then we can write

$$\mathbb{V} [\nu_f(D)] = -\frac{1}{\pi^2} \int_{\partial D} \int_{\partial D} \bar{\partial}_z \bar{\partial}_w \text{Cov} \left[ \log |\hat{f}(z)|, \log |\hat{f}(w)| \right] d\bar{z} d\bar{w}.$$

Notice that

$$\mathbb{E} \left[ \hat{f}(z) \overline{\hat{f}(z)} \right] = \mathbb{E} \left[ |\hat{f}(z)|^2 \right] = \mathbb{E} \left[ \frac{|f(z)|^2}{\mathcal{K}_f(z, z)} \right] = \frac{\mathcal{K}_f(z, z)}{\mathcal{K}_f(z, z)} = 1.$$

Also we get

$$\Theta(z, w) := \mathbb{E} \left[ \hat{f}(z) \overline{\hat{f}(w)} \right] = \frac{\mathbb{E} \left[ \hat{f}(z) \overline{\hat{f}(w)} \right]}{\sqrt{\mathcal{K}_f(z, z)} \sqrt{\mathcal{K}_f(w, w)}} = \frac{\mathcal{K}_f(z, w)}{\sqrt{\mathcal{K}_f(z, z)} \sqrt{\mathcal{K}_f(w, w)}}.$$

**Lemma 3.1.2** ([2], p. 44-46). *Let  $Z_1$  and  $Z_2$  be complex normal random variables such that  $\mathbb{E} [Z_1 \overline{Z_1}] = \mathbb{E} [Z_2 \overline{Z_2}] = 1$  and  $\mathbb{E} [Z_1 \overline{Z_2}] = \theta$ , then*

$$\text{Cov} [\log |Z_1|, \log |Z_2|] = \sum_{j=1}^{+\infty} \frac{|\theta|^{2j}}{4j^2} = \frac{1}{4} \text{Li}_2 (|\theta|^2).$$

The function

$$\text{Li}_2(x) := \sum_{j=1}^{+\infty} \frac{x^j}{j^2}, \quad x \in [0, 1),$$

is called the dilogarithm.

Applying this lemma in our case:

$$\mathbb{V} [\nu_f(D)] = -\frac{1}{4\pi^2} \int_{\partial D} \int_{\partial D} \bar{\partial}_z \bar{\partial}_w \text{Li}_2 (|\Theta(z, w)|^2) d\bar{z} d\bar{w}.$$

We shall compute  $\bar{\partial}_z \bar{\partial}_w \text{Li}_2 (K(z, w))$ . For simplicity, denote (3.1.2) by  $K$ . We have

$$\bar{\partial}_w \text{Li}_2 (K) = \bar{\partial}_w \sum_{j=1}^{+\infty} \frac{K^j}{j^2} = \bar{\partial}_w (K) \sum_{j=1}^{+\infty} \frac{K^{j-1}}{j},$$

and

$$\bar{\partial}_z \bar{\partial}_w \text{Li}_2 (K) = \bar{\partial}_z \left( \bar{\partial}_w (K) \sum_{j=1}^{+\infty} \frac{K^{j-1}}{j} \right) \tag{3.1.3}$$

$$= \frac{\partial^2 K}{\partial \bar{z} \partial \bar{w}} \sum_{j=1}^{+\infty} \frac{K^{j-1}}{j} + \bar{\partial}_w (K) \bar{\partial}_z (K) \sum_{j=1}^{+\infty} \frac{j-1}{j} K^{j-2}. \tag{3.1.4}$$

We claim that:

$$\frac{\partial^2 K}{\partial \bar{z} \partial \bar{w}} = \frac{1}{K} \frac{\partial K}{\partial \bar{z}} \frac{\partial K}{\partial \bar{w}}.$$

On the one hand

$$\frac{\partial K}{\partial \bar{z}} = \frac{\partial}{\partial \bar{z}} \left( \frac{|\mathcal{K}_f(z, w)|^2}{\mathcal{K}_f(z, z) \mathcal{K}_f(w, w)} \right) = \frac{1}{\mathcal{K}_f(w, w)} \frac{\partial}{\partial \bar{z}} \left( \frac{\mathcal{K}_f(z, w) \overline{\mathcal{K}_f(w, z)}}{\mathcal{K}_f(z, z)} \right) \quad (3.1.5)$$

$$= \frac{\mathcal{K}_f(z, w)}{\mathcal{K}_f(w, w)} \frac{\partial}{\partial \bar{z}} \left( \frac{\mathcal{K}_f(w, z)}{\mathcal{K}_f(z, z)} \right). \quad (3.1.6)$$

On the other hand

$$\frac{\partial K}{\partial \bar{w}} = \frac{\partial}{\partial \bar{w}} \left( \frac{|\mathcal{K}_f(z, w)|^2}{\mathcal{K}_f(z, z) \mathcal{K}_f(w, w)} \right) = \frac{1}{\mathcal{K}_f(z, z)} \frac{\partial}{\partial \bar{w}} \left( \frac{\mathcal{K}_f(z, w) \overline{\mathcal{K}_f(w, z)}}{\mathcal{K}_f(w, w)} \right) \quad (3.1.7)$$

$$= \frac{\mathcal{K}_f(w, z)}{\mathcal{K}_f(z, z)} \frac{\partial}{\partial \bar{w}} \left( \frac{\mathcal{K}_f(z, w)}{\mathcal{K}_f(w, w)} \right). \quad (3.1.8)$$

All combined

$$\begin{aligned} \frac{1}{K} \frac{\partial K}{\partial \bar{z}} \frac{\partial K}{\partial \bar{w}} &= \frac{1}{K} \frac{\mathcal{K}_f(z, w) \mathcal{K}_f(w, z)}{\mathcal{K}_f(w, w) \mathcal{K}_f(z, z)} \frac{\partial}{\partial \bar{z}} \left( \frac{\mathcal{K}_f(w, z)}{\mathcal{K}_f(z, z)} \right) \frac{\partial}{\partial \bar{w}} \left( \frac{\mathcal{K}_f(z, w)}{\mathcal{K}_f(w, w)} \right) \\ &= \frac{\partial^2 K}{\partial \bar{z} \partial \bar{w}}. \end{aligned}$$

Now we can use this identity to finish the computations of  $\bar{\partial}_z \bar{\partial}_w \text{Li}_2(K)$ . Going back to (3.1.3), we have

$$\begin{aligned} \bar{\partial}_z \bar{\partial}_w \text{Li}_2(K) &= \frac{\partial^2 K}{\partial \bar{z} \partial \bar{w}} \sum_{j=1}^{+\infty} \frac{K^{j-1}}{j} + \bar{\partial}_w(K) \bar{\partial}_z(K) \sum_{j=1}^{+\infty} \frac{j-1}{j} K^{j-2} \\ &= \frac{1}{K} \frac{\partial K}{\partial \bar{z}} \frac{\partial K}{\partial \bar{w}} \sum_{j=1}^{+\infty} \frac{K^{j-1}}{j} + \frac{\partial K}{\partial \bar{z}} \frac{\partial K}{\partial \bar{w}} \sum_{j=1}^{+\infty} \frac{j-1}{j} K^{j-2} \\ &= \frac{1}{K} \frac{\partial K}{\partial \bar{z}} \frac{\partial K}{\partial \bar{w}} \left( \sum_{j=1}^{+\infty} K^{j-1} \right) = \frac{1}{K(1-K)} \frac{\partial K}{\partial \bar{z}} \frac{\partial K}{\partial \bar{w}}, \end{aligned}$$

where the last equality follows because  $|K| \leq 1$  (see Lemma 2.2.3). Hence

$$\mathbb{V}[\nu_f(D)] = -\frac{1}{4\pi^2} \int_{\partial D} \int_{\partial D} \frac{1}{K(1-K)} \frac{\partial K}{\partial \bar{z}} \frac{\partial K}{\partial \bar{w}} d\bar{z} d\bar{w}.$$

Using (3.1.5) and (3.1.7), we finish with the expression

$$\mathbb{V}[\nu_f(D)] = -\frac{1}{4\pi^2} \int_{\partial D} \int_{\partial D} \frac{1}{1-K} \frac{\partial}{\partial \bar{z}} \left( \frac{\mathcal{K}_f(w, z)}{\mathcal{K}_f(z, z)} \right) \frac{\partial}{\partial \bar{w}} \left( \frac{\mathcal{K}_f(z, w)}{\mathcal{K}_f(w, w)} \right) d\bar{z} d\bar{w},$$

as we wanted.  $\square$

## 3.2 Fluctuation of the zero set of an $\mathbb{S}^2$ -GAF

In this section we will use Theorem 3.1.1 to get expressions for disks of the variance for a  $\mathbb{C}$ -GAF with real parameter  $L > 0$  and for a  $\mathbb{S}^2$ -GAF with parameter  $L \in \mathbb{N}$ . We will compare these values with the analogous ones for a Poisson process with the same average of points. Also we will show that the Bargmann-Fock space of parameter  $L = 1$  can be seen as the limit of parabolic spaces as  $L \rightarrow +\infty$ .

### 3.2.1 Variance of a $\mathbb{C}$ -GAF

The starting point of our analysis is the explicit formula of the variance for a  $\mathbb{C}$ -GAF.

**Theorem 3.2.1.** *Let  $f_L$  be a  $\mathbb{C}$ -GAF of real parameter  $L > 0$ . For a disk  $D(z_0, r) \subset \mathbb{C}$  for  $z_0 \in \mathbb{C}$  and  $r > 0$  we have:*

$$\mathbb{V}[\nu_{f_L}(D(z_0, r))] = \frac{\sqrt{Lr^2}}{2\pi} \int_0^{4Lr^2} \frac{1}{e^x - 1} \frac{\sqrt{x}}{\sqrt{1 - \frac{x}{4Lr^2}}} dx. \quad (3.2.1)$$

*Proof.* The value does not depend on  $z_0$  by the invariance of the translations. Then we can assume that  $z_0 = 0$ . Denote  $D := D(0, r)$ .

We apply Theorem 3.1.1 to this case. We know that  $\mathcal{K}_{f_L}(z, w) = e^{Lz\bar{w}}$ . Then, by (3.1.2),  $K(z, w) = e^{-L|z-w|^2}$ . It is trivial to compute the following:

$$\frac{\mathcal{K}_{f_L}(z, w)}{\mathcal{K}_{f_L}(w, w)} = e^{Lz\bar{w} - L|w|^2}, \quad \frac{\partial}{\partial \bar{w}} \left( \frac{\mathcal{K}_{f_L}(z, w)}{\mathcal{K}_{f_L}(w, w)} \right) = L(z - w)e^{Lz\bar{w} - L|w|^2},$$

$$\frac{\mathcal{K}_{f_L}(w, z)}{\mathcal{K}_{f_L}(z, z)} = e^{Lw\bar{z} - L|z|^2}, \quad \frac{\partial}{\partial \bar{z}} \left( \frac{\mathcal{K}_{f_L}(w, z)}{\mathcal{K}_{f_L}(z, z)} \right) = L(w - z)e^{Lw\bar{z} - L|z|^2},$$

$$\frac{1}{1 - K(z, w)} = \frac{1}{1 - e^{-L|z-w|^2}}.$$

Hence

$$\mathbb{V}[\nu_f(D)] = \frac{L^2}{4\pi^2} \int_{\partial D} \int_{\partial D} \frac{(w - z)^2}{e^{L|z-w|^2} - 1} d\bar{z}d\bar{w}.$$

Denoting  $z = re^{i\theta}$ ,  $w = re^{i\phi}$ , for all  $\theta, \phi \in (0, 2\pi)$ , we have

$$\begin{aligned} (w - z)^2 |\mathcal{J}(\theta, \phi)| &= r^2(e^{2i\theta} - 2e^{i(\theta+\phi)} + e^{2i\phi}) \begin{vmatrix} -ire^{-i\theta} & 0 \\ 0 & -ire^{-i\phi} \end{vmatrix} \\ &= -r^4(e^{i(\theta-\phi)} + e^{i(\phi-\theta)} - 2) = r^4 |1 - e^{i(\theta-\phi)}|^2, \end{aligned}$$

where  $\mathcal{J}$  is the Jacobian matrix. Applying the change of variables with respect to  $(\theta, \phi)$ , we get

$$\mathbb{V}[\nu_{f_L}(D)] = \frac{Lr^2}{4\pi^2} \int_0^{2\pi} \int_0^{2\pi} \frac{Lr^2 |1 - e^{i(\theta-\phi)}|^2}{e^{Lr^2 |1 - e^{i(\theta-\phi)}|^2} - 1} d\theta d\phi \stackrel{(*)}{=} \frac{Lr^2}{2\pi} \int_0^{2\pi} \frac{Lr^2 |1 - e^{it}|^2}{e^{Lr^2 |1 - e^{it}|^2} - 1} dt,$$

where in (\*) we applied the change  $t = \theta - \phi$ . Now, notice that  $|1 - e^{it}|^2 = 2 - 2 \cos t$ . Using the change of variable

$$x = 2Lr^2(1 - \cos t), \quad dx = 2Lr^2 \sin t dt,$$

where

$$\sin t = \sqrt{1 - \cos^2 t} = \sqrt{\frac{x}{Lr^2} \sqrt{1 - \frac{x}{4Lr^2}}},$$

we have, by using the fact that the integrand is even,

$$\mathbb{V}[\nu_{f_L}(D)] = \frac{\sqrt{Lr^2}}{2\pi} \int_0^{4Lr^2} \frac{1}{e^x - 1} \frac{\sqrt{x}}{\sqrt{1 - \frac{x}{4Lr^2}}} dx.$$

□

### 3.2.2 Variance of an $\mathbb{S}^2$ -GAF

Let us give the variance expression for an  $\mathbb{S}^2$ -GAF.

**Theorem 3.2.2.** *Let  $f_L$  be an  $\mathbb{S}^2$ -GAF of parameter  $L \in \mathbb{N}$ . For a chordal disk  $D_{ch}(z_0, \tilde{\rho}) \subset \mathbb{C}$  for  $z_0 \in \mathbb{C}$  and  $\tilde{\rho} > 0$  we have:*

$$\mathbb{V}[\nu_{f_L}(D_{ch}(z_0, \tilde{\rho}))] = \frac{L^2 \tilde{\rho} \sqrt{4 - \tilde{\rho}^2}}{2\pi \cdot 4} \int_0^{\frac{\tilde{\rho}^2(4 - \tilde{\rho}^2)}{4}} \frac{(1-x)^L}{1 - (1-x)^L} \frac{\sqrt{x}}{1-x} \frac{1}{\sqrt{1 - \frac{4}{\tilde{\rho}^2(4 - \tilde{\rho}^2)}x}} dx, \quad (3.2.2)$$

*Proof.* The value does not depend on  $z_0$  by the invariance of the rotations. Then we can assume  $z_0 = 0$ . By i) of Remark 2.4.3, we have that, if  $z \in D_{ch}(0, \tilde{\rho})$ :

$$|z| \leq \frac{\tilde{\rho}}{\sqrt{4 - \tilde{\rho}^2}} = \hat{\rho}.$$

We shall apply Theorem 3.1.1 to the case in  $D := D(0, \hat{\rho})$ . Since  $f_L$  is an  $\mathbb{S}^2$ -GAF, it is easy to verify that

$$K(z, w) = \frac{|1 + z\bar{w}|^{2L}}{(1 + |z|^2)^L (1 + |w|^2)^L}.$$

By straightforward computations we get

$$\frac{\mathcal{K}_{f_L}(z, w)}{\mathcal{K}_{f_L}(w, w)} = \frac{(1 + z\bar{w})^L}{(1 + |w|^2)^L}, \quad \frac{\partial}{\partial \bar{w}} \left( \frac{\mathcal{K}_{f_L}(z, w)}{\mathcal{K}_{f_L}(w, w)} \right) = L(z - w) \frac{(1 + z\bar{w})^{L-1}}{(1 + |w|^2)^{L+1}},$$

$$\frac{\mathcal{K}_{f_L}(w, z)}{\mathcal{K}_{f_L}(z, z)} = \frac{(1 + w\bar{z})^L}{(1 + |z|^2)^L}, \quad \frac{\partial}{\partial \bar{z}} \left( \frac{\mathcal{K}_{f_L}(w, z)}{\mathcal{K}_{f_L}(z, z)} \right) = -L(z - w) \frac{(1 + w\bar{z})^{L-1}}{(1 + |z|^2)^{L+1}},$$

$$\frac{1}{1 - K(z, w)} = \frac{(1 + |z|^2)^L (1 + |w|^2)^L}{(1 + |z|^2)^L (1 + |w|^2)^L - |1 + z\bar{w}|^{2L}}.$$

Thus

$$\begin{aligned} \mathbb{V}[\nu_{f_L}(D_{ch}(0, \tilde{\rho}))] &= \\ &= \frac{L^2}{4\pi^2} \int_{\partial D} \int_{\partial D} \frac{|1 + z\bar{w}|^{2L-2} (z-w)^2 d\bar{z}d\bar{w}}{(1+|z|^2)^{L+1}(1+|w|^2)^{L+1} - |1+z\bar{w}|^{2L}(1+|z|^2)(1+|w|^2)}. \end{aligned}$$

Fixing  $\hat{\rho} > 0$  of the disk  $D$  and denoting  $z = \hat{\rho}e^{i\theta}$ ,  $w = \hat{\rho}e^{i\phi}$ , for all  $\theta, \phi \in (0, 2\pi)$ , we have

$$\begin{aligned} (z-w)^2 |\mathcal{J}(\theta, \phi)| &= \hat{\rho}^2 (e^{2i\theta} - 2e^{i(\theta+\phi)} + e^{2i\phi}) \begin{vmatrix} -i\hat{\rho}e^{-i\theta} & 0 \\ 0 & -i\hat{\rho}e^{-i\phi} \end{vmatrix} \\ &= -\hat{\rho}^4 (e^{i(\theta-\phi)} + e^{i(\phi-\theta)} - 2) = \hat{\rho}^4 |1 - e^{i(\theta-\phi)}|^2, \end{aligned}$$

where  $\mathcal{J}$  stands for the Jacobian matrix. By trivial calculations and applying the change of variable  $t = \theta - \phi$ , we get now

$$\begin{aligned} \mathbb{V}[\nu_{f_L}(D_{ch}(0, \tilde{\rho}))] &= \frac{L^2 \hat{\rho}^4}{4\pi^2} \int_0^{2\pi} \int_0^{2\pi} \frac{|1 + \hat{\rho}^2 e^{i(\theta-\phi)}|^{2L-2} |1 - e^{i(\theta-\phi)}|^2}{(1 + \hat{\rho}^2)^{2L+2} - |1 + \hat{\rho}^2 e^{i(\theta-\phi)}|^{2L} (1 + \hat{\rho}^2)^2} d\theta d\phi \\ &= \frac{L^2}{2\pi} \frac{\hat{\rho}^4}{(1 + \hat{\rho}^2)^2} \int_0^{2\pi} \frac{|1 + \hat{\rho}^2 e^{it}|^{2L}}{(1 + \hat{\rho}^2)^{2L} - |1 + \hat{\rho}^2 e^{it}|^{2L}} \frac{|1 - e^{it}|^2}{|1 + \hat{\rho}^2 e^{it}|^2} dt. \end{aligned}$$

Notice that

$$|1 + \hat{\rho}^2 e^{it}|^2 = (1 + \hat{\rho}^2)^2 \left[ 1 - \frac{2\hat{\rho}^2}{(1 + \hat{\rho}^2)^2} (1 - \cos t) \right], \quad |1 - e^{it}|^2 = 2 - 2\cos t.$$

By using these expressions and that the integrand is even we have

$$\begin{aligned} \mathbb{V}[\nu_{f_L}(D_{ch}(0, \tilde{\rho}))] &= \\ &= \frac{2L^2}{\pi} \frac{\hat{\rho}^4}{(1 + \hat{\rho}^2)^4} \int_0^\pi \frac{\left[ 1 - \frac{2\hat{\rho}^2}{(1 + \hat{\rho}^2)^2} (1 - \cos t) \right]^L}{1 - \left[ 1 - \frac{2\hat{\rho}^2}{(1 + \hat{\rho}^2)^2} (1 - \cos t) \right]^L} \frac{1 - \cos t}{1 - \frac{2\hat{\rho}^2}{(1 + \hat{\rho}^2)^2} (1 - \cos t)} dt. \end{aligned}$$

If we use the change of variable

$$x = \frac{2\hat{\rho}^2}{(1 + \hat{\rho}^2)^2} (1 - \cos t), \quad dx = \frac{2\hat{\rho}^2}{(1 + \hat{\rho}^2)^2} \sin t dt,$$

where

$$\sin t = \sqrt{1 - \cos^2 t} = \frac{1 + \hat{\rho}^2}{\hat{\rho}} \sqrt{x} \sqrt{1 - \frac{(1 + \hat{\rho}^2)^2}{4\hat{\rho}^2} x},$$

we have

$$\begin{aligned} \mathbb{V}[\nu_{f_L}(D_{ch}(0, \tilde{\rho}))] &= \frac{L^2}{2\pi} \frac{\hat{\rho}}{1 + \hat{\rho}^2} \int_0^{\frac{4\hat{\rho}^2}{(1 + \hat{\rho}^2)^2}} \frac{(1-x)^L}{1 - (1-x)^L} \frac{\sqrt{x}}{1-x} \frac{1}{\sqrt{1 - \frac{(1 + \hat{\rho}^2)^2}{4\hat{\rho}^2} x}} dx \\ &= \frac{L^2}{2\pi} \frac{\tilde{\rho} \sqrt{4 - \tilde{\rho}^2}}{4} \int_0^{\frac{\tilde{\rho}^2(4 - \tilde{\rho}^2)}{4}} \frac{(1-x)^L}{1 - (1-x)^L} \frac{\sqrt{x}}{1-x} \frac{1}{\sqrt{1 - \frac{4}{\tilde{\rho}^2(4 - \tilde{\rho}^2)} x}} dx. \end{aligned}$$

□

We are going to use a simplified version of (3.2.2), which is just the variance by applying the change of variable  $\tilde{\rho} = 2\rho$ :

$$\mathbb{V}[\nu_{f_L}(D_{ch}(0, 2\rho))] = \frac{L^2}{2\pi} \rho \sqrt{1 - \rho^2} \int_0^{4\rho^2(1-\rho^2)} \frac{(1-x)^L}{1 - (1-x)^L} \frac{\sqrt{x}}{1-x} \frac{1}{\sqrt{1 - \frac{x}{4\rho^2(1-\rho^2)}}} dx. \quad (3.2.3)$$

We are going to use this formula to study the asymptotic behaviour of the zero set of an  $\mathbb{S}^2$ -GAF.

### 3.2.3 Asymptotics as $L \rightarrow +\infty$

Our first case of asymptotic behaviour is when the degree of the polynomial of an  $\mathbb{S}^2$ -GAF,  $L \in \mathbb{N}$ , tends to infinite. This can be translated as increasing the average number of points in a chordal disk  $D_{ch}(z_0, 2\rho)$ ,  $z_0 \in \mathbb{C}$ , which is  $L$  times the surface of the chordal disk. In physical models where the points represent gas particles this is called the *transition to the liquid phase*.

We want to show the next:

**Proposition 3.2.3.** *Let  $f_L$  be an  $\mathbb{S}^2$ -GAF of intensity  $L \in \mathbb{N}$ . Consider a chordal disk  $D_{ch}(z_0, 2\rho)$ , with  $z_0 \in \mathbb{C}$ . Then*

$$\mathbb{V}[\nu_{f_L}(D_{ch}(z_0, 2\rho))] = \left( \frac{\sqrt{L}}{4\sqrt{\pi}} \zeta(3/2) \rho \sqrt{1 - \rho^2} \right) (1 + o(1)), \quad \text{as } L \rightarrow +\infty,$$

where  $\zeta$  stands for the Riemann's zeta function and  $o(1)$  is a term tending to 0 as  $L \rightarrow +\infty$ .

**Remark 3.2.4.** Recall that the variance of the Poisson process was of order  $L$  (see (1.3.3)). However, the variance of the zero set points of an  $\mathbb{S}^2$ -GAF is of order  $\sqrt{L}$ . In this sense, the process  $\mathcal{Z}_{f_L}$  is more *rigid* than the Poisson process as  $L \rightarrow +\infty$ .

*Proof.* By invariance, we can choose  $z_0 = 0$ . By denoting  $D_{ch} := D_{ch}(0, 2\rho)$  and  $\tau = 4\rho^2(1 - \rho^2) \in (0, 1)$ , we get from (3.2.3):

$$\mathbb{V}[\nu_{f_L}(D_{ch})] = \frac{L^2}{4\pi} \sqrt{\tau} \int_0^\tau \frac{(1-x)^L}{1 - (1-x)^L} \frac{\sqrt{x}}{1-x} \frac{1}{\sqrt{1 - x/\tau}} dx. \quad (3.2.4)$$

We want to isolate the leading term as  $L \rightarrow +\infty$ . Specifically, we want to show that this expression can be rewritten as

$$\mathbb{V}[\nu_{f_L}(D_{ch})] = \left( \frac{L^2}{4\pi} \sqrt{\tau} \int_0^\tau \frac{(1-x)^L}{1 - (1-x)^L} \frac{\sqrt{x}}{1-x} dx \right) (1 + o(1)), \quad L \rightarrow +\infty.$$

Denote

$$A_L = \int_0^\tau \frac{(1-x)^L}{1 - (1-x)^L} \frac{\sqrt{x}}{1-x} \frac{1}{\sqrt{1 - x/\tau}} dx, \quad \tilde{A}_L = \int_0^\tau \frac{(1-x)^L}{1 - (1-x)^L} \frac{\sqrt{x}}{1-x} dx.$$

By (3.2.4), the variance can be expressed as

$$\mathbb{V}[\nu_{f_L}(D_{ch})] = \frac{L^2}{4\pi} \sqrt{\tau} \tilde{A}_L \left( 1 + \frac{A_L - \tilde{A}_L}{\tilde{A}_L} \right).$$

In order to get the result, we have to consider

$$\lim_{L \rightarrow +\infty} \frac{A_L - \tilde{A}_L}{\tilde{A}_L} = \lim_{L \rightarrow +\infty} \frac{\int_0^\tau \frac{(1-x)^L}{1-(1-x)^L} \frac{\sqrt{x}}{1-x} \left( \frac{1}{\sqrt{1-x/\tau}} - 1 \right) dx}{\int_0^\tau \frac{(1-x)^L}{1-(1-x)^L} \frac{\sqrt{x}}{1-x} dx} = 0, \quad (3.2.5)$$

and ensure that the order of  $A_L - \tilde{A}_L$  is bigger than the order of  $\tilde{A}_L$ . The following lemma addresses this:

**Lemma 3.2.5.** *It is satisfied that:*

1.  $\tilde{A}_L = \mathcal{O}(L^{-3/2})$ .
2.  $A_L - \tilde{A}_L = \mathcal{O}(L^{-5/2})$ .

Using Lemma 3.2.5, we verify (3.2.5) and we can write

$$\mathbb{V}[\nu_{f_L}(D_{ch})] = \left( \frac{L^2}{\pi} \sqrt{\tau} \int_0^\tau \frac{(1-x)^L}{1-(1-x)^L} \frac{\sqrt{x}}{1-x} dx \right) (1 + o(1)).$$

Noticing that the mass of the integral is concentrated around zero, the integral from 0 to 1 is of the same order. Now, using the geometric series for  $(1-x)^L$  since  $|1-x| < 1$ , the fact that a power series converges uniformly over compact sets in the interval of convergence and using the definition of Beta function (2.3.3), we have that the variance is

$$\begin{aligned} \mathbb{V}[\nu_{f_L}(D_{ch})] &= \left( \frac{L^2}{4\pi} \sqrt{\tau} \int_0^1 \sum_{k=1}^{+\infty} (1-x)^{kL-1} \sqrt{x} dx \right) (1 + o(1)) \\ &= \left( \frac{L^2}{4\pi} \sqrt{\tau} \sum_{k=1}^{+\infty} \int_0^1 (1-x)^{kL-1} \sqrt{x} dx \right) (1 + o(1)) \\ &= \left( \frac{L^2}{4\pi} \sqrt{\tau} \sum_{k=1}^{+\infty} \text{B}(3/2, kL) \right) (1 + o(1)) \\ &= \left( \frac{L^2}{8\sqrt{\pi}} \sqrt{\tau} \sum_{k=1}^{+\infty} \frac{\Gamma(kL)}{\Gamma(kL + 3/2)} \right) (1 + o(1)). \end{aligned}$$

By the asymptotics of  $\Gamma$ , (2.3.2), we conclude

$$\mathbb{V}[\nu_{f_L}(D_{ch})] = \left( \frac{L^2}{8\sqrt{\pi}} \sqrt{\tau} \sum_{k=1}^{+\infty} (kL)^{-3/2} \right) (1 + o(1)) = \left( \frac{\sqrt{L}}{8\sqrt{\pi}} \zeta(3/2) \sqrt{\tau} \right) (1 + o(1)).$$

□

*Proof of Lemma 3.2.5.* 1. Notice that  $\tilde{A}_L$  is equivalent to

$$\begin{aligned}\tilde{A}_L &= \int_0^\tau \frac{(1-x)^L}{1-(1-x)^L} \frac{\sqrt{x}}{1-x} dx = \int_0^1 \frac{(1-x)^L}{1-(1-x)^L} \frac{\sqrt{x}}{1-x} dx - \int_\tau^1 \frac{(1-x)^L}{1-(1-x)^L} \frac{\sqrt{x}}{1-x} dx \\ &= \left( \int_0^1 \frac{(1-x)^L}{1-(1-x)^L} \frac{\sqrt{x}}{1-x} dx \right) \left( 1 - \frac{\int_\tau^1 \frac{(1-x)^L}{1-(1-x)^L} \frac{\sqrt{x}}{1-x} dx}{\int_0^1 \frac{(1-x)^L}{1-(1-x)^L} \frac{\sqrt{x}}{1-x} dx} \right) \\ &= \left( \int_0^1 \frac{(1-x)^L}{1-(1-x)^L} \frac{\sqrt{x}}{1-x} dx \right) \left( 1 - \frac{J_1}{J_2} \right),\end{aligned}$$

where

$$J_1 := \int_\tau^1 \frac{(1-x)^L}{1-(1-x)^L} \frac{\sqrt{x}}{1-x} dx$$

and

$$J_2 := \int_0^1 \frac{(1-x)^L}{1-(1-x)^L} \frac{\sqrt{x}}{1-x} dx.$$

Let us check that the order of the numerator  $J_1$  is bigger than the order of the denominator  $J_2$ . On the one hand, since  $|1-x| < 1$  and using that a power series converges uniformly over compact sets in the interval of convergence, we have, by the Beta function (2.3.3),

$$\begin{aligned}J_2 &= \sum_{k=1}^{+\infty} \int_0^1 (1-x)^{Lk-1} \sqrt{x} dx = \sum_{k=1}^{+\infty} B(3/2, Lk) = \sum_{k=1}^{+\infty} \frac{\Gamma(3/2)\Gamma(Lk)}{\Gamma(Lk+3/2)} \\ &= \frac{\sqrt{\pi}}{2} \sum_{k=1}^{+\infty} \frac{\Gamma(Lk)}{\Gamma(Lk+3/2)}.\end{aligned}$$

By the asymptotics of  $\Gamma$  given in (2.3.2), we get that

$$J_2 = \frac{\sqrt{\pi}}{2} \sum_{k=1}^{+\infty} \frac{\Gamma(Lk)}{\Gamma(Lk+3/2)} \simeq \frac{\sqrt{\pi}}{2} \sum_{k=1}^{+\infty} (Lk)^{-3/2} = \frac{\sqrt{\pi}}{2} \zeta(3/2) L^{-3/2}.$$

Then  $J_2 = \mathcal{O}(L^{-3/2})$ . On the other hand

$$J_1 = \int_\tau^1 \frac{(1-x)^L}{1-(1-x)^L} \frac{\sqrt{x}}{1-x} dx \leq \int_\tau^1 \frac{(1-\tau)^{L-1}}{1-(1-\tau)^L} dx = \frac{(1-\tau)^L}{1-(1-\tau)^L}.$$

Hence  $J_1 = \mathcal{O}((1-\tau)^L)$ . Thus  $\lim_{L \rightarrow +\infty} J_1/J_2 = 0$  and

$$\tilde{A}_L = \left( \int_0^1 \frac{(1-x)^L}{1-(1-x)^L} \frac{\sqrt{x}}{1-x} dx \right) (1 + o(1)).$$

As we did for  $J_2$ , we obtain, as  $L \rightarrow +\infty$ ,

$$\begin{aligned} \int_0^1 \frac{(1-x)^L}{1-(1-x)^L} \frac{\sqrt{x}}{1-x} dx &= \sum_{k=1}^{+\infty} \int_0^1 (1-x)^{Lk-1} \sqrt{x} dx = \frac{\sqrt{\pi}}{2} \sum_{k=1}^{+\infty} \frac{\Gamma(Lk)}{\Gamma(Lk+3/2)} \\ &= \left( \frac{\sqrt{\pi}}{2} \sum_{k=1}^{+\infty} (Lk)^{-3/2} \right) (1+o(1)) \\ &= \left( \frac{\sqrt{\pi}}{2} \zeta(3/2) L^{-3/2} \right) (1+o(1)) \neq 0. \end{aligned}$$

2. We have that

$$A_L - \tilde{A}_L = \int_0^\tau \frac{(1-x)^L}{1-(1-x)^L} \frac{\sqrt{x}}{1-x} \left( \frac{1}{\sqrt{1-x/\tau}} - 1 \right) dx.$$

Notice that the integrand takes high values as  $x$  approaches to 0. For this reason we are going to split the integral into two:

$$\begin{aligned} A_L - \tilde{A}_L &= \int_0^{\tau/2} \frac{(1-x)^L}{1-(1-x)^L} \frac{\sqrt{x}}{1-x} \left( \frac{1}{\sqrt{1-x/\tau}} - 1 \right) dx \\ &\quad + \int_{\tau/2}^\tau \frac{(1-x)^L}{1-(1-x)^L} \frac{\sqrt{x}}{1-x} \left( \frac{1}{\sqrt{1-x/\tau}} - 1 \right) dx =: J_3 + J_4. \end{aligned}$$

Let us focus on  $J_3$ . Developing  $(1-x/\tau)^{-1/2}$  by Taylor at zero we have that, for  $x \leq \tau/2$ :

$$(1-x/\tau)^{-1/2} = 1 + \frac{x}{2\tau} + \mathcal{O}(x^2) \leq 1 + \frac{x}{\tau}.$$

Thus  $(1-x/\tau)^{-1/2} - 1 \leq x/\tau$ . Therefore

$$J_3 \leq \int_0^{\tau/2} \frac{(1-x)^L}{1-(1-x)^L} \frac{\sqrt{x}}{1-x} \frac{x}{\tau} dx = \frac{1}{\tau} \int_0^{\tau/2} \frac{(1-x)^{L-1}}{1-(1-x)^L} x^{3/2} dx.$$

Recall that the mass of the integral is concentrated around zero, so the integral from 0 to 1 is of the same order. Using again the uniform convergence of a power series over compact sets of the interval of convergence:

$$\begin{aligned} J_3 &\leq \frac{1}{\tau} \sum_{k=1}^{+\infty} \int_0^1 (1-x)^{kL-1} x^{3/2} dx = \frac{1}{\tau} \sum_{k=1}^{+\infty} \text{B}(5/2, kL) = \frac{1}{\tau} \sum_{k=1}^{+\infty} \frac{\Gamma(5/2)\Gamma(kL)}{\Gamma(kL+5/2)} \\ &= \frac{3\sqrt{\pi}}{4\tau} \sum_{k=1}^{+\infty} \frac{\Gamma(kL)}{\Gamma(kL+5/2)}. \end{aligned}$$

By using again the asymptotics of  $\Gamma$  given in (2.3.2) we have that

$$J_3 \leq \frac{3\sqrt{\pi}}{4\tau} \sum_{k=1}^{+\infty} \frac{\Gamma(kL)}{\Gamma(kL+5/2)} \simeq \frac{3\sqrt{\pi}}{4\tau} \sum_{k=1}^{+\infty} (kL)^{-5/2} = \frac{3\sqrt{\pi}}{4\tau} \zeta(5/2) L^{-5/2}.$$

Thus  $J_3 = \mathcal{O}(L^{-5/2})$ . With this we conclude that  $\lim_{L \rightarrow +\infty} J_3/\tilde{A}_L = 0$ . For  $J_4$  we have that

$$\begin{aligned} J_4 &= \int_{\tau/2}^{\tau} \frac{(1-x)^L}{1-(1-x)^L} \frac{\sqrt{x}}{1-x} \left( \frac{1}{\sqrt{1-x/\tau}} - 1 \right) dx \\ &\leq \int_{\tau/2}^{\tau} \frac{(1-\tau/2)^L}{1-(1-\tau/2)^L} \frac{\sqrt{\tau}}{1-\tau/2} \frac{dx}{\sqrt{1-x/\tau}}. \end{aligned}$$

By the change of variable  $y = x/\tau$ :

$$J_4 \leq \frac{(1-\tau/2)^L}{1-(1-\tau/2)^L} \frac{\tau^{3/2}}{1-\tau/2} \int_{1/2}^1 \frac{dy}{\sqrt{1-y}} = \frac{(1-\tau/2)^L}{1-(1-\tau/2)^L} \frac{\sqrt{2}\tau^{3/2}}{1-\tau/2}.$$

Therefore  $J_4 = \mathcal{O}((1-\tau/2)^L)$ . This tends to zero much faster than any power of  $L$  as  $L \rightarrow +\infty$ . Then  $\lim_{L \rightarrow +\infty} J_4/\tilde{A}_L = 0$ . □

### 3.2.4 Asymptotics as $\rho \rightarrow 0$

Here we are going to study the case when  $\rho \rightarrow 0$  and  $L$  is fixed. In such a case it is intuitive that the variance of an  $\mathbb{S}^2$ -GAF tends to zero as  $\rho$  does, but here we also see how fast it goes to zero.

We want to show the following:

**Proposition 3.2.6.** *Let  $f_L$  be an  $\mathbb{S}^2$ -GAF of intensity  $L \in \mathbb{N}$ . Consider a chordal disk  $D_{ch}(z_0, 2\rho)$ , with  $z_0 \in \mathbb{C}$ . Then*

$$\mathbb{V}[\nu_{f_L}(D_{ch}(z_0, 2\rho))] = L\rho^2(1 + o(1)) \quad \text{as } \rho \rightarrow 0.$$

**Remark 3.2.7.** The Poisson process and the zero set process of an  $\mathbb{S}^2$ -GAF have the same speed of convergence.

*Proof.* By invariance, we can choose  $z_0 = 0$ . Denote  $D_{ch} := D_{ch}(0, 2\rho)$ . We start from the expression (3.2.4), which is

$$\mathbb{V}[\nu_{f_L}(D_{ch})] = \frac{L^2}{4\pi} \sqrt{\tau} \int_0^{\tau} \frac{(1-x)^L}{1-(1-x)^L} \frac{\sqrt{x}}{1-x} \frac{1}{\sqrt{1-x/\tau}} dx,$$

where  $\tau = 4\rho^2(1-\rho^2)$ . Notice that, since  $\tau \rightarrow 0$ , the Taylor series at a neighbourhood of zero guarantees us that

$$(1-x)^L = e^{L \log(1-x)} = e^{-Lx}(1 + o(1))$$

and

$$1 - (1-x)^L = (1 - e^{-Lx})(1 + o(1)) = Lx(1 + o(1)).$$

Thus, noticing that the integral takes high values near of zero and making the change of variable  $t = x/\tau$ :

$$\begin{aligned}\mathbb{V}[\nu_{f_L}(D_{ch})] &= \left( \frac{L^2}{4\pi} \sqrt{\tau} \int_0^\tau \frac{e^{-Lx}}{Lx} \frac{\sqrt{x}}{1-x} \frac{1}{\sqrt{1-x/\tau}} dx \right) (1 + o(1)) \\ &= \left( \frac{L}{4\pi} \sqrt{\tau} \int_0^\tau e^{-Lx} \frac{dx}{\sqrt{x}\sqrt{1-x/\tau}} \right) (1 + o(1)) \\ &= \left( \frac{L}{4\pi} \sqrt{\tau} \int_0^1 e^{-L\tau t} \frac{\tau}{\sqrt{t\tau}\sqrt{1-t}} dt \right) (1 + o(1)) \\ &= \left( \frac{L}{4\pi} \tau \int_0^1 e^{-L\tau t} \frac{dt}{\sqrt{t}\sqrt{1-t}} \right) (1 + o(1)).\end{aligned}$$

Denote

$$F(\tau) := \int_0^1 e^{-L\tau t} \frac{dt}{\sqrt{t}\sqrt{1-t}}.$$

We have that

$$\begin{aligned}F(\tau) &= \sum_{k=0}^{+\infty} \frac{(-1)^k L^k}{k!} \left( \int_0^1 t^{k-1/2} \frac{dt}{\sqrt{1-t}} \right) \tau^k = \sum_{k=0}^{+\infty} \frac{(-1)^k L^k}{k!} \frac{\Gamma(k+1/2)\Gamma(1/2)}{\Gamma(k+1)} \tau^k \\ &= \sqrt{\pi} \sum_{k=0}^{+\infty} \frac{(-1)^k L^k}{(k!)^2} \Gamma(k+1/2) \tau^k = \sqrt{\pi} \Gamma(1/2) - L\sqrt{\pi} \Gamma(3/2) \tau + \dots \\ &= \pi - \frac{L\pi}{2} \tau + \dots\end{aligned}$$

Hence, as  $\rho \rightarrow 0$ :

$$\begin{aligned}\mathbb{V}[\nu_{f_L}(D_{ch})] &= \left[ \frac{L}{4\pi} \tau \left( \pi - \frac{L\pi}{2} \tau + \dots \right) \right] (1 + o(1)) \\ &= (L\rho^2(1 - \rho^2)) (1 + o(1)) = L\rho^2(1 + o(1)).\end{aligned}$$

□

### 3.2.5 $\mathcal{F}_1$ as the limit of $\mathcal{P}_L$ as $L \rightarrow +\infty$

There is a result stating that  $\mathcal{F}_1$  can be understood as a limit space of  $\mathcal{P}_L$  as  $L \rightarrow +\infty$ . More concretely:

**Lemma 3.2.8.** *Given a GAF  $f_1^{\mathbb{C}} \in \mathcal{F}_1$  and a constant  $M > 0$ , there is  $L_0 \in \mathbb{N}$  such that for all  $L \geq L_0$ , there exist GAFs  $f_L^{\mathbb{S}^2} \in \mathcal{P}_L$  such that*

$$\int_{\{|z| \leq M/\sqrt{L}\}} \left| f_1^{\mathbb{C}}(\sqrt{L}z) - f_L^{\mathbb{S}^2}(z) \right|^2 e^{-L|z|^2} dm(z) \lesssim \frac{1}{L} \|f_1^{\mathbb{C}}\|_{\mathcal{F}_1}^2$$

and

$$\int_{\{|z| > M/\sqrt{L}\}} \frac{|f_L^{\mathbb{S}^2}(z)|^2}{\pi(1+|z|^2)^{L+2}} dm(z) \lesssim \frac{1}{L} \|f_1^{\mathbb{C}}\|_{\mathcal{F}_1}^2.$$

A deeper explanation can be found in [7], p. 32.

In accordance with Lemma 3.2.8, we are going to see whether the limit as  $L \rightarrow +\infty$  of the variance of the random variables of the zero point process of an  $\mathbb{S}^2$ -GAF  $f_L^{\mathbb{S}^2}$  of parameter  $L \in \mathbb{N}$  in  $D_{ch}(z_0, 2r/\sqrt{L})$ , with  $z_0 \in \mathbb{C}$ , is the variance of a  $\mathbb{C}$ -GAF  $f_1^{\mathbb{C}}$  of parameter  $L = 1$ .

**Proposition 3.2.9.** *The limit of the variance of an  $\mathbb{S}^2$ -GAF as  $L \rightarrow +\infty$  coincides with the variance of a  $\mathbb{C}$ -GAF of parameter  $L = 1$ , that is,*

$$\lim_{L \rightarrow +\infty} \mathbb{V} \left[ \nu_{f_L^{\mathbb{S}^2}} \left( D_{ch}(z_0, 2r/\sqrt{L}) \right) \right] = \mathbb{V} \left[ \nu_{f_1^{\mathbb{C}}} (D(z_0, r)) \right], \quad z_0 \in \mathbb{C}.$$

*Proof.* By invariance, we can choose  $z_0 = 0$ . Denote  $D_{ch} := D_{ch}(0, 2r/\sqrt{L})$  and  $D := D(0, r)$ . Choosing  $\rho = r/\sqrt{L}$ , using the new expression of  $\tau$  in (3.2.4) and applying the change of variable  $y = (L^2x)/(4r^2(L-r^2))$ , we have that

$$\begin{aligned} \mathbb{V} \left[ \nu_{f_L^{\mathbb{S}^2}} (D_{ch}) \right] &= \frac{L^2}{4\pi} \sqrt{\frac{4r^2(L-r^2)}{L^2}} \int_0^{\frac{4r^2(L-r^2)}{L^2}} \frac{(1-x)^L}{1-(1-x)^L} \frac{\sqrt{x}}{1-x} \frac{dx}{\sqrt{1-\frac{L^2x}{4r^2(L-r^2)}}} \\ &= \frac{4}{\pi} \frac{r^4(L-r^2)^2}{L^2} \int_0^1 \frac{\left(1 - \frac{4r^2(L-r^2)}{L^2}y\right)^{L-1}}{1 - \left(1 - \frac{4r^2(L-r^2)}{L^2}y\right)^L} y^{1/2}(1-y)^{-1/2} dy. \end{aligned}$$

Denote

$$g_{L,r}(y) = \frac{\left(1 - \frac{4r^2(L-r^2)}{L^2}y\right)^{L-1}}{1 - \left(1 - \frac{4r^2(L-r^2)}{L^2}y\right)^L},$$

which is a decreasing function with respect to  $L$ , bounded below by 0 and

$$\lim_{L \rightarrow +\infty} g_{L,r}(y) = \frac{1}{e^{4r^2y} - 1}.$$

By Monotone Convergence Theorem we get that

$$\lim_{L \rightarrow +\infty} \int_0^1 g_{L,r}(y) dy = \int_0^1 \lim_{L \rightarrow +\infty} g_{L,r}(y) dy = \int_0^1 \frac{dy}{e^{4r^2y} - 1}.$$

Notice that

$$\lim_{L \rightarrow +\infty} \frac{r^4(L-r^2)^2}{L^2} = r^4.$$

Since both limits exist:

$$\begin{aligned} \lim_{L \rightarrow +\infty} \mathbb{V} \left[ \nu_{f_L^{\mathbb{S}^2}} (D_{ch}) \right] &= \frac{4}{\pi} \lim_{L \rightarrow +\infty} \left( \frac{r^4(L-r^2)^2}{L^2} \int_0^1 g_{L,r}(y) y^{1/2}(1-y)^{-1/2} dy \right) \\ &= \frac{4}{\pi} \left( \lim_{L \rightarrow +\infty} \frac{r^4(L-r^2)^2}{L^2} \right) \left( \lim_{L \rightarrow +\infty} \int_0^1 g_{L,r}(y) y^{1/2}(1-y)^{-1/2} dy \right) \\ &= \frac{4}{\pi} r^4 \int_0^1 \frac{y^{1/2}(1-y)^{-1/2}}{e^{4r^2y} - 1} dy. \end{aligned}$$

Now, let us compute the variance of a  $\mathbb{C}$ -GAF for  $L = 1$ . Due to (3.2.1), we can write:

$$\mathbb{V} \left[ \nu_{f_1^{\mathbb{C}}}(D) \right] = \frac{r}{2\pi} \int_0^{4r^2} \frac{1}{e^x - 1} \frac{\sqrt{x}}{\sqrt{1 - \frac{x}{4r^2}}} dx.$$

Applying the change of variable  $y = x/(4r^2)$ :

$$\mathbb{V} \left[ \nu_{f_1^{\mathbb{C}}}(D) \right] = \frac{4}{\pi} r^4 \int_0^1 \frac{y^{1/2}(1-y)^{-1/2}}{e^{4r^2y} - 1} dy,$$

and the result follows.  $\square$

### 3.3 Variance of an $\mathbb{S}^2$ -GAF via linear statistics

In this section we study the fluctuations of an  $\mathbb{S}^2$ -GAF  $f_L$  of parameter  $L \in \mathbb{N}$  through *linear statistics*, that is, for all test-functions  $\varphi \in \mathcal{C}_c^2(\Omega)$ , where  $\Omega \subseteq \mathbb{C}$ ,

$$I_L(\varphi) = \int_{\Omega} \varphi d\nu_{f_L}.$$

We have the following theorem:

**Theorem 3.3.1** ([2], p. 42-44). *Consider a linear statistic  $I_L(\varphi)$ , where  $\varphi \in \mathcal{C}_c^2(\Omega)$  and  $\Omega \subseteq \mathbb{C}$ . Then*

$$\mathbb{V} [I_L(\varphi)] = \frac{\zeta(3)}{32L} \|\Delta^* \varphi\|_{L^2(\Omega, m^*)}^2 + \mathcal{O} \left( \frac{\log L}{L^2} \right),$$

as  $L \rightarrow +\infty$ , where  $\Delta^* := (1 + |z|^2)^2 \Delta$ .

**Remark 3.3.2.**  $\Delta^*$  is called the *invariant Laplacian*, and it is invariant in the sense

$$\Delta^*(u \circ \phi_a) = \Delta^* u \circ \phi_a, \quad u \in \mathcal{C}_c^2(\mathbb{C}),$$

where  $\phi_a$  is the transformation defined in (2.5.2).

*Proof.* By definition

$$\mathbb{V} [I_L(\varphi)] = \mathbb{E} \left[ (I_L(\varphi) - \mathbb{E} [I_L(\varphi)])^2 \right],$$

where

$$I_L(\varphi) = \frac{1}{2\pi} \int_{\Omega} \varphi(z) \Delta \log |f_L(z)|$$

and by the Edelman-Kostlan formula (see Theorem 2.5.3)

$$\mathbb{E} [I_L(\varphi)] = \frac{1}{2\pi} \int_{\Omega} \varphi(z) \Delta \log \sqrt{\mathcal{K}_{f_L}(z, z)}.$$

Letting  $\hat{f}_L \sim N_{\mathbb{C}}(0, 1)$  we get

$$\begin{aligned} I_L(\varphi) - \mathbb{E}[I_L(\varphi)] &= \frac{1}{2\pi} \int_{\Omega} \varphi(z) \Delta \log \frac{|f_L(z)|}{\sqrt{\mathcal{K}_{f_L}(z, z)}} = \frac{1}{2\pi} \int_{\Omega} \varphi(z) \Delta \log |\hat{f}_L(z)| \\ &= \frac{1}{2\pi} \int_{\Omega} \Delta \varphi(z) \log |\hat{f}_L(z)|. \end{aligned}$$

Thus

$$\begin{aligned} \mathbb{V}[I_L(\varphi)] &= \frac{1}{4\pi^2} \mathbb{E} \left[ \int_{\Omega} \Delta \varphi(z) \log |\hat{f}_L(z)| dm(z) \int_{\Omega} \Delta \varphi(w) \log |\hat{f}_L(w)| dm(w) \right] \\ &= \frac{1}{4\pi^2} \int_{\Omega} \int_{\Omega} \Delta \varphi(z) \Delta \varphi(w) \mathbb{E} \left[ \log |\hat{f}_L(z)| \log |\hat{f}_L(w)| \right] dm(z) dm(w). \end{aligned}$$

Since  $\hat{f}_L \sim N_{\mathbb{C}}(0, 1)$ , we have that  $\mathbb{E} \left[ \log |\hat{f}_L(z)| \right]$  is constant (independent of  $z$ ). Therefore:

$$\bar{\partial}_z \bar{\partial}_w \text{Cov} \left[ \log |\hat{f}_L(z)|, \log |\hat{f}_L(w)| \right] = \bar{\partial}_z \bar{\partial}_w \mathbb{E} \left[ \log |\hat{f}_L(z)| \log |\hat{f}_L(w)| \right],$$

and hence

$$\begin{aligned} \mathbb{V}[I_L(\varphi)] &= \frac{1}{4\pi^2} \int_{\Omega} \int_{\Omega} \Delta \varphi(z) \Delta \varphi(w) \text{Cov} \left[ \log |\hat{f}_L(z)|, \log |\hat{f}_L(w)| \right] dm(z) dm(w) \\ &= \frac{1}{4} \int_{\Omega} \int_{\Omega} \Delta^* \varphi(z) \Delta^* \varphi(w) \text{Cov} \left[ \log |\hat{f}_L(z)|, \log |\hat{f}_L(w)| \right] dm^*(z) dm^*(w). \end{aligned}$$

For simplicity, denote  $\kappa_L(z, w) = \text{Cov} \left[ \log |\hat{f}_L(z)|, \log |\hat{f}_L(w)| \right]$ . To estimate the integrals we will separate the points that are close and far from the diagonal due to:

**Lemma 3.3.3.** *For all  $x \in [0, 1]$ , the next estimate holds:*

$$\frac{x}{4} \leq \sum_{j=1}^{+\infty} \frac{x^j}{4j^2} \leq \frac{x}{2}.$$

By using Lemma 3.3.3, we get the estimates

$$\frac{(1 + |\phi_w(z)|^2)^{-L}}{4} \leq \kappa_L(z, w) \leq \frac{(1 + |\phi_w(z)|^2)^{-L}}{2}. \quad (3.3.1)$$

Hence, for a value  $C_L$ , that we are going to specify later, we can write:

$$\Omega^2 = \Omega \times \Omega = \{(z, w) : \kappa_L(z, w) \leq C_L\} \cup \{(z, w) : \kappa_L(z, w) > C_L\}.$$

So, the points  $(z, w) \in \Omega^2$  such that  $\kappa_L(z, w) > C_L$  are those near the diagonal. Otherwise, the points  $(z, w) \in \Omega^2$  such that  $\kappa_L(z, w) \leq C_L$  are those far from the diagonal. We will see that the variance above is only relevant near the diagonal,

and its principal term will be  $\|\Delta^*\varphi\|_{L^2(\Omega, m^*)}^2$ .

We have

$$\begin{aligned}
\mathbb{V}[I_L(\varphi)] &= \frac{1}{4} \int_{\Omega} \int_{\Omega} \Delta^*\varphi(z) \Delta^*\varphi(w) \kappa_L(z, w) dm^*(z) dm^*(w) \\
&= \frac{1}{4} \left[ \int_{\{(z,w) : \kappa_L(z,w) \leq C_L\}} \Delta^*\varphi(z) \Delta^*\varphi(w) \kappa_L(z, w) dm^*(z) dm^*(w) \right. \\
&\quad \left. + \int_{\{(z,w) : \kappa_L(z,w) > C_L\}} \Delta^*\varphi(z) \Delta^*\varphi(w) \kappa_L(z, w) dm^*(z) dm^*(w) \right] \\
&= \frac{1}{4} \left[ \int_{\{(z,w) : \kappa_L(z,w) \leq C_L\}} \Delta^*\varphi(z) \Delta^*\varphi(w) \kappa_L(z, w) dm^*(z) dm^*(w) \right. \\
&\quad + \int_{\{(z,w) : \kappa_L(z,w) > C_L\}} (\Delta^*\varphi(w) - \Delta^*\varphi(z)) \Delta^*\varphi(z) \kappa_L(z, w) dm^*(z) dm^*(w) \\
&\quad \left. + \int_{\{(z,w) : \kappa_L(z,w) > C_L\}} (\Delta^*\varphi(z))^2 \kappa_L(z, w) dm^*(z) dm^*(w) \right].
\end{aligned}$$

Let us name the integrals as:

$$I_1 := \int_{\{(z,w) : \kappa_L(z,w) \leq C_L\}} \Delta^*\varphi(z) \Delta^*\varphi(w) \kappa_L(z, w) dm^*(z) dm^*(w),$$

$$I_2 := \int_{\{(z,w) : \kappa_L(z,w) > C_L\}} (\Delta^*\varphi(w) - \Delta^*\varphi(z)) \Delta^*\varphi(z) \kappa_L(z, w) dm^*(z) dm^*(w)$$

and

$$I_3 := \int_{\{(z,w) : \kappa_L(z,w) > C_L\}} (\Delta^*\varphi(z))^2 \kappa_L(z, w) dm^*(z) dm^*(w).$$

The bound for  $I_1$  is straightforward. Indeed:

$$\begin{aligned}
|I_1| &\leq \int_{\{(z,w) : \kappa_L(z,w) \leq C_L\}} |\Delta^*\varphi(z)| |\Delta^*\varphi(w)| \kappa_L(z, w) dm^*(z) dm^*(w) \\
&\leq C_L \int_{\{(z,w) : \kappa_L(z,w) \leq C_L\}} |\Delta^*\varphi(z)| |\Delta^*\varphi(w)| dm^*(z) dm^*(w) \\
&\leq C_L \int_{\Omega} |\Delta^*\varphi(z)| dm^*(z) \int_{\Omega} |\Delta^*\varphi(w)| dm^*(w) = C_L \|\Delta^*\varphi\|_{L^1(\Omega, m^*)}^2.
\end{aligned}$$

For  $I_2$ , since  $\varphi \in \mathcal{C}_c^2(\Omega)$ ,  $\Delta^*\varphi$  is uniformly continuous in  $\text{supp } \varphi$ . Hence it exists  $\varepsilon(t)$  with  $\lim_{t \rightarrow 1} \varepsilon(t) = 0$  such that, for all  $z, w \in \Omega$  and taking  $\phi_w$  as in (2.5.2) we have

$$|\Delta^*\varphi(w) - \Delta^*\varphi(z)| \leq \varepsilon((1 + |\phi_w(z)|^2)^{-1}).$$

Observe that, from (2.4.6),

$$(1 + |\phi_w(z)|^2)^{-L} = \left[ 1 - \left( \frac{d_{ch}(z, w)}{2} \right)^2 \right]^L = |\Theta_L(z, w)|^2,$$

recalling that

$$\Theta_L(z, w) := \frac{\mathcal{K}_{f_L}(z, w)}{\sqrt{\mathcal{K}_{f_L}(z, z)}\sqrt{\mathcal{K}_{f_L}(w, w)}}.$$

For  $(z, w) \in \Omega^2$  such that  $\kappa_L(z, w) > C_L$ , we get that

$$(2C_L)^{1/L} < (1 + |\phi_w(z)|^2)^{-1},$$

and therefore, since the function  $\varepsilon(t)$  is decreasing,

$$\varepsilon((2C_L)^{1/L}) > \varepsilon((1 + |\phi_w(z)|^2)^{-1}).$$

Thus, by denoting  $\Omega_{\kappa, \varphi} = \{(z, w) : \kappa_L(z, w) > C_L\} \cap (\text{supp } \varphi \times \text{supp } \varphi)$ :

$$\begin{aligned} |I_2| &\leq \int_{\Omega_{\kappa, \varphi}} |\Delta^* \varphi(w) - \Delta^* \varphi(z)| |\Delta^* \varphi(z)| \kappa_L(z, w) dm^*(z) dm^*(w) \\ &\leq \frac{1}{2} \int_{\Omega_{\kappa, \varphi}} \varepsilon((1 + |\phi_w(z)|^2)^{-1}) |\Delta^* \varphi(z)| (1 + |\phi_w(z)|^2)^{-L} dm^*(z) dm^*(w) \\ &\leq \frac{\varepsilon((2C_L)^{1/L})}{2} \int_{\Omega_{\kappa, \varphi}} |\Delta^* \varphi(z)| (1 + |\phi_w(z)|^2)^{-L} dm^*(z) dm^*(w) \\ &\stackrel{(*)}{=} \frac{\varepsilon((2C_L)^{1/L})}{2} \int_{\text{supp } \varphi} |\Delta^* \varphi(z)| dm^*(z) \int_{\{w \in \Omega : \kappa_L(0, w) > C_L\}} (1 + |w|^2)^{-L} dm^*(w) \\ &= \frac{\varepsilon((2C_L)^{1/L})}{2} \|\Delta^* \varphi\|_{L^1(\Omega, m^*)} \int_{\{w \in \Omega : \kappa_L(0, w) > C_L\}} (1 + |w|^2)^{-L} dm^*(w), \end{aligned}$$

where in (\*) we just applied the invariance of the measure  $dm^*$ . Since  $\varepsilon(t) \lesssim |1 - t|$  for all  $t$  in a neighbourhood of 1, we obtain by choosing  $C_L = 1/(2L^2)$  and applying Taylor in a neighbourhood of  $1/L \simeq 0$  (as  $L \rightarrow +\infty$ ):

$$\varepsilon((2C_L)^{1/L}) \lesssim 1 - (2C_L)^{1/L} = 1 - L^{-2/L} \simeq \frac{\log L}{L}.$$

Hence

$$|I_2| \lesssim \frac{\log L}{L} \frac{\|\Delta^* \varphi\|_{L^1(\Omega, m^*)}}{2} \int_{\{w \in \Omega : \kappa_L(0, w) > C_L\}} (1 + |w|^2)^{-L} dm^*(w).$$

For  $I_3$  notice that, by using again the invariance of  $dm^*$ :

$$I_3 = \|\Delta^* \varphi\|_{L^2(\Omega, m^*)}^2 \int_{\{w \in \Omega : \kappa_L(0, w) > C_L\}} (1 + |w|^2)^{-L} dm^*(w).$$

By the fact that  $\lim_{L \rightarrow +\infty} C_L^{1/L} = 1$  we can write that  $I_2 = o(I_3)$  and

$$\mathbb{V}[I_L(\varphi)] = I_3 \left[ 1 + \mathcal{O}\left(\frac{\log L}{L}\right) \right] + \mathcal{O}(C_L).$$

We must compute

$$J_w := \int_{\{w \in \Omega : \kappa_L(0, w) > C_L\}} (1 + |w|^2)^{-L} dm^*(w) = \int_{\{w \in \Omega : \kappa_L(0, w) > C_L\}} |\Theta_L(0, w)|^2 dm^*(w).$$

By the definition of  $\Theta_L$  and using Lemma 3.1.2 we have

$$\begin{aligned} J_w &= \int_{\{w \in \Omega : \kappa_L(0, w) > C_L\}} \sum_{j=1}^{+\infty} \frac{(1 + |w|^2)^{-jL}}{4j^2} dm^*(w) \\ &= \sum_{j=1}^{+\infty} \frac{1}{4j^2} \int_{\{w \in \Omega : \kappa_L(0, w) > C_L\}} (1 + |w|^2)^{-jL} dm^*(w). \end{aligned}$$

From (3.3.1),

$$\{|w|^2 < (4C_L)^{-1/L} - 1\} \subset \{\kappa_L(0, w) > C_L\} \subset \{|w|^2 < (2C_L)^{-1/L} - 1\}.$$

From here we see that

$$\begin{aligned} \{\kappa_L(0, w) > C_L\} &= \\ &= \{|w|^2 < (2C_L)^{-1/L} - 1\} \setminus (\{|w|^2 < (2C_L)^{-1/L} - 1\} \cap \{\kappa_L(0, w) \leq C_L\}). \end{aligned}$$

Therefore

$$\begin{aligned} J_w &= \sum_{j=1}^{+\infty} \frac{1}{4j^2} \left[ \int_{\{|w|^2 < (2C_L)^{-1/L} - 1\}} (1 + |w|^2)^{-jL} dm^*(w) \right. \\ &\quad \left. - \int_{\{|w|^2 < (2C_L)^{-1/L} - 1\} \cap \{\kappa_L(0, w) \leq C_L\}} (1 + |w|^2)^{-jL} dm^*(w) \right]. \end{aligned}$$

The negative sum is negligible. More precisely:

**Lemma 3.3.4.** *It is satisfied that*

$$\sum_{j=1}^{+\infty} \frac{1}{4j^2} \int_{\{|w|^2 < (2C_L)^{-1/L} - 1\} \cap \{\kappa_L(0, w) \leq C_L\}} (1 + |w|^2)^{-jL} dm^*(w) = \mathcal{O}\left(\frac{\log L}{L^3}\right).$$

By Lemma 3.3.4 we get

$$J_w = \sum_{j=1}^{+\infty} \frac{1}{4j^2} \int_{\{|w|^2 < (2C_L)^{-1/L} - 1\}} (1 + |w|^2)^{-jL} dm^*(w) + \mathcal{O}\left(\frac{\log L}{L^3}\right).$$

By denoting

$$I_j = \int_{\{|w|^2 < (2C_L)^{-1/L} - 1\}} (1 + |w|^2)^{-jL} dm^*(w),$$

calling  $r_L = (2C_L)^{-1/L} - 1$  and changing into polar coordinates

$$\begin{aligned} I_j &= \int_{\{|w|^2 < r_L\}} (1 + |w|^2)^{-jL} dm^*(w) = \int_0^{\sqrt{r_L}} (1 + r^2)^{-jL-2} r dr \\ &= \frac{1}{2} \int_0^{r_L} (1 + t)^{-jL-2} dt = \frac{1}{2(1+jL)} [1 - (1+r_L)^{-jL-1}]. \end{aligned}$$

We reach the expression:

$$J_w = \frac{1}{8} \sum_{j=1}^{+\infty} \frac{1}{j^2(1+jL)} [1 - (1+r_L)^{-jL-1}] + \mathcal{O}\left(\frac{\log L}{L^3}\right).$$

Notice that

$$\frac{1}{j^2(1+jL)} \leq \frac{1}{j^2} \quad \text{and} \quad \frac{1}{j^2(1+jL)(1+r_L)^{jL+1}} \leq \frac{1}{j^2}.$$

So we can separate  $J_w$  into two sums because both sums are convergent:

$$J_w = \frac{1}{8} \left[ \sum_{j=1}^{+\infty} \frac{1}{j^2(1+jL)} - \sum_{j=1}^{+\infty} \frac{(1+r_L)^{-jL-1}}{j^2(1+jL)} \right] + \mathcal{O}\left(\frac{\log L}{L^3}\right).$$

Again, the negative sum is negligible. That is:

**Lemma 3.3.5.** *It is satisfied that*

$$\sum_{j=1}^{+\infty} \frac{(1+r_L)^{-jL-1}}{j^2(1+jL)} = \mathcal{O}\left(\frac{\log L}{L^2}\right).$$

By Lemma 3.3.5, we can state that

$$J_w = \frac{1}{8} \sum_{j=1}^{+\infty} \frac{1}{j^2(1+jL)} + \mathcal{O}\left(\frac{\log L}{L^2}\right).$$

We have that  $\lim_{L \rightarrow +\infty} (jL)/(1+jL) = 1$ , uniformly in  $j \geq 1$ . Thus

$$J_w = \frac{1}{8} \sum_{j=1}^{+\infty} \frac{1}{j^3 L} + \mathcal{O}\left(\frac{\log L}{L^2}\right) = \frac{\zeta(3)}{8L} + \mathcal{O}\left(\frac{\log L}{L^2}\right).$$

Hence, as  $L \rightarrow +\infty$ , and recalling that  $I_1 = \mathcal{O}(L^{-2})$  for such a chosen  $C_L$ , we finally have

$$\mathbb{V}[I_L(\varphi)] = \frac{\zeta(3)}{32L} \|\Delta^* \varphi\|_{L^2(\Omega, m^*)}^2 + \mathcal{O}\left(\frac{\log L}{L^2}\right).$$

□

Let us show the three lemmas stated along the proof of Theorem 3.3.1.

*Proof of Lemma 3.3.3.* For all  $x \in [0, 1]$  we must see that

$$\frac{x}{4} \leq \sum_{j=1}^{+\infty} \frac{x^j}{4j^2} \leq \frac{x}{2}.$$

The left bound is clear, since it is the  $j = 1$  term of the sum. For the right bound denote

$$F(x) = x - \frac{1}{2} \sum_{j=1}^{+\infty} \frac{x^j}{j^2}.$$

We want to see that  $F(x) \geq 0$  for all  $x \in [0, 1]$ . By computing the derivative we have

$$F'(x) = 1 - \frac{1}{2x} \sum_{j=1}^{+\infty} \frac{x^j}{j} = 1 - \frac{1}{2x} \log(1+x),$$

and notice that  $F'(x) \geq 0$  since  $\log(1+x) \leq 2x$  by Taylor for  $x \geq 0$ . Then the upper bound holds.  $\square$

*Proof of Lemma 3.3.4.* We want to show that

$$S = \sum_{j=1}^{+\infty} \frac{1}{4j^2} \int_{\{|w|^2 < (2C_L)^{-1/L} - 1\} \cap \{\kappa_L(0, w) \leq C_L\}} (1 + |w|^2)^{-jL} dm^*(w) = \mathcal{O}\left(\frac{\log L}{L^3}\right).$$

We have

$$\begin{aligned} S &\leq \sum_{j=1}^{+\infty} \frac{1}{4j^2} \int_{\{(2C_L)^{1/L} \leq (1+|w|^2)^{-1} \leq (4C_L)^{1/L}\}} (1 + |w|^2)^{-jL} dm^*(w) \\ &= \frac{1}{8\pi} \sum_{j=1}^{+\infty} \frac{1}{j^2} \int_{\{(2C_L)^{1/L} \leq (1+|w|^2)^{-1} \leq (4C_L)^{1/L}\}} (1 + |w|^2)^{-jL-2} dm(w) \\ &\leq \frac{(4C_L)^{1+2/L}}{8\pi} m\left(\{(4C_L)^{-1/L} - 1 \leq |w|^2 \leq (2C_L)^{-1/L} - 1\}\right) \sum_{j=1}^{+\infty} \frac{1}{j^2} \\ &\leq \frac{\pi}{48} (4C_L)^{1+2/L} [(2C_L)^{-1/L} - 1] = \frac{\pi}{48} \left(\frac{2}{L^2}\right)^{1+2/L} (L^{2/L} - 1) \\ &= \frac{\pi}{24} \frac{1}{L^2} 2^{2/L} L^{-4/L} (L^{2/L} - 1). \end{aligned}$$

By Taylor series of  $\exp(x)$  around 0, we have:

$$\begin{aligned} 2^{2/L} &= e^{\frac{2}{L} \log 2} = 1 - \frac{2}{L} \log 2 + \mathcal{O}\left(\frac{1}{L^2}\right) = 1 + \mathcal{O}\left(\frac{1}{L}\right), \\ L^{-4/L} &= e^{-\frac{4}{L} \log L} = 1 - \frac{4}{L} \log L + \mathcal{O}\left(\frac{\log^2 L}{L^2}\right) = 1 + \mathcal{O}\left(\frac{\log L}{L}\right), \\ L^{2/L} - 1 &= e^{\frac{2}{L} \log L} - 1 = \frac{2}{L} \log L + \mathcal{O}\left(\frac{\log^2 L}{L^2}\right) = \mathcal{O}\left(\frac{\log L}{L}\right). \end{aligned}$$

Therefore,

$$\begin{aligned} S &\leq \frac{\pi}{24} \frac{1}{L^2} \left(1 + \mathcal{O}\left(\frac{1}{L}\right)\right) \left(1 + \mathcal{O}\left(\frac{\log L}{L}\right)\right) \left(\mathcal{O}\left(\frac{\log L}{L}\right)\right) \\ &= \mathcal{O}\left(\frac{\log L}{L^3}\right). \end{aligned}$$

□

*Proof of Lemma 3.3.5.* We have to see that

$$S = \sum_{j=1}^{+\infty} \frac{(1+r_L)^{-jL-1}}{j^2(1+jL)} = \mathcal{O}\left(\frac{\log L}{L^2}\right).$$

Recall that  $r_L = (2C_L)^{-1/L} - 1$ . By direct computation we get that  $(1+r_L)^{-jL-1} = L^{-2j-\frac{2}{L}}$ . Since  $L \rightarrow +\infty$ , we obtain  $1+jL \simeq jL$  and  $L^{-2/L} = -2\log L/L + \mathcal{O}(\log^2 L/L^2) \leq \mathcal{O}(\log L/L)$ . Thus

$$S = \sum_{j=1}^{+\infty} \frac{L^{-2j-\frac{2}{L}}}{j^2(1+jL)} \leq \left(\sum_{j=1}^{+\infty} \frac{1}{j^3 L}\right) \mathcal{O}\left(\frac{\log L}{L}\right) = \mathcal{O}\left(\frac{\log L}{L^2}\right).$$

□

# Chapter 4

## Large deviations and the Hole probability of an $\mathbb{S}^2$ -GAF

In this chapter we are going to develop another point of view to understand the fluctuation of a process generated by the zero set of an  $\mathbb{S}^2$ -GAF. For the first section of this chapter, consider a test-function  $\varphi \in \mathcal{C}_c^2(\mathbb{C})$ , an  $\mathbb{S}^2$ -GAF  $f_L$  of parameter  $L \in \mathbb{N}$  and the linear statistic

$$I_L(\varphi) = \int_{\mathbb{C}} \varphi d\nu_{f_L}.$$

As seen before, its mean is

$$\mathbb{E}[I_L(\varphi)] = \mathbb{E} \left[ \int_{\mathbb{C}} \varphi d\nu_{f_L} \right] = \int_{\mathbb{C}} \varphi d\mathbb{E}[\nu_{f_L}] = L \int_{\mathbb{C}} \varphi dm^*.$$

We study how much  $I_L$  deviates from its mean  $\mathbb{E}[I_L(\varphi)]$  by a fraction of this same mean. In particular, we will see that such deviation happens with very small probability.

The second part of this chapter is a consequence of this study. We estimate the probability that there is a hole, that is, a disk without any zeros of  $f_L$ .

We are going to prove all the results of this chapter in terms of the chordal distance. This chapter is strongly based on [6]. Also [4] was helpful to write these pages.

### 4.1 Large deviations

We begin stating the main theorem.

**Theorem 4.1.1** (Large deviations). *For all  $\varphi \in \mathcal{C}_c^2(\mathbb{C})$  and for all  $\delta > 0$ , there exist constants  $c = c(\varphi, \delta) > 0$  and  $L_0 = L_0(\varphi, \delta) \in \mathbb{N}$  such that for all  $L \geq L_0$*

$$\mathbb{P} [|I_L(\varphi) - \mathbb{E}[I_L(\varphi)]| > \delta \mathbb{E}[I_L(\varphi)]] \leq e^{-cL^2}.$$

Notice that this can be rewritten as

$$\mathbb{P} \left[ \left| \frac{I_L(\varphi)}{\mathbb{E}[I_L(\varphi)]} - 1 \right| > \delta \right] \leq e^{-cL^2}.$$

To prove Theorem 4.1.1 we need a sequence of results.

*Proof of Theorem 4.1.1.* Consider  $\varphi \in \mathcal{C}_c^2(\mathbb{C})$ . By the definition of distributional derivative we have

$$I_L(\varphi) - \mathbb{E}[I_L(\varphi)] = \frac{1}{2\pi} \int_{\mathbb{C}} \varphi \Delta \log |\hat{f}_L| = \frac{1}{4\pi} \int_{\text{supp } \varphi} (\Delta \varphi) \log |\hat{f}_L|^2,$$

where  $\hat{f}_L = f_L / \sqrt{\mathcal{K}_{f_L}(z, z)} \sim N_{\mathbb{C}}(0, 1)$ . Hence

$$\begin{aligned} |I_L(\varphi) - \mathbb{E}[I_L(\varphi)]| &\leq \|\Delta \varphi\|_{\infty} \int_{\text{supp } \varphi} \left| \log |\hat{f}_L(z)|^2 \right| (1 + |z|^2)^2 dm^*(z) \\ &\leq \sup_{z \in \text{supp } \varphi} [(1 + |z|^2)^2] \|\Delta \varphi\|_{\infty} \int_{\text{supp } \varphi} \left| \log |\hat{f}_L(z)|^2 \right| dm^*(z) \\ &= C_{\varphi} \int_{\text{supp } \varphi} \left| \log |\hat{f}_L(z)|^2 \right| dm^*(z), \end{aligned}$$

where  $C_{\varphi}$  is a constant that depends on  $\varphi$ . Now we have to see that this integral is small. This is not strange because  $\log |\hat{f}_L|^2$  has values near to zero since  $|\hat{f}_L|^2$  has mean one.

**Lemma 4.1.2.** *For all regular compact set  $K \subset \mathbb{C}$  and for all  $\delta > 0$  there is a constant  $c = c(K, \delta)$  such that*

$$\mathbb{P} \left[ \int_K \left| \log |\hat{f}_L(z)|^2 \right| dm^*(z) > \delta L \right] \leq e^{-cL^2}.$$

To show this we need another lemma that controls the average of  $\left| \log |\hat{f}_L(z)|^2 \right|$  over disks.

**Lemma 4.1.3.** *There is a constant  $c > 0$  such that for all  $D_{ch} := D_{ch}(z_0, \rho) \subset \mathbb{C}$ , with  $z_0 \in \mathbb{C}$ ,*

$$\mathbb{P} \left[ \frac{1}{m^*(D_{ch})} \int_{D_{ch}} \left| \log |\hat{f}_L(w)|^2 \right| dm^*(w) > 5L \right] \leq e^{-cL^2}.$$

Let us see how Lemma 4.1.2 is implied by Lemma 4.1.3.

*Proof of Lemma 4.1.2.* Since  $K$  is a compact set, we can use a minimal covering consisting on a finite number of open disks  $D_j := D_{ch}(z_j, \varepsilon)$ , for all  $j = 1, \dots, N$ , such that all the disks have the same invariant volume, say  $m^*(D_j) = v$ . We will choose  $v$  later on.

By Lemma 4.1.3, outside an exceptional set of probability  $N \exp(-cL^2) \leq \exp(-c'L^2)$ :

$$\begin{aligned} \int_K \left| \log |\hat{f}_L(w)|^2 \right| dm^*(w) &\leq \int_{\bigcup_{j=1}^N D_j} \left| \log |\hat{f}_L(w)|^2 \right| dm^*(w) \\ &\leq \sum_{j=1}^N \int_{D_j} \left| \log |\hat{f}_L(w)|^2 \right| dm^*(w) \\ &\leq \sum_{j=1}^N 5Lm^*(D_j) = 5NLv. \end{aligned}$$

By choosing  $v = \delta/(5N)$ , we conclude that

$$\int_K \left| \log |\hat{f}_L(w)|^2 \right| dm^*(w) \leq \delta L.$$

□

Now we shall show Lemma 4.1.3. For this we need more estimates.

**Lemma 4.1.4.** *For all  $\rho < 1$  and  $\delta > 0$  there exist  $c = c(\rho, \delta) > 0$  and  $L_0 = L_0(\rho, \delta) \in \mathbb{N}$  such that for all  $L \geq L_0$  and  $z_0 \in \mathbb{C}$ :*

- (a)  $\mathbb{P} \left[ \max_{z \in D_{ch}(z_0, \rho)} \log |\hat{f}_L(z)|^2 < -\delta L \right] \leq e^{-cL^2}$ ,
- (b)  $\mathbb{P} \left[ \max_{z \in D_{ch}(z_0, \rho)} \log |\hat{f}_L(z)|^2 > \delta L \right] \leq e^{-ce^{L\delta/2}}$ .

Combining both estimates, we get

$$\mathbb{P} \left[ \max_{z \in D_{ch}(z_0, \rho)} \left| \log |\hat{f}_L(z)|^2 \right| > \delta L \right] \leq e^{-cL^2}.$$

*Proof.* We know that the distribution of  $\mathcal{Z}_{f_L}$  is invariant by translations, hence we can choose  $z_0 = 0$ . By i) of Remark 2.4.3, the condition  $z \in D_{ch}(0, \rho)$  is equivalent to

$$|z| \leq \hat{\rho} := \frac{\rho}{\sqrt{4 - \rho^2}}.$$

(a) Consider the event

$$\mathcal{E}_1 := \left\{ \max_{|z| \leq \hat{\rho}} \log |\hat{f}_L(z)|^2 < -\delta L \right\}.$$

Under this event  $\mathcal{E}_1$ :

$$\log |\hat{f}_L(z)|^2 = \log |f_L(z)|^2 - L \log(1 + |z|^2) < -\delta L$$

and therefore, for  $|z| \leq \hat{\rho}$ :

$$\log |f_L(z)|^2 < L (\log(1 + |z|^2) - \delta) \leq L (\log(1 + \hat{\rho}^2) - \delta).$$

Since  $\log(1 + |z|^2)$  is subharmonic in  $\mathbb{C}$ , we can use the Maximum principle to state:

$$\mathcal{E}_1 \subset \left\{ \max_{|z|=\hat{\rho}} \log |f_L(z)|^2 \leq L (\log(1 + \hat{\rho}^2) - \delta) \right\}.$$

Thus

$$\begin{aligned} \mathbb{P}[\mathcal{E}_1] &\leq \mathbb{P} \left[ \max_{|z|=\hat{\rho}} \frac{\log |f_L(z)|}{L} \leq \frac{1}{2} (\log(1 + \hat{\rho}^2) - \delta) \right] \\ &= \mathbb{P} \left[ \max_{|z|=\hat{\rho}} \frac{\log |f_L(z)|}{L} \leq \left( \frac{1}{2} - \tilde{\delta} \right) \log(1 + \hat{\rho}^2) \right], \end{aligned}$$

where  $\tilde{\delta} = \delta / [2 \log(1 + \hat{\rho}^2)]$ . To continue we need another lemma.

**Lemma 4.1.5.** *For all  $\delta \in (0, 1/2)$  and  $r > 0$  there exist  $c = c(r, \delta)$  and  $L_0 = L_0(r, \delta) \in \mathbb{N}$  such that for all  $L \geq L_0$ :*

$$\mathbb{P} \left[ \max_{|z|=r} \frac{\log |f_L(z)|}{L} \leq \left( \frac{1}{2} - \delta \right) \log(1 + r^2) \right] \leq e^{-cL^2}.$$

*Proof.* Under such event we have

$$\max_{|z|=r} |f_L(z)| \leq (1 + r^2)^{L(\frac{1}{2}-\delta)}.$$

We shall see that this forces some coefficients of the series of  $f_L$  to be small, something that can happen only with probability  $\exp(-cL^2)$ .

Since

$$f_L(z) = \sum_{n=0}^L \frac{f_L^{(n)}(0)}{n!} z^n = \sum_{n=0}^L \xi_n \binom{L}{n}^{1/2} z^n,$$

where  $(\xi_n)_{n=0}^L$  is a sequence of i.i.d.  $N_{\mathbb{C}}(0, 1)$  random variables, we have

$$\xi_n = \binom{L}{n}^{-1/2} \frac{f_L^{(n)}(0)}{n!},$$

for all  $n \in \mathbb{N}$ . Cauchy estimates yield

$$|\xi_n| \leq \binom{L}{n}^{-1/2} \frac{1}{r^n} \max_{|z|=r} |f_L(z)| \leq \binom{L}{n}^{-1/2} \frac{1}{r^n} (1 + r^2)^{L(\frac{1}{2}-\delta)}.$$

Taking squares:

$$|\xi_n|^2 \leq \frac{(1 + r^2)^{L(1-\delta)}}{\binom{L}{n} r^{2n}},$$

where  $\delta$  stands for the old  $\delta^2$ .

This event happens with low probability. Assume that  $n \in \mathbb{N}$  takes high values, that is,  $n/L \geq \varepsilon$  for some  $\varepsilon > 0$  with  $n \leq L$ . By Stirling's formula:

$$n! \sim \sqrt{2\pi n} \left(\frac{n}{e}\right)^n$$

as  $n \rightarrow +\infty$ . Hence

$$\binom{L}{n} \sim \frac{1}{\sqrt{L-n}} \left(\frac{L}{L-n}\right)^{L-n} \left(\frac{L}{n}\right)^n.$$

Therefore, and using  $L-n \leq L$ , we get

$$\begin{aligned} |\xi_n|^2 &\leq \frac{(1+r^2)^{L(1-\delta)}}{\binom{L}{n} r^{2n}} \lesssim \sqrt{L} \frac{(1+r^2)^{L(1-\delta)}}{r^{2n}} \left(\frac{L-n}{L}\right)^{L-n} \left(\frac{n}{L}\right)^n \\ &= \exp \left[ \frac{1}{2} \log L + L(1-\delta) \log(1+r^2) - n \log r^2 + (L-n) \log \left(1 - \frac{n}{L}\right) + n \log \left(\frac{n}{L}\right) \right]. \end{aligned}$$

Denote

$$\Phi(t) = \frac{1}{2} \log L + L(1-\delta) \log(1+r^2) - t \log r^2 + (L-t) \log \left(1 - \frac{t}{L}\right) + t \log \left(\frac{t}{L}\right).$$

We have that  $|\xi_n|^2 \leq \exp \Phi(n)$ . Let us study the behaviour of  $\Phi(t)$ , for all  $t \in [1, L]$ . We have that

$$\Phi(L) = \frac{1}{2} \log L + L(1-\delta) \log(1+r^2) - L \log r^2$$

and

$$\Phi'(t) = \log \left[ \frac{t}{L r^2 \left(1 - \frac{t}{L}\right)} \right] > 0.$$

Define

$$\tilde{\Phi}(s) = \Phi(sL) = \frac{1}{2} \log L + L(1-\delta) \log(1+r^2) - sL \log r^2 + L(1-s) \log(1-s) + sL \log(s).$$

Since

$$\tilde{\Phi}'(s) = -L \log r^2 - L \log(1-s) + L \log s,$$

we have that  $\tilde{\Phi}'(\alpha) = 0$  at  $\alpha = r^2/(1+r^2)$ . In this minimum,

$$\tilde{\Phi}(\alpha) = \frac{1}{2} \log L - \delta L \log(1+r^2),$$

where the negative term is the dominant one as  $L \rightarrow +\infty$ . So, for  $L \in \mathbb{N}$  large enough and taking  $s \in [(1-\delta/2)\alpha, (1+\delta/2)\alpha]$ , we have

$$\tilde{\Phi}(s) \leq -\frac{\delta}{2} L \log(1+r^2).$$

With this, for  $t \in I_{r,L,\delta} := [(1 - \delta/2)L\alpha, (1 + \delta/2)L\alpha]$ , we have

$$\Phi(t) \leq -\frac{\delta}{2}L \log(1 + r^2).$$

In particular, for all  $n \in I_{r,L,\delta}$ ,

$$|\xi_n|^2 \leq e^{\Phi(n)} \leq e^{-\frac{\delta}{2}L \log(1+r^2)}.$$

Denote  $N = \#\{n \in \mathbb{N} : n \in I_{r,L,\delta}\}$ . Note that  $N \simeq \delta L\alpha$ . Thus

$$\begin{aligned} \mathbb{P} \left[ \max_{|z|=r} \frac{\log |f_L(z)|}{L} \leq \left( \frac{1}{2} - \delta \right) \log(1 + r^2) \right] &\leq \prod_{n \in I_{r,L,\delta}} \mathbb{P} \left[ |\xi_n|^2 \leq e^{-\frac{\delta}{2}L \log(1+r^2)} \right] \\ &= \left( \mathbb{P} \left[ |\xi_n|^2 \leq e^{-\frac{\delta}{2}L \log(1+r^2)} \right] \right)^N. \end{aligned}$$

Using the Taylor development of  $1 - \exp(-x)$  around zero we get

$$\left( 1 - e^{-e^{-\frac{\delta}{2}L \log(1+r^2)}} \right)^N \leq e^{-\frac{\delta}{2}NL \log(1+r^2)} = e^{-cL^2},$$

for some  $c$  that depends on  $r$  and  $\delta$ , specifically

$$c = \frac{\delta^2}{2} \frac{r^2}{1+r^2} \log(1+r^2).$$

□

Applying Lemma 4.1.5, we finish the proof of section (a).

(b) Consider the event

$$\mathcal{E}_2 := \left\{ \max_{|z| \leq \rho} \log |\hat{f}_L(z)|^2 > \delta L \right\} = \left\{ \max_{|z| \leq \rho} [\log |f_L(z)|^2 - L \log(1 + |z|^2)] > \delta L \right\}.$$

To estimate the probability of  $\mathcal{E}_2$  we will control the coefficients of the series of  $f_L$ . Let  $C > 0$  be a constant such that  $C\delta < 1$ . From the definition of  $\mathbb{S}^2$ -GAF (2.4.8) we have that

$$|f_L(z)| \leq \sum_{n=0}^{[C\delta L]} |\xi_n| \binom{L}{n}^{1/2} |z|^n + \sum_{n=[C\delta L]+1}^L |\xi_n| \binom{L}{n}^{1/2} |z|^n =: S_1 + S_2.$$

Let us estimate both parts separately.

For  $S_1$  we use Cauchy-Schwarz inequality:

$$\begin{aligned} S_1 &\leq \left( \sum_{n=0}^{[C\delta L]} |\xi_n|^2 \right)^{1/2} \left( \sum_{n=0}^{[C\delta L]} \binom{L}{n} |z|^{2n} \right)^{1/2} \leq \left( \sum_{n=0}^{[C\delta L]} |\xi_n|^2 \right)^{1/2} \left( \sum_{n=0}^L \binom{L}{n} |z|^{2n} \right)^{1/2} \\ &= \left( \sum_{n=0}^{[C\delta L]} |\xi_n|^2 \right)^{1/2} (1 + |z|^2)^{L/2}. \end{aligned}$$

For  $S_2$  we have, using that  $|z| \leq \hat{\rho}$ :

$$S_2 \leq \sum_{n=[C\delta L]+1}^L |\xi_n| \binom{L}{n}^{1/2} \hat{\rho}^n.$$

In the proof of Lemma 4.1.2 we chose disks with very small radii. So, for convenience of the actual proof, we can select  $\rho^2 < 2$  so that  $\hat{\rho} < 1$ . Hence we can consider  $\beta > 0$  such that  $\hat{\rho} = \exp(-\beta)$ . Also take  $\gamma \in (0, \beta)$  and  $\varepsilon > 0$  such that  $0 < \gamma < \gamma + \varepsilon < \beta$ . Define the event

$$\mathcal{A} := \{|\xi_n| \leq e^{\gamma n} : n > [C\delta L]\}.$$

Applying again the Stirling's formula we get

$$\binom{L}{n}^{1/2} \lesssim \frac{L^{L/2}}{n^{n/2}(L-n)^{(L-n)/2}}.$$

Therefore

$$S_2 \leq \sum_{[C\delta L]+1}^L e^{(\gamma-\beta)n} \binom{L}{n}^{1/2} \leq \sum_{[C\delta L]+1}^L e^{(\gamma-\beta)n} \frac{L^{L/2}}{n^{n/2}(L-n)^{(L-n)/2}}.$$

**Lemma 4.1.6.** *For a given  $\varepsilon > 0$  there is  $D > 0$  large enough such that for all  $n > [C\delta L]$ :*

$$\frac{L^L}{n^n(L-n)^{(L-n)}} \leq De^{\varepsilon n}.$$

*Proof.* Notice that

$$\frac{L^L}{n^n(L-n)^{(L-n)}} = e^{L \log L - n \log n - (L-n) \log(L-n)}.$$

Hence, we have to check whether for some  $\varepsilon > 0$  there is  $D > 0$  large enough such that for all  $n > [C\delta L]$ :

$$L \log L - n \log n - (L-n) \log(L-n) \leq \varepsilon n + \log D.$$

Define

$$\Phi_\varepsilon(t) := L \log L - t \log t - (L-t) \log(L-t) - \varepsilon t$$

for all  $t \in [[C\delta L] + 1, L]$ . We want to see that

$$\Phi_\varepsilon(t) \leq \log D.$$

By taking  $t = Ls$  we define the rescaling of  $\Phi_\varepsilon$ :

$$\tilde{\Phi}_\varepsilon(s) := \Phi_\varepsilon(Ls) = L \log L - Ls \log(Ls) - (L-Ls) \log(L-Ls) - \varepsilon Ls$$

for all  $s \in (\alpha, 1]$ . Since

$$\tilde{\Phi}'_\varepsilon(s) = L \log(1-s) - L \log s - L\varepsilon = L \left[ \log \left( \frac{1-s}{s} \right) - \varepsilon \right],$$

we see that  $\tilde{\Phi}'_\varepsilon(s) < 0$  if  $1/(1 + \exp(\varepsilon)) \leq s$ . So, choosing  $\alpha = 1/(1 + \exp(\varepsilon))$ ,  $\tilde{\Phi}_\varepsilon$  is a decreasing function. Clearly, the maximum is at  $s_\varepsilon = 1/(1 + \exp(\varepsilon))$ . Therefore, for all  $s \in (s_\varepsilon, 1]$ :

$$\tilde{\Phi}_\varepsilon(s) \leq C_{\varepsilon,L} := \tilde{\Phi}_\varepsilon(s_\varepsilon).$$

Thus,

$$\Phi_\varepsilon(t) \leq C_{\varepsilon,L}$$

if  $t/L \geq s_\varepsilon$ , i.e., if  $t \geq s_\varepsilon L$ . □

Taking  $D$  appropriately as in Lemma 4.1.6 and since  $0 < \gamma < \gamma + \varepsilon < \beta$ , we get

$$S_2 \lesssim \sum_{n=[C\delta L]+1}^L e^{-[\beta-(\gamma+\varepsilon)]n} \leq \sum_{n=0}^{+\infty} e^{-[\beta-(\gamma+\varepsilon)]n} = \frac{1}{1 - e^{-[\beta-(\gamma+\varepsilon)]}}.$$

Let us show that  $\mathcal{A}$  happens with high probability. Indeed,

$$\mathbb{P}[\mathcal{A}] = \prod_{n=[C\delta L]+1}^L \mathbb{P}[|\xi_n|^2 \leq e^{2\gamma n}] = \prod_{n=[C\delta L]+1}^L \left(1 - e^{-e^{2\gamma n}}\right).$$

Using the fact that  $\log(1-x) \simeq -x$  for  $x \simeq 0$ , we write

$$\log \mathbb{P}[\mathcal{A}] = \sum_{n=[C\delta L]+1}^L \log \left(1 - e^{-e^{2\gamma n}}\right) \simeq - \sum_{n=[C\delta L]+1}^L e^{-e^{2\gamma n}}.$$

There is a constant  $L_0 = L_0(\rho, \delta) \in \mathbb{N}$  such that for all  $L \geq L_0$  and  $n > [C\delta L]$ :

$$-e^{-e^{2\gamma n}} \geq -e^{-e^{\gamma C\delta L}}.$$

Thus

$$\log \mathbb{P}[\mathcal{A}] \geq - \sum_{n=[C\delta L]+1}^L e^{-e^{\gamma C\delta L}} \simeq -e^{-e^{\gamma C\delta L}},$$

and this implies that

$$\mathbb{P}[\mathcal{A}] \geq e^{-e^{-e^{\gamma C\delta L}}}.$$

What we get after all is, under  $\mathcal{A}$ ,

$$|f(z)| \leq \left( \sum_{n=0}^{[C\delta L]} |\xi_n|^2 \right)^{1/2} (1 + |z|^2)^{L/2} + \frac{1}{1 - e^{-[\beta-(\gamma+\varepsilon)]}}.$$

The condition in  $\mathcal{E}_2$ , that is  $|f_L(z)|^2/(1+|z|^2)^L > e^{\delta L}$ , implies

$$\begin{aligned} \sum_{n=0}^{[C\delta L]} |\xi_n|^2 &\geq \left[ (1+|z|^2)^{-L/2} \left( |f_L(z)| - (1 - e^{-[\beta-(\gamma+\varepsilon)]})^{-1} \right) \right]^2 \\ &= \left[ \frac{|f_L(z)|}{(1+|z|^2)^{L/2}} - (1+|z|^2)^{-L/2} (1 - e^{-[\beta-(\gamma+\varepsilon)]})^{-1} \right]^2 \\ &> \left[ e^{\delta L/2} - (1+|z|^2)^{-L/2} (1 - e^{-[\beta-(\gamma+\varepsilon)]})^{-1} \right]^2 > e^{\delta L}. \end{aligned}$$

Hence

$$\begin{aligned} \mathbb{P}[\mathcal{E}_2 \cap \mathcal{A}] &\leq \mathbb{P} \left[ \sum_{n=0}^{[C\delta L]} |\xi_n|^2 \geq e^{\delta L} \right] \leq \sum_{n=0}^{[C\delta L]} \mathbb{P} [|\xi_n|^2 \geq e^{\delta L}] = \sum_{n=0}^{[C\delta L]} e^{-e^{\delta L}} \\ &\lesssim e^{-e^{\delta L/2}}. \end{aligned}$$

Finally

$$\mathbb{P}[\mathcal{E}_2] = \frac{\mathbb{P}[\mathcal{E}_2 \cap \mathcal{A}]}{\mathbb{P}[\mathcal{A}]} \leq e^{-e^{\delta L/2}} e^{e^{-\gamma C\delta L}} \leq e^{-ce^{L\delta/2}},$$

for a suitable  $c > 0$ .

□

We have to show Lemma 4.1.3. However, we need a result that estimates the average of  $\log |\hat{f}_L(z_0)|^2$ , for  $z_0 \in \mathbb{C}$ .

**Lemma 4.1.7.** *Let  $D_{ch}^{z_0} := D_{ch}(z_0, \rho)$  be a disk with  $z_0 \in \mathbb{C}$  and  $\rho > 0$ . Then*

$$\log |\hat{f}_L(z_0)|^2 \leq \frac{1}{m^*(D_{ch}^{z_0})} \int_{D_{ch}^{z_0}} \log |\hat{f}_L(w)|^2 dm^*(w) + L\varepsilon(\hat{\rho}, z_0),$$

where, by *i)* of Remark 2.4.3,

$$\hat{\rho} := \frac{\rho}{\sqrt{4 - \rho^2}}$$

and

$$\varepsilon(\hat{\rho}, z_0) := \begin{cases} 1 - \frac{\log(1 + \hat{\rho}^2)}{\hat{\rho}^2}, & \text{if } |z_0 \hat{\rho}| < 1, \\ \left( 1 - \frac{\log(1 + \hat{\rho}^2)}{\hat{\rho}^2} \right) - \frac{1}{m^*(D_{ch}^0)} \int_{1/|z_0|}^{\hat{\rho}} \log |\overline{z_0} r|^2 \frac{2r}{(1+r^2)^2} dr, & \text{if } |z_0 \hat{\rho}| > 1. \end{cases}$$

For any case,  $\varepsilon(\hat{\rho}, z_0) \leq 1$ .

**Remark 4.1.8.** The function  $\varepsilon$  can be written in terms of  $\rho$  and the integral is computable, but we do not need that. We will need only the inequality of  $\varepsilon$ .

*Proof.* By the subharmonicity of  $\log |\hat{f}_L|^2$ , we have

$$\begin{aligned} \log |\hat{f}_L(z_0)|^2 &= \log |f_L(z_0)|^2 - L \log \mathcal{K}_{f_L}(z_0, z_0) = \log |f_L(z_0)|^2 - L \log(1 + |z_0|^2) \\ &\leq \frac{1}{m^*(D_{ch}^{z_0})} \int_{D_{ch}^{z_0}} \log |f_L(w)|^2 dm^*(w) - L \log(1 + |z_0|^2) \\ &= \frac{1}{m^*(D_{ch}^{z_0})} \int_{D_{ch}^{z_0}} \log |\hat{f}_L(w)|^2 dm^*(w) \\ &\quad + L \left[ \frac{1}{m^*(D_{ch}^{z_0})} \int_{D_{ch}^{z_0}} \log(1 + |w|^2) dm^*(w) - \log(1 + |z_0|^2) \right]. \end{aligned}$$

Name

$$I := \int_{D_{ch}^{z_0}} \log(1 + |w|^2) dm^*(w).$$

By (2.4.5), we get

$$\begin{aligned} I &= \int_{D_{ch}^0} \log(1 + |\phi_{z_0}(w)|^2) dm^*(w) = \int_{D_{ch}^0} \log \left[ \frac{(1 + |z_0|^2)(1 + |w|^2)}{|1 + \bar{z}_0 w|^2} \right] dm^*(w) \\ &= \int_{D_{ch}^0} \log(1 + |z_0|^2) dm^*(w) + \int_{D_{ch}^0} \log(1 + |w|^2) dm^*(w) - \int_{D_{ch}^0} \log |1 + \bar{z}_0 w|^2 dm^*(w). \end{aligned}$$

Since  $m^*(D_{ch}^{z_0}) = m^*(D_{ch}^0)$  and using the new expression of  $I$ , we have

$$\begin{aligned} \log |\hat{f}_L(z_0)|^2 &= \frac{1}{m^*(D_{ch}^{z_0})} \int_{D_{ch}^{z_0}} \log |\hat{f}_L(w)|^2 dm^*(w) \\ &\quad + L \left[ \frac{1}{m^*(D_{ch}^0)} \int_{D_{ch}^0} \log(1 + |w|^2) dm^*(w) - \frac{1}{m^*(D_{ch}^0)} \int_{D_{ch}^0} \log |1 + \bar{z}_0 w|^2 dm^*(w) \right]. \end{aligned}$$

Using a polar coordinate change of variables,

$$\int_{D_{ch}^0} \log |1 + \bar{z}_0 w|^2 dm^*(w) = \frac{1}{\pi} \int_0^{\hat{\rho}} \left( \int_0^{2\pi} \log |1 + \bar{z}_0 r e^{i\theta}|^2 d\theta \right) \frac{r}{(1 + r^2)^2} dr.$$

Let us study the integral in  $\theta$  by distinguishing cases:

- If  $|z_0 \hat{\rho}| < 1$ , the harmonicity of the integrand implies

$$\int_0^{2\pi} \log |1 + \bar{z}_0 r e^{i\theta}|^2 d\theta = 2\pi \log(1) = 0.$$

- If  $|z_0 \hat{\rho}| > 1$ , by using the previous case and the change of variable  $\psi = -\theta$  we can conclude that

$$\begin{aligned} \int_0^{2\pi} \log |1 + \bar{z}_0 r e^{i\theta}|^2 d\theta &= \int_0^{2\pi} \log |\bar{z}_0 r|^2 d\theta + \int_0^{2\pi} \log |1 + (\bar{z}_0 r)^{-1} e^{-i\theta}|^2 d\theta \\ &= 2\pi \log |\bar{z}_0 r|^2 + \int_0^{2\pi} \log |1 + (\bar{z}_0 r)^{-1} e^{-i\psi}|^2 d\psi \\ &= 2\pi \log |\bar{z}_0 r|^2. \end{aligned}$$

Thus, if  $|z_0\hat{\rho}| > 1$ :

$$\begin{aligned} \int_{D_{ch}^0} \log |1 + \bar{z}_0 w|^2 dm^*(w) &= \frac{1}{\pi} \int_{1/|z_0|}^{\hat{\rho}} \left( \int_0^{2\pi} \log |1 + \bar{z}_0 r e^{i\theta}|^2 d\theta \right) \frac{r}{(1+r^2)^2} dr \\ &= \int_{1/|z_0|}^{\hat{\rho}} \log |\bar{z}_0 r|^2 \frac{2r}{(1+r^2)^2} dr. \end{aligned}$$

By using another change into polar coordinates, the change of variable  $t = 1 + r^2$  and integration by parts with  $u = \log t$ ,  $dv = dt/t^2$ , we have

$$\begin{aligned} \int_{D_{ch}^0} \log(1 + |w|^2) dm^*(w) &= \int_0^{\hat{\rho}} \log(1 + r^2) \frac{2r}{(1+r^2)^2} dr = \int_1^{1+\hat{\rho}^2} \frac{\log t}{t^2} dt \\ &= -\frac{\log t}{t} \Big|_{t=1}^{t=1+\hat{\rho}^2} + \int_1^{1+\hat{\rho}^2} \frac{dt}{t^2} = \frac{\hat{\rho}^2 - \log(1 + \hat{\rho}^2)}{1 + \hat{\rho}^2} \\ &= m^*(D_{ch}^0) \left( 1 - \frac{\log(1 + \hat{\rho}^2)}{\hat{\rho}^2} \right). \end{aligned}$$

Hence the expression of  $\varepsilon(\hat{\rho}, z_0)$  follows and it is trivial that  $\varepsilon(\hat{\rho}, z_0) \leq 1$  for all  $\hat{\rho} > 0$  and  $z_0 \in \mathbb{C}$ . □

*Proof of Lemma 4.1.3.* Using section (a) of Lemma 4.1.4, outside of an exceptional set of probability  $\exp(-cL^2)$ , there exists  $\lambda \in D_{ch} := D_{ch}(z_0, \rho) \subset \mathbb{C}$  such that

$$\log |\hat{f}_L(\lambda)|^2 > -Lm^*(D_{ch}).$$

By Lemma 4.1.7,

$$-Lm^*(D_{ch}) < \frac{1}{m^*(D_{ch})} \int_{D_{ch}} \log |\hat{f}_L(w)|^2 dm^*(w) + L,$$

which implies

$$0 < \frac{1}{m^*(D_{ch})} \int_{D_{ch}} \log |\hat{f}_L(w)|^2 dm^*(w) + 2L.$$

Separating the logarithm into the positive and negative part:

$$\frac{1}{m^*(D_{ch})} \int_{D_{ch}} \log^- |\hat{f}_L(w)|^2 dm^*(w) \leq \frac{1}{m^*(D_{ch})} \int_{D_{ch}} \log^+ |\hat{f}_L(w)|^2 dm^*(w) + 2L.$$

Adding the positive part of the logarithm,

$$\frac{1}{m^*(D_{ch})} \int_{D_{ch}} \left| \log |\hat{f}_L(w)|^2 \right| dm^*(w) \leq \frac{2}{m^*(D_{ch})} \int_{D_{ch}} \log^+ |\hat{f}_L(w)|^2 dm^*(w) + 2L.$$

By Lemma 4.1.4, outside an exceptional set of probability  $\exp(-cL^2)$  we conclude

$$\begin{aligned} \frac{1}{m^*(D_{ch})} \int_{D_{ch}} \left| \log |\hat{f}_L(w)|^2 \right| dm^*(w) &\leq 2 \max_{w \in D_{ch}} \log^+ |\hat{f}_L(w)|^2 + 2L \\ &\leq 3Lm^*(D_{ch}) + 2L \leq 5L. \end{aligned}$$

□

By applying Lemma 4.1.3, the proof of Theorem 4.1.1 is finished.  $\square$

A direct consequence of Theorem 4.1.1, useful to show the Hole probability theorem, is the following.

**Corollary 4.1.9.** *Let  $U$  be a bounded open set in  $\mathbb{C}$ . For all  $\delta \in (0, 1)$  there are constants  $c = c(U, \delta) > 0$  and  $L_0 = L_0(U, \delta) \in \mathbb{N}$  such that for all  $L \geq L_0$*

$$\mathbb{P} \left[ \left| \frac{I_L(\mathbb{1}_U)}{\mathbb{E}[I_L(\mathbb{1}_U)]} - 1 \right| > \delta \right] = \mathbb{P} \left[ \left| \frac{\nu_{f_L}(U)}{Lm^*(U)} - 1 \right| > \delta \right] \leq e^{-cL^2}.$$

*Proof.* Consider test-functions  $\varphi_1, \varphi_2 \in \mathcal{C}_c^2(\mathbb{C})$  such that  $0 \leq \varphi_1 \leq \mathbb{1}_U \leq \varphi_2 \leq 1$ ,

$$\int_{\mathbb{C}} \varphi_1 dm^* \geq m^*(U)(1 - \delta)$$

and

$$\int_{\mathbb{C}} \varphi_2 dm^* \leq m^*(U)(1 + \delta).$$

By Theorem 4.1.1 we have, outside an exceptional set of probability  $\exp(-cL^2)$ :

$$\begin{aligned} \nu_{f_L}(U) &= \int_U d\nu_{f_L} \leq \int_{\mathbb{C}} \varphi_2 d\nu_{f_L} \leq (1 + \delta) \mathbb{E} \left[ \int_{\mathbb{C}} \varphi_2 d\nu_{f_L} \right] = (1 + \delta) \int_{\mathbb{C}} \varphi_2 d\mathbb{E}[\nu_{f_L}] \\ &= L(1 + \delta) \int_{\mathbb{C}} \varphi_2 dm^* \leq L(1 + \delta)^2 m^*(U). \end{aligned}$$

Analogously, outside an exceptional set of probability  $\exp(-cL^2)$ , we get:

$$\nu_{f_L}(U) \geq L(1 - \delta)^2 m^*(U).$$

Thus:

$$\frac{\nu_{f_L}(U)}{Lm^*(U)} - 1 \leq 2\delta + \delta^2 \leq 3\delta.$$

Also

$$\frac{\nu_{f_L}(U)}{Lm^*(U)} - 1 \geq -2\delta + \delta \geq -3\delta.$$

$\square$

## 4.2 Hole probability

In the last section of the project we focus on the *Hole probability*.

**Theorem 4.2.1.** *For a given  $\rho > 0$ , there exist  $C_1 = C_1(\rho) > 0$ ,  $C_2 = C_2(\rho) > 0$  and  $L_0 \in \mathbb{N}$  such that, for all  $L \geq L_0$  and for all  $z_0 \in \mathbb{C}$ ,*

$$e^{-C_1 L^2} \leq \mathbb{P}[\mathcal{Z}_{f_L} \cap D_{ch}(z_0, \rho) = \emptyset] \leq e^{-C_2 L^2},$$

where  $\mathcal{Z}_{f_L}$  is the zero set of the  $\mathbb{S}^2$ -GAF  $f_L$ .

*Proof. Upper bound.* Since

$$\{\nu_{f_L}(D_{ch}(z_0, \rho)) = 0\} \subset \left\{ \left| \frac{\nu_{f_L}(D_{ch}(z_0, \rho))}{Lm^*(D_{ch}(z_0, \rho))} - 1 \right| > \delta \right\},$$

for all  $\delta < 1$ , Corollary 4.1.9 gives us the upper bound.

*Lower bound.* By the invariance we can assume that  $z_0 = 0$ . Denote  $D_{ch} := D_{ch}(0, \rho)$ . We are going to choose two events that force  $f_L$  to have  $\mathcal{Z}_{f_L} \cap D_{ch} = \emptyset$ . Clearly

$$|f_L(z)| \geq |\xi_0| - \left| \sum_{n=1}^L \xi_n \binom{L}{n}^{1/2} z^n \right|.$$

The first event is

$$\mathcal{E}_1 := \{|\xi_0| \geq 1\}.$$

By i) of Proposition 2.1.2,  $\mathcal{E}_1$  has probability

$$\mathbb{P}[\mathcal{E}_1] = \mathbb{P}[|\xi_0|^2 \geq 1] = e^{-1}.$$

By i) of Remark 2.4.3, take

$$\hat{\rho} := \frac{\rho}{\sqrt{4 - \rho^2}}.$$

For the second term we use the Cauchy-Schwarz inequality for  $|z| \leq \hat{\rho}$ :

$$\begin{aligned} \left| \sum_{n=1}^L \xi_n \binom{L}{n}^{1/2} z^n \right| &\leq \sum_{n=1}^L |\xi_n| \binom{L}{n}^{1/2} \hat{\rho}^n \leq \left( \sum_{n=1}^L |\xi_n|^2 \right)^{1/2} \left( \sum_{n=1}^L \binom{L}{n} \hat{\rho}^{2n} \right)^{1/2} \\ &\leq \left( \sum_{n=1}^L |\xi_n|^2 \right)^{1/2} (1 + \hat{\rho}^2)^{L/2}. \end{aligned}$$

Choosing the second event as

$$\mathcal{E}_2 := \left\{ |\xi_n|^2 \leq \frac{1}{16L(1 + \hat{\rho}^2)^L}, n = 1, \dots, L \right\},$$

we see that, under  $\mathcal{E}_2$ :

$$\left| \sum_{n=1}^L \xi_n \binom{L}{n}^{1/2} z^n \right| \leq \frac{1}{4}.$$

Knowing that  $1 - \exp(-x) \geq x/2$  for  $x \leq 1$ , we get

$$\begin{aligned} \mathbb{P}[\mathcal{E}_2] &= \prod_{n=1}^L \mathbb{P} \left[ |\xi_n|^2 \leq \frac{1}{16L(1 + \hat{\rho}^2)^L} \right] = \left[ 1 - \exp \left( -\frac{1}{16L(1 + \hat{\rho}^2)^L} \right) \right]^L \\ &\geq \left( \frac{1}{32L(1 + \hat{\rho}^2)^L} \right)^L = e^{-L^2 \log[32L(1 + \hat{\rho}^2)]} = e^{-CL^2}, \end{aligned}$$

where  $C = C(\hat{\rho}, L) := \log[32L(1 + \hat{\rho}^2)]$ .

All combined, under the event  $\mathcal{E}_1 \cap \mathcal{E}_2$  we have that  $f_L(z) \geq 3/4 > 0$  and

$$\mathbb{P}[\mathcal{E}_1 \cap \mathcal{E}_2] = \mathbb{P}[\mathcal{E}_1] \mathbb{P}[\mathcal{E}_2] \geq e^{-1} e^{-CL^2} \geq e^{-C_1 L^2},$$

for a suitable  $C_1 = C_1(\rho) > 0$ . □

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