# Percentile charts for speeding based on telematics information 

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#### Abstract

Reference charts are widely used as a graphical tool for assessing and monitoring children's growth given gender and age. Here, we propose a similar approach to the assessment of driving risk. Based on telematics data, and using quantile regression models, our methodology estimates the percentiles of the distance driven at speeds above the legal limit depending on drivers' characteristics and the journeys made. We refer to the resulting graphs as percentile charts for speeding and illustrate their use for a sample of drivers with Pay-How-You-Drive insurance policies. We find that percentiles of distance driven at excessive speeds depend mainly on total distance driven, the percentage of driving in urban areas and the driver's gender. However, the impact on the estimated percentile for these covariates is not constant. We conclude that the heterogeneity in the risk of driving long distances above the speed limit can be easily represented using reference charts and that, conversely, individual drivers can be scored by calculating an estimated percentile for their specific case. The dynamics of this risk score can be assessed by recording drivers as they accumulate driving experience and cover more kilometres. Our methodology should be useful for accident prevention and, in the context of Manage-How-You-Drive insurance, reference charts can provide real-time alerts and enhance recommendations for ensuring safety.


Keywords: motor insurance; speed; telematics; quantile regression; reference curves; risk score.

## 1. Introduction

Growth reference charts are used worldwide to provide a simple graphical tool for monitoring the evolution in children's height and weight. As such, they enable doctors and parents to track a child's estimated percentile path and observe his or her position with respect to that of their corresponding reference population of either boys or girls. Here, we seek to design a similar tool for assessing driving risk, based on the distance driven above the posted speed limit as an indicator of peril. A driver's risk evolution is then analysed with respect to total distance driven and other circumstances that need to be taken into consideration, primarily driving zone. The tool developed is both highly informative and simple, and can be directly used to communicate driving risk.

Recent research on road traffic safety specifically related to speeding highlights that speed increases both the risk and severity of an accident (see Dissanayake and Lu, 2002; Ossiander and Cummings, 2002; Jun et al.,

2007, 2011; Vernon et al., 2004 and Viallon and Laumon, 2013). Viallon and Laumon (2013) analysed the effectiveness of the speed regulation policies introduced during the period 2001-2010 in France with respect to high-level speeding. The authors built a model which relates the number of fatal crashes to speed, and highlighted the effectiveness of speed regulation policies. More recently, Arvin et al. (2019) investigated the relationship between pre-crash driving instability and crash intensity. They found that higher instability in driving increases the probability of a severe crash, and that the speed prior a crash is highly correlated with its intensity. Bogstrand et al. (2015) analysed fatal road traffic crashes in Norway during 2005-2010, and found statistically significant associations between impairment by alcohol or amphetamines and driving unbelted or speeding. The authors also found that excessive speeding is one of the main reasons for traffic crashes. Additionally, speeding and being unbelted are the main reasons for a fatal outcome. More recently, Høye (2020) intestigated fatal car crashes in Norway as well, but from 2005 to 2015. They found that individuals who drive under the influence of alcohol and/or drugs are more often male, unbelted and unlicensed. Moreover, they normally drive old cars, and are involved in single-vehicle crashes at night, in the weekend and on low-volume roads.

There are also evidences that drivers are not homogeneous with respect to their level of risk and driving style. Specifically, men present riskier driving patterns, driving more kilometres per day, during the night and at speeds above the limit than women (Ayuso et al., 2014, 2016a, 2016b). All these factors have been shown to be associated with a greater number of accidents (Gao et al., 2019a; Gao and Wüthrich, 2019; Guillen et al., 2019). Moreover, Paefgen et al. (2014) report that the risk of accident is higher on urban roads, during weekends, at nightfall and at low- ( $0-30 \mathrm{~km} / \mathrm{h}$ ) or high-range speeds ( $90-120 \mathrm{~km} / \mathrm{h}$ ). Indeed, Pérez-Marín and Guillen (2019) concluded that if excess speeds could be eliminated, the expected number of accident claims would be reduced by half. Interestingly, Pérez-Marín et al. (2019a) showed that young drivers tend to reduce posted speed limit violations after an accident, probably because they are more aware of the risk.

Speed and driving distance have been exhaustively analysed in transport research (see, for example, Hewson, 2008 or Plötz et al., 2017). Moreover, analyses of speeding in traffic safety research have focused not only on the average speed, but also on its quantiles. Specifically, Hewson (2008) explored the benefits of using quantile regression to evaluate whether or not an intervention is able to significantly modify the $85^{\text {th }}$ percentile speed. Recently, Pérez-Marín et al. (2019b) applied quantile regression to an analysis of the effects of telematics information (location and time of driving and the total distance driven) on a range of percentiles of the distance driven at speeds above the limit by using a sample of drivers covered by a Pay-How-You-Drive (PHYD) insurance policy. In PHYD policies, the premium is calculated based on the customer's driving pattern (such as speeding, harsh acceleration, sudden braking or hard cornering). Based on these patterns, a driver's risk score can be obtained and used to calculate his or her premium (see a survey in Arumugam and Bhargavi, 2019).

In this paper, we propose a methodology for displaying percentiles that allows us to quantify a driver's risk score. To do so, we use a graphical representation of the percentiles of distance driven at speeds above the limit, depending on specific driver characteristics and on the sort of trips they make. Employing charts similar to the well-known reference curves for child growth, we develop a new methodology in the context of speeding that should prove useful when a large number of covariates can influence a driver's behaviour on the road and, hence, their risk profile.

Specifically, we call our graphs percentile charts for speeding, as they provide each driver with their corresponding percentile of distance driven at speeds above the legal limits, given all available information on that driver. This proves to be a straightforward risk score for the driver. We take the article by Perez-Marín et al. (2019) as our starting point, and use the methodology proposed by Wei et al. (2006) in the context of growth charts (based on quantile regression) to produce percentile charts for speeding. We use the same data as presented in Perez-Marín et al. (2019b) and explore alternative model formulations in the context of generalized linear models (GLMs) and quantile regression. In particular, we investigate in-depth the relationship between distance driven at speeds above the legal limits (the dependent variable in our regression
models) and total distance driven. We conclude that their relationship is not linear, but exponential. This exponential relationship determines the shape of the percentile charts for speeding. As a result, we also observe that our methodology substantially improves the initial results obtained in Perez-Marín et al. (2019b).

The rest of this paper is organized as follows. In section 2, the quantile regression model and the database used in our study are presented. In section 3, the main results of the regression models are summarized and the percentile charts are provided. In section 4, the main results are discussed, and finally in section 5 the main conclusions of the paper are presented.

## 2. Material and Methods

### 2.1. Methods

Percentile charts for speeding are obtained by means of quantile regression, where each curve corresponds to a percentile level. A web application is easily designed, so that when a user enters his or her covariate information and observed mileage above the speed limit, a graph is displayed, locating the specific driver on the chart. Quantile regression is specially recommended when the response variable is asymmetric conditional on the explanatory variables. OLS and Gaussian assumptions would produce biased estimates of extreme quantiles (see Khattak et al , 2016, for a comparison in a case study on transportation). Eide and Showalter (1998) presented one of the earliest comparisons between OLS and quantiles regression in the field of education. They found a differential effect of school size at different points of the students' test score gain conditional distribution, meaning that school size has increasing return of scale on the gain in students' scores (see also Castellano and Ho, 2013). Similar examples of quantile regression versus OLS can be found in Bel et al. (2015) and O'Garra and Mourato (2006) in environmental economics.

In this paper, we also fit a GLM model prior to quantile regression; specifically, we fit a gamma model because the dependent variable, which is mileage above the speed limit, is expected to be asymmetric. That is, while a large number of drivers can be expected not to exceed the speed limit over a certain number of kilometres, only a few are expected to exceed the limit over a high percentage of the distance driven.

The $\tau$-quantile of a continuous random variable $Y$ is the value $c_{\tau}$ for which $P\left(Y \leq c_{\tau}\right)=\tau$. In the financial and actuarial industries, the $\tau$-quantile, or the percentile at the level $\tau$, is known as the value-atrisk at level $\tau$. Quantile regression is used in order to estimate conditional quantiles, as the model assumes that the $c_{\tau}(Y)$ depends on certain explanatory variables. Specifically,

$$
\begin{equation*}
c_{\tau}\left(Y_{i} \mid X_{1 i}, \ldots, X_{k i}\right)=\beta_{0}^{\tau}+\beta_{1}^{\tau} X_{1 i}+\beta_{2}^{\tau} X_{2 i}+\ldots+\beta_{k}^{\tau} X_{k i} \tag{1}
\end{equation*}
$$

where $Y_{i}$ is the dependent variable for the $i$-th individual, with $i=1, \ldots, n$, and $X_{j i}$ are the observations of the explanatory variables, with $j=1, \ldots, k$. It can be proved (Koenker and Bassett, 1978) that

$$
\begin{equation*}
\widehat{\beta^{\tau}}=\underset{b}{\operatorname{argmin}}\left[\sum_{Y_{i} \geq X_{i}^{\prime} b} \tau\left|Y_{i}-X_{i}^{\prime} \mathrm{b}\right|+\sum_{Y_{i}<X_{i}^{\prime} b}(1-\tau)\left|Y_{i}-X_{i}^{\prime} \mathrm{b}\right|\right] \tag{2}
\end{equation*}
$$

The objective function (2) corresponds to the sum of $n$ components, called $\rho_{\tau}\left(Y_{i}-X_{i}^{\prime} \mathrm{b}\right)$ that are expressed as follows:

$$
\begin{align*}
\rho_{\tau}\left(Y_{i}-X_{i}^{\prime} \mathrm{b}\right)=\tau\left(Y_{i}-X_{i}^{\prime} \mathrm{b}\right) I_{\left\{Y_{i} \geq X_{i}^{\prime} b\right\}} & +(\tau-1)\left(Y_{i}-X_{i}^{\prime} \mathrm{b}\right) I_{\left\{Y_{i}<X_{i}^{\prime} b\right\}} \\
& =  \tag{3}\\
& =\left(Y_{i}-X_{i}^{\prime} \mathrm{b}\right)\left(\tau-I_{\left\{Y_{i}<X_{i}^{\prime} b\right\}}\right)
\end{align*}
$$

where $I_{\{.\}}$is an indicator function equal to 1 if the condition in the subindex is fulfilled, and 0 otherwise. A quantile regression model can be easily fitted, for example in R , by using the function qr of the quantreg R package (Koenker et al., 2018).

Koenker and Machado (1999) proposed an expression to measure the goodness-of-fit of the quantile regression based on a comparison of the values of the objective functions of the estimated model and of the constrained model that only includes an intercept term. Specifically, let

$$
\begin{equation*}
\widehat{V}(\tau)=\sum_{i=1}^{n} \rho_{\tau}\left(Y_{i}-X_{i}^{\prime} \widehat{\beta^{\tau}}\right) \tag{4}
\end{equation*}
$$

be the value of the objective function of the estimated model and

$$
\begin{equation*}
\tilde{V}(\tau)=\sum_{i=1}^{n} \rho_{\tau}\left(Y_{i}-\beta_{0}^{\tau}\right) \tag{5}
\end{equation*}
$$

be the value of the objective function of the constrained model that only includes the intercept term. Then, the goodness-of-fit measure proposed by Koenker and Machado (1999) is

$$
\begin{equation*}
R^{1}(\tau)=1-\hat{V}(\tau) / \tilde{V}(\tau) \tag{6}
\end{equation*}
$$

which is similar to the $R^{2}$ in the multiple linear regression model. Additional details of quantile regression implementation in R can be found in Uribe and Guillen (2020).

### 2.2. Data

The dataset used in this article is the same as that employed in Pérez-Marín et al. (2019b). The data contain the complete portfolio of 9,585 drivers aged 35 years or less with a PHYD policy in a Spanish motor insurance company. All insureds covered in 2010 were included. The description of the variables is presented in Table 1. We know the gender (variable Gender) and age of the driver at the beginning of 2010 (variable Age). Additionally, we also know the total number of kilometres driven during $2010(\mathrm{Km})$, the number of kilometres driven at speeds above the posted limit (Tolerkm, which is our dependent variable), the percentage of kilometres driven on urban roads (Urban) and, finally, the percentage of kilometres driven at night (Night). An urban road is a segment of a road within the boundaries of a builtup area, which is an area with entries and exits especially sign-posted as such and where speed is limited by law at least to $50 \mathrm{Km} / \mathrm{h}$. In order to fit the gamma model, note that 29 observations with zero kilometres driven at speeds above the posted limit - $0.3 \%$ of the size of the dataset - had to be removed from the original data base.

Table 1. Description of variables used in the insurance dataset

| Variable | Description |
| :---: | :---: |
| Tolerkm | Number of kilometres driven at speeds above the posted limit during 2010 |
| Km | Total number of kilometres driven during 2010* |
| Urban | \% of kilometres driven on urban roads during 2010* |
| Night | $\%$ of kilometres driven at night (between midnight and 6 am ) during 2010 |
| Age | Age of the driver at the beginning of 2010 |
| Gender | 1 male, $0=$ female |

As shown in Table 2, Tolerkm presents a positive asymmetry (skewness coefficient $=3.64$ ), with a long tail. The sample comprises $49 \%$ women and $51 \%$ men. The average age of drivers is 24.78 years. The average number of kilometres driven during the observed year was $13,099.91$ (standard deviation of $7,698.98)$. On average, drivers travelled $26.2 \%$ of kilometres on urban roads, $7.02 \%$ of kilometres at night and $1,402.44$ kilometres at speeds above the limit.

Table 2. Descriptive statistics of the insurance data set

|  | Min | $1^{\text {st }}$ Quartile | Median | Mean | $3^{\text {rd }}$ Quartile | Max | St. Dev. | Skewness |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Tolerkm | 0.03 | 285.78 | 692.92 | $1,402.44$ | $1,710.44$ | $23,500.19$ | $1,996.90$ | 3.64 |
| Km | 27.79 | $7,575.15$ | $11,719.83$ | $13,099.91$ | $17,350.12$ | $57,756.98$ | $7,698.98$ | 1.08 |
| Urban | 0.00 | 15.59 | 23.36 | 26.20 | 34.25 | 96.41 | 14.04 | 0.99 |
| Night | 0.00 | 2.49 | 5.32 | 7.02 | 9.85 | 78.56 | 6.12 | 1.67 |
| Age | 18.11 | 22.66 | 24.63 | 24.78 | 26.88 | 35.00 | 2.82 | 0.11 |

## 3. Results

The aim of the paper is to fit the conditional percentile of Tolerkm. In order to build the model, we need to know Tolerkm and some explanatory variables (age, gender, total distance travelled and percentage of urban and nigh ttime driving).

Firstly, we employed a gamma regression model ${ }^{1}$ to fit the conditional mean of Tolerkm and tried different transformations of Km and Urban (including logarithmic and power transformations) to minimize the Akaike information criterion (AIC). We also tried to transform the other two continuous explanatory variables (Night and Age), but this had almost no impact on the AIC score. Power transformations were more effective in order to reduce AIC compared to logarithmic transformations. We sought many power transformations on Km and Urban, specifically, $\mathrm{Km}^{i}$ and $U r b a n^{j}$ where $i=0.05$ to 0.5 increasing by 0.05 , and $j=0.1$ to 1 increasing by 0.1 , and finally the combination that reduced the AIC the most (equal to $149,299.5$ ) was $K m \_t g=\mathrm{Km}^{0.1}$ and Urban_tg $=$ Urban $^{0.7}$ for the gamma model.

The parameter estimates of the corresponding gamma regression model are shown in Table 3. Coefficient estimates with a p-value lower than $1 \%$ correspond to gender, the transformed total number of kilometres driven (Km_tg) and the transformed percentage of kilometres driven in urban areas (Urban_tg). Age effect is only significant at the $10 \%$ level ( $p$-value $=0.0807$ ), probably because the insurance policies were sold exclusively to young drivers. Likewise, the positive effect of percentage of kilometres driven at night (Night) is only significant at the $10 \%$ level ( p -value=$=0.0987$ ), which would indicate that drivers with a higher percentage of night time driving tend to have an average excess speed distance greater than those with a lower percentage of night time driving. $K m \_t g$ has a positive parameter estimate, indicating that an increase in the total number of kilometres driven contributes to increasing the expected number of kilometres driven at speeds above the posted limits. In contrast, Urban_tg presents the opposite effect: the higher the percentage of kilometres driven on urban roads, the lower the expected number of kilometres driven at speeds above the posted limit. Finally, gender (baseline reference: female) has a positive parameter estimate, indicating that men seem to drive more kilometres at speeds above the posted limit than women.

Table 3. Results of the gamma regression model for the insurance data set. Dependent variable is the number of kilometres driven above posted speed limits

|  | Parameter estimate <br> $(\mathrm{p}$-value $)$ |
| :--- | :---: |
| Intercept | -5.126659 |
|  | $(<0.0001)$ |
| Km_tg | 4.966361 |
|  | $(<0.0001)$ |
| Urban_tg | -0.065209 |
| Night | $(<0.0001)$ |
|  | 0.002475 |
|  | $(0.0987)$ |

[^0]| Age | -0.005587 |
| :--- | :---: |
|  | $(0.0807)$ |
| Gender | 0.207654 |
|  | $(<0.0001)$ |

For the quantile regression model we proceed in the same way as in the gamma model, and searched different transformations (including logarithmic and power transformations) and finally those that reduced the AIC the most were $K m \_t q r=K m^{1.7}$ and Urban_tqr $=$ Urban $^{0.1}$.

The parameter estimates and goodness-of-fit of the quantile regression models at different levels ( $\tau=$ $0.5,0.75,0.90,0.95,0.975,0.99$ ) are shown in Table 4. We see that Km_tqr has a significant effect, with a positive parameter estimate, for all levels of the quantile. This means that, for a specific quantile, increasing the total number of kilometres driven increases the quantile of the number of kilometres driven at speeds above the posted limits, ceteris paribus. In contrast, while Urban_tqr also has a significant effect, it has a negative parameter estimate. Thus, as the percentage of kilometres driven in urban areas increases, the quantile of the number of kilometres driven at speeds above the limits decreases. Night has a significant effect only when estimating the median of the kilometres driven at speeds above the limits, but for other levels of the quantile, it has no significant effect. In the case of the median, the parameter estimate is positive, indicating that increasing the percentage of kilometres driven at night increases the median kilometres driven at speeds above the limits. Age has a significant effect only when estimating the quantiles at the $95^{\text {th }}$ and $97.5^{\text {th }}$ levels. In both cases, the corresponding parameters are positive; thus, increasing the driver's age also increases the corresponding percentiles of the distance driven at speeds above the limits. Finally, gender (baseline reference: female) has a significant parameter for all levels of the quantiles up to the $95^{\text {th }}$. The coefficient is positive; thus, men have higher percentile values of distance driven at speeds above the limits than women. In the case of the goodness-of-fit criterion, it is apparent that the contribution explaining the quantiles of the model with covariates vs. the model without increases with the increase in percentile level, reaching $61.22 \%$ at the $99^{\text {th }}$ level. Additionally, in Figure A1 in the Appendix we also provide the marginal effect (estimated parameter) of each explanatory variable in the quantile regression models, as a function of the level of the estimated quantile, showing that the impact of covariates on different percentile levels is not always constant, which highlights the great utility of reference charts as graphical tools.

Figure 1 shows the percentile charts for speeding for males and females, respectively, together with the sample data. The plots show Tolerkm vs. Km, and additionally the grey lines represent the estimated quantiles at different levels. The red line represents the conditional of Tolerkm estimated using the gamma regression model in Table 3. In Figure 1, the values of Urban, Night and Age have been fixed at the mean values in the sample for men and women, respectively. Table 5 provides various examples of percentiles obtained when using the speed reference curves in Figure 1. For example, if a male driver drives $10,000 \mathrm{~km}$ per year and $2,000 \mathrm{~km}$ are above the limit, he is in the 90th percentile curve. On the other hand, the same driver is in the $54^{\text {th }}$ percentile curve if he drives $20,000 \mathrm{~km}$ per year, and finally, he is in the $29^{\text {th }}$ percentile curve if he drives $30,000 \mathrm{~km}$ per year. The corresponding percentiles for women are also shown in Table 5, and are very similar if just a little higher, indicating that women seem to drive at speeds above the posted limit speed less than men. For example, if a woman drives $10,000 \mathrm{~km}$ per year and 2,000 of them at speeds above the limits, she is in the 92th percentile curve.

Table 4. Parameter estimates of the quantile regression model for different percentiles of mileage above the speed limit

|  | $50^{\text {th }}$ <br> percentile <br> (p-value) | $75^{\text {th }}$ <br> percentile <br> (p-value) | $90^{\text {th }}$ <br> percentile (p-value) | $95^{\text {th }}$ <br> percentile (p-value) | $97.5^{\mathrm{th}}$ <br> percentile (p-value) | $99^{\text {th }}$ <br> percentile <br> (p-value) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 1077.1401 | 3359.33609 | 6581.18447 | 8506.75790 | 10173.74522 | 11288.44387 |
|  | (<0.00001) | (<0.00001) | (<0.00001) | (<0.00001) | (<0.00001) | (<0.00001) |
| Km_tqr | 0.00008 | 0.00013 | 0.00020 | 0.00024 | 0.00028 | 0.00032 |


|  | $(<0.00001)$ | $(<0.00001)$ | $(<0.00001)$ | $(<0.00001)$ | $(<0.00001)$ | $(<0.00001)$ |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Urban_tqr | -739.33632 | -2285.83883 | -4529.95550 | -5938.64364 | -7176.65122 | -7960.61577 |
|  | $(<0.00001)$ | $(<0.00001)$ | $(<0.00001)$ | $(<0.00001)$ | $(<0.00001)$ | $(<0.00001)$ |
| Night | 2.36200 | 1.08333 | -0.94645 | 4.77960 | 9.13793 | -0.57912 |
|  | $(0.00224)$ | $(0.43992)$ | $(0.64086)$ | $(0.18409)$ | $(0.26052)$ | $(0.95843)$ |
| Age | -1.80286 | -1.22676 | 6.18758 | 17.61657 | 28.01504 | 37.93405 |
|  | $(0.20585)$ | $0.70373)$ | $(0.25224)$ | $(0.02532)$ | $(0.00894)$ | $(0.09753)$ |
| Gender | 104.81103 | 167.33072 | 167.22760 | 140.47488 | 101.91653 | 189.44456 |
|  | $(<0.00001)$ | $(<0.00001)$ | $(<0.00001)$ | $(0.00360)$ | $(0.16106)$ | $(0.23436)$ |
| Goodness of fit $(\%)$ | 23.43 | 33.59 | 44.56 | 50.89 | 55.55 | 61.22 |

As discussed above, the speeding reference curves shown in Figure 1 have been obtained by assuming that the other explanatory variables (Urban, Night and Age) are equal to the corresponding sample means for men and women, respectively. In Figure 2 we show how these reference curves for the $95^{\text {th }}$ percentile change for different values of Urban (which is the most relevant explanatory variable, apart from Km , for explaining Tolerkm). Specifically, for men we considered values of Urban equal to $8.70,23.45$ and $52.47 \% ~\left(5^{\text {th }}, 50^{\text {th }}\right.$ and $95^{\text {th }}$ percentiles of Urban in the male sample, respectively), and we refer to these values as low, median and high levels of urban driving. Similarly, for women we considered values of Urban equal to 8.37, 23.01 and $53.81 \%\left(5^{\text {th }}, 50^{\text {th }}\right.$ and $95^{\text {th }}$ percentiles of Urban in the female sample, respectively), and similarly we refer to them as low, median and high levels of urban driving. The corresponding reference speed curves for the $95^{\text {th }}$ percentile of Tolerkm are represented in Figure 2 for men and women, respectively, where the red lines are used to represent the corresponding curves for the average values obtained with the gamma regression model (Table 3). We observe that, as the percentage of urban driving increases, all curves move downwards, as Urban_tqr has a negative coefficient. Specifically, in Table 6 we show some examples of the $95^{\text {th }}$ percentile of Tolerkm for certain values of Urban and Km . For a male driver driving $10,000 \mathrm{~km}$ per year the $95^{\text {th }}$ percentile of Tolerkm is equal to $1,753.22 \mathrm{~km}$ if he has a high percentage of urban driving, $2,466.03 \mathrm{~km}$ if he has a median percentage and $3,234.56 \mathrm{~km}$ if he has a low percentage. When the distance driven by the male driver increases to $20,000 \mathrm{~km}$ per year, then the $95^{\text {th }}$ percentile of Tolerkm is equal to $5,120.66,5,803.48$ and $6,572.01 \mathrm{~km}$ for high, median and low percentages of urban driving, respectively. Table 6 also shows the results corresponding to women drivers, and we observe that they are slightly lower than those for male drivers.

Table 5. Percentiles obtained using the percentile charts for examples of speeding (Tolerkm) and total distance driven (Km) by gender.

|  |  | Km |  |  |  |
| :--- | :--- | :--- | :---: | :---: | :---: |
| Tolerkm | Men | 2,000 | 90,000 | 20,000 | 30,000 |
|  |  | 3,000 | $98^{\text {th }}$ | $54^{\text {th }}$ | $29^{\text {th }}$ |
|  | Women | 2,000 | $92^{\text {th }}$ | $44^{\text {th }}$ |  |
|  | 3,000 | $98^{\text {th }}$ | $57^{\text {th }}$ | $30^{\text {th }}$ | $45^{\text {th }}$ |

Other explanatory variables (Urban, Night and Age) are equal to their sample means for men and women, respectively.


Figure 1. Percentile chart for speeding for male drivers (left) and female drivers (right). Tolerkm vs. Km, where grey lines represent the estimated quantiles at different levels. The red line represents the mean of Tolerkm estimated using the gamma regression model in Table 3. Age, urban and night driving are fixed at the mean level by gender.


Figure 2. Percentile chart for speeding for male drivers (left) and female drivers (right) at the $95^{\text {th }}$ level.
Tolerkm vs. Km, where grey lines represent the estimated $95^{\text {th }}$ percentile and red lines represent the mean of Tolerkm estimated using the gamma regression model, for different values of Urban (dashed = low level of urban driving, solid = median level, and dotted = high level). Age and night driving are fixed at the sample mean level by gender.

Table 6. Estimated Tolerkm for the $95^{\text {th }}$ percentile chart for different values of Urban and Km for men and women. Age and night driving are fixed at the sample mean level.

|  |  | Km |  |  |
| :--- | :--- | ---: | ---: | ---: |
|  |  | 10,000 | 20,000 | 30,000 |
| Men | Low Urban | $3,234.56$ | $6,572.01$ | $11,356.31$ |
|  | Median Urban | $2,466.03$ | $5,803.48$ | $10,587.78$ |
|  | High Urban | $1,753.22$ | $5,120.66$ | $9,904.96$ |
| Women | Low Urban | $3,110.06$ | $6,447.51$ | $11,231.81$ |
|  |  | 8 |  |  |


| Median Urban | $2,328.48$ | $5,665.93$ | $10,450.22$ |
| :--- | ---: | ---: | ---: |
| High Urban | $1,607.98$ | $4,945.43$ | $9,729.72$ |

An interactive graphical tool that displays the evolution of a driver's speeding percentile as a function of total distance driven, night-time driving, gender and principal driving zone can be seen in Figure 3 and it can also be accessed online ${ }^{2}$.


Figure 3. Example of interactive speeding percentile chart that locates a particular driver (black dot), given total distance driven, total speeding kilometres and all other reference characteristics stated in the left panel.

## 4. Discussion

We have found that the most relevant variables explaining the number of kilometres driven at speeds above the limits are: total distance driven, percentage of urban driving and gender. This result is in line with those obtained previously by Perez-Marín et al. (2019b), but with a clear improvement in the methodology, more refined results and providing a tool enhancing their applicability (the percentile charts for speeding). Moreover, we also conclude that men have riskier driving patterns compared to women, which is in line with previous research (Høye, 2020; Ayuso et al., 2014; 2016a; 2016b). In most of the models for these data, we found that age and night time driving do not have a significant impact, although they are relevant factors explaining the risk of an accident (see Paefgen et al, 2014, Ayuso et al., 2014 and 2016b). In both cases, this appears to be due to the lack of variability in the PHYD policies, which in our sample were sold exclusively to young drivers. Note that, night time diving could have been included as a categorical variable (indicating whether or not the drivers use the car during the night). Nevertheless, we have percent driving at night (percentage of kilometres travelled at night) and this is why we use this covariate instead of a qualitative indicator.

We analysed the relationship between the distance driven at speeds above the limits and the total distance driven, and found this relationship not to be linear, but rather exponential. This means that as the total distance driven increases, the number of kilometres driven at speeds above the limits also increases, but at an everincreasing rate. This might be due to the driving experience gained or to an excess of confidence on the part of the driver. A nonlinear effect was also found between the driving experience (measured by the distance travelled) and the risk of accident (see Boucher et al. 2013). In that case, the authors found that the risk associated to a higher number of driven kilometres is fully balanced by the larger experience of the driver and

[^1]the other safety factors, and, under certain conditions, the expected frequency of claims is not increased by the number of driven kilometres.

The exponential relationship introduced in the covariates of quantile regressions determines the shape of the reference risk curves for driving at excess speeds. Such models allow the factors associated with higher quantile values to be identified and, hence, for risky drivers to be detected. Our results contribute to calculating the percentile risk score for each driver by controlling for their specific characteristics (and not for the whole population of drivers). Based on these quantile regression models, risk reference curves have been obtained. These graphical tools provide, for each driver, the corresponding percentile of the distance driven at speeds above the limits (which constitutes that driver's risk score), as a function of the total distance driven. Moreover, these curves can be easily obtained for particular types of driver, depending on their characteristics (gender, percentage of urban driving, etc.).

We consider this methodology of risk quantification to be very useful in application with Manage-How-YouDrive (MHYD) insurance products, where the premium is calculated using the same procedure as that used in PHYD insurance, but, in addition, drivers are provided with real-time alerts and recommendations for guaranteeing their safety (Arumugam and Bhargavi, 2019). As such, MHYD insurance improves both customer service and protection in the sector. In this context, the methodology presented here is able to deliver valuable graphical information in terms of preventive early warnings. Estimating just how a driver ranks with respect to distance driven above the posted speed limit is personalized information that should constitute interesting feedback for policy holders (Pérez-Marín et al., 2019b). However, it is necessary to remark that if the result of this study is provided to the driver, drivers could drive at speeds above the legal limit with confidence before they reach a certain speeding mileage. So, we advocate that communication should warn that speeding is always dangerous and should be avoided in any circumstances. Here, it should be stressed that excess speed is perhaps the only feature a driver can easily modify, given that other factors, such as percentage of urban driving, are largely determined by external circumstances and drivers are essentially unable to change them. Indeed, Pérez-Marín et al. (2019a) report that young drivers have a tendency to reduce speed limit violations after an accident, probably because of their greater awareness of the associated risks. As speed is the main cause of severe accidents, those who present lower risk scores (a lower percentile on the risk reference curve) should probably have lower insurance premiums. As to how this ranking should be translated into an insurance price is a question we leave for further research, but there is no doubt that direct bonuses rewarding careful drivers could easily be introduced.

One limitation of the analysis reported here, and which should be pointed out, is that the degree to which drivers exceeded the posted speed limit was not recorded and, therefore, we do not know the severity of the speed violation.

## 5. Conclusions

In this paper, we have presented the design for a prototype graphical tool that displays the evolution of a driver's speeding risk percentile as a function of total distance driven, night time driving, gender and principal driving zone. We conclude that the effect of these covariates on the estimated percentile is not constant. The heterogeneity in the risk of driving above the speed limit can be easily represented using our graphical tool. Moreover, our interface produces a personalized percentile that provides immediate feedback to the user. By measuring a driver's current speeding based on telemetry, that driver can see their evolution over distance driven and they can be provided with a score that is based on their peers' driving records. We firmly believe such reference charts are set to become a standard in the visualization of driving risk.

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## Appendix

Figure A1. Parameter estimates of quantile regression for total kilometres driven above the speed limit at different levels of the quantile. Confidence intervals at a $5 \%$ level of significance are shown as shaded bands. The horizontal red line represents the corresponding parameter estimate in a classical linear regression model.



[^0]:    ${ }^{1}$ We also used lognormal (but it provided a higher AIC score) and inverse Gaussian regressions (but it was eventually discarded because of convergence problems in the algorithm).

[^1]:    ${ }^{2}$ https://riskcenter.shinyapps.io/speeding_risk_reference_chart/

